Non-Bayesian Methods

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Lecture Outline

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- 1. Limitations of Bayesian Decision Theory
- 2. Neyman Pearson Task
- 3. Minimax Task
- 4. Wald Task

Bayesian Decision Theory

Recall:

X set of observations

K set of hidden states

D set of decisions

 $p_{XK}: X \times K \to \mathbb{R}$: joint probability

 $W: K \times D \rightarrow \mathbb{R}: loss function,$

 $q: X \to D$: strategy

R(q): risk:

$$R(q) = \sum_{x \in X} \sum_{k \in K} p_{XK}(x, k) \ W(k, q(x))$$
 (1)

Bayesian strategy q^* :

$$q^* = \operatorname*{argmin}_{q \in X \to D} R(q) \tag{2}$$



The limitations follow from the very ingredients of the Bayesian Decision Theory — the necessity to know all the probabilities and the loss function.

- lacktriangle The loss function W must make sense, but in many tasks it wouldn't
 - medical diagnosis task (W): price of medicines, staff labor, etc. but what penalty in case of patient's death?) Uncomparable penalties on different axes of X.
 - nuclear plant
 - judicial error
- The prior probabilities $p_K(k)$: must exist and be known. But in some cases it does not make sense to talk about probabilities because the events are not random.
 - $K=\{1,2\}\equiv$ {own army plane, enemy plane}; p(x|1), p(x|2) do exist and can be estimated, but p(1) and p(2) don't.
- ♦ The conditionals may be subject to non-random intervention; $p(x \mid k, z)$ where $z \in Z = \{1, 2, 3\}$ are different interventions.
 - a system for handwriting recognition: The training set has been prepared by 3 different persons. But the test set has been constructed by one of the 3 persons only. This **cannot** be done:

(!)
$$p(x | k) = \sum p(z)p(x | k, z)$$
 (3)

- \bullet $K = \{D, N\}$ (dangerous state, normal state)
- X set of observations
- lacktriangle Conditionals $p(x \mid D)$, $p(x \mid N)$ are given
- lacktriangle The priors $p(\mathsf{D})$ and $p(\mathsf{N})$ are unknown or do not exist
- \bullet $q: X \to K$ strategy

The Neyman Person Task looks for the optimal strategy q^* for which

- i) the error of classification of the dangerous state is lower than a predefined threshold $\bar{\epsilon}_D$ (0 < $\bar{\epsilon}_D$ < 1), while
- ii) the classification error for the normal state is as low as possible.

This is formulated as an optimization task with an inequality constraint:

$$q^* = \underset{q:X \to K}{\operatorname{argmin}} \sum_{x:q(x) \neq \mathsf{N}} p(x \mid \mathsf{N}) \tag{4}$$

subject to:
$$\sum_{x: a(x) \neq D} p(x \mid D) \le \overline{\epsilon}_D.$$
 (5)

Neyman Pearson Task



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(copied from the previous slide:)

$$q^* = \underset{q:X \to K}{\operatorname{argmin}} \sum_{x:q(x) \neq \mathsf{N}} p(x \mid \mathsf{N}) \tag{4}$$

subject to:
$$\sum_{x:q(x)\neq D} p(x \mid D) \le \overline{\epsilon}_D.$$
 (5)

A strategy is characterized by the classification error values ϵ_N and ϵ_D :

$$\epsilon_{\mathsf{N}} = \sum_{x:q(x)\neq\mathsf{N}} p(x\,|\,\mathsf{N})$$
 (false alarm) (6)

$$\epsilon_{\mathsf{D}} = \sum_{x: q(x) \neq \mathsf{D}} p(x \mid \mathsf{D})$$
 (overlooked danger) (7)

Example: Male/Female Recognition (Neyman Pearson) (1)

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An aging student at CTU wants to marry. He can't afford to miss recognizing a girl when he meets her, therefore he sets the threshold on female classification error to $\bar{\epsilon}_D = 0.2$. At the same time, he wants to minimize mis-classifying boys for girls.

- $K = \{D, N\} \equiv \{F, M\}$ (female, male)
- ullet measurements $X = \{ ext{short, normal, tall} \} imes \{ ext{ultralight, light, avg, heavy} \}$
- Prior probabilities do not exist.
- Conditionals are given as follows:

p(x F)						
short	.197	.145	.094	.017		
normal	.077	.299	.145	.017		
tall	.001	.008	.000	.000		
u-light		light	avg	heavy		

p(x w)					
short	.011	.005	.011	.011	
normal	.005	.071	.408	.038	
tall	.002	.014	.255	.169	
	u-light	light	avg	heavy	

 $m(m|\mathbf{M})$

(8)

The optimal strategy q^* for a given $x \in X$ is constructed using the likelihood ratio $\frac{p(x \mid N)}{n(x \mid D)}$.

Let there be a constant $\mu \geq 0$. Given this μ , a strategy q is constructed as follows:

$$\frac{p(x \mid \mathsf{N})}{p(x \mid \mathsf{D})} > \mu \quad \Rightarrow \quad q(x) = \mathsf{N},$$

$$\frac{p(x \mid \mathsf{N})}{p(x \mid \mathsf{D})} \le \mu \quad \Rightarrow \quad q(x) = \mathsf{D}.$$
(9)

$$\frac{p(x \mid \mathsf{N})}{p(x \mid \mathsf{D})} \le \mu \quad \Rightarrow \quad q(x) = \mathsf{D} \,. \tag{10}$$

The optimal strategy q^* is obtained by selecting the minimal μ for which there still holds that $\epsilon_{\rm D} \leq \bar{\epsilon}_{\rm D}$.

Let us show this on an example.

Example: Male/Female Recognition (Neyman Pearson) (2)



p(x F)					
short	.197	.145	.094	.017	
normal	.077	.299	.145	.017	
tall	.001	.008	.000	.000	
	light	ght	Ng Ng	eavy	

p(x w)						
short	.011	.005	.011	.011		
normal	.005	.071	.408	.038		
tall	.002	.014	.255	.169		
	u-light	light	avg	heavy		

n(x|M)

r(x) = p(x M)/p(x F)					
short	0.056	0.034	0.117	0.647	
normal	0.065	0.237	2.814	2.235	
tall	2.000	1.750	∞	∞	
	u-light	light	avg	heavy	

rank order of $p(x w)/p(x F)$				
short	2	1	4	6
normal	3	5	10	9
tall	8	7	11	12
	u-light	light	avg	heavy

rapk arder of m(m|M)/m(m|E)

Here, different μ 's can produce 11 different strategies.

First, let us take $2.814 < \mu < \infty$, e.g. $\mu = 3$. This produces a strategy $q^*(x) = \mathsf{F}$ everywhere except where $p(x|\mathsf{F}) = 0$. Obviously, classification error ϵ_F for F is $\epsilon_\mathsf{F} = 0$, and $\epsilon_\mathsf{M} = 1 - .255 - .169 = .576$.

m p

Example: Male/Female Recognition (Neyman Pearson) (3)

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p(x F)						
short	.197	.145	.094	.017		
normal	.077	.299	.145	.017		
tall	.001	.008	.000	.000		
	ı-light	light	avg	леаvу		

p(x M)					
short	.011	.005	.011	.011	
normal	.005	.071	.408	.038	
tall	.002	.014	.255	.169	
	u-light	light	avg	heavy	

	r(x) = p(x M)/p(x F)					
short	0.056	0.034	0.117	0.647		
normal	0.065	0.237	2.814	2.235		
tall	2.000	1.750	∞	∞		
	J-light	light	avg	heavy		

rank, and $q^*(x) = \{F, M\}$ for $\mu = 2.5$				
short	2	1	4	6
normal	3	5	10	9
tall	8	7	11	12
	u-light	light	avg	heavy

Next, take μ which satisfies

$$r_9 < \mu < r_{10} \quad (\text{e.g. } \mu = 2.5)$$
 (11)

(where r_i is the likelihood ratios indexed by its rank.)

Here,
$$\epsilon_{\rm F}=.145$$
, and $\epsilon_{\rm M}=1-.255-.169-.408=.168$.

m p

Example: Male/Female Recognition (Neyman Pearson) (4)

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p(x F)						
short	.197	.145	.094	.017		
normal	.077	.299	.145	.017		
tall	.001	.008	.000	.000		
	-light	light	avg	теаvу		

p(x M)					
short	.011	.005	.011	.011	
normal	.005	.071	.408	.038	
tall	.002	.014	.255	.169	
	u-light	light	avg	heavy	

	r(x) = r	p(x M)/y	p(x F)	
short	0.056	0.034	0.117	0.647
normal	0.065	0.237	2.814	2.235
tall	2.000	1.750	∞	∞
	light	ght	90 20	avy

rank, and	$d q^*(a)$	x) =	$\{F,N$	$\{ \}$ for $\mu=2.1$
short	2	1	4	6
normal	3	5	10	9
tall	8	7	11	12
	u-light	light	avg	heavy

Do the same for μ satisfying

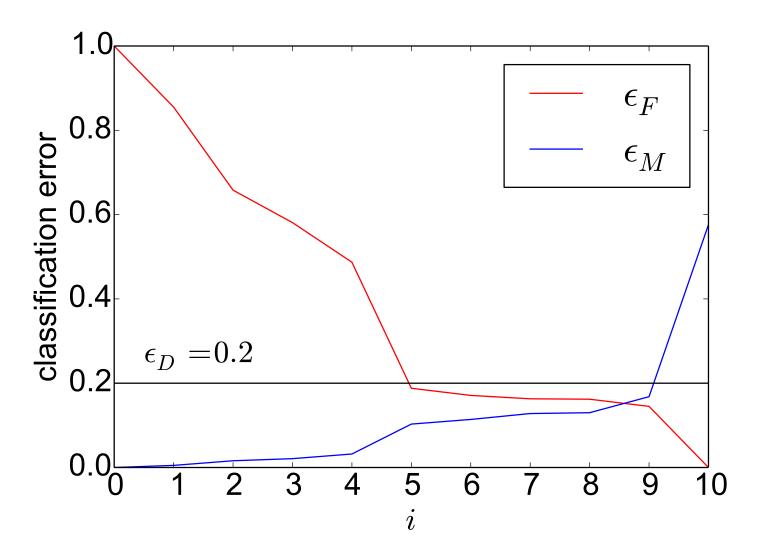
$$r_8 < \mu < r_9$$
 (e.g. $\mu = 2.1$) (12)

$$\Rightarrow \epsilon_{\mathsf{F}} = .162$$
, and $\epsilon_{\mathsf{M}} = 0.13$.

Example: Male/Female Recognition (Neyman Pearson) (5)

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Classification errors for F and M, for $\mu_i = \frac{r_i + r_{i+1}}{2}$ and $\mu_0 = 0$.



The optimum is reached for $r_5 < \mu < r_6$; $\epsilon_{\rm F} = .188$, $\epsilon_{\rm M} = .103$

Neyman Pearson Solution: Illustration of Principle

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Lagrangian of the Neyman Pearson Task is

$$L(q) = \sum_{x: q(x) = D} p(x \mid N) + \mu \left(\sum_{x: q(x) = N} p(x \mid D) - \bar{\epsilon}_D \right)$$
(13)

$$= \underbrace{1 - \sum_{x:q(x)=N} p(x \mid N)}_{p(x \mid N)} + \mu \left(\sum_{x:q(x)=N} p(x \mid D) \right) - \mu \overline{\epsilon}_{D}$$
 (14)

$$=1 - \mu \bar{\epsilon}_{\mathsf{D}} + \sum_{x: q(x)=\mathsf{N}} \underbrace{\{\mu \, p(x \, | \, \mathsf{D}) - p(x \, | \, \mathsf{N})\}}_{T(x)} \tag{15}$$

If T(x) is negative for an x then it will decrease the objective function and the optimal strategy q^* will decide $q^*(x) = N$. This illustrates why the solution to the Neyman Pearson Task has the form

$$\frac{p(x \mid \mathsf{N})}{p(x \mid \mathsf{D})} > \mu \quad \Rightarrow \quad q(x) = \mathsf{N} \,, \tag{9}$$

$$\frac{p(x \mid \mathsf{N})}{p(x \mid \mathsf{D})} \le \mu \quad \Rightarrow \quad q(x) = \mathsf{D} \,. \tag{10}$$

m p

Neyman Pearson: Derivation (1)

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$$q^* = \min_{q:X \to K} \sum_{x:q(x) \neq \mathsf{N}} p(x \mid \mathsf{N}) \qquad \text{subject to: } \sum_{x:q(x) \neq \mathsf{D}} p(x \mid \mathsf{D}) \leq \bar{\epsilon}_{\mathsf{D}}. \tag{16}$$

Let us rewrite this as

$$q^* = \min_{q:X \to K} \sum_{x \in X} \alpha(x) p(x \mid \mathsf{N}) \qquad \text{subject to:} \qquad \sum_{x \in X} [1 - \alpha(x)] p(x \mid \mathsf{D}) \le \bar{\epsilon}_{\mathsf{D}}. \tag{17}$$

and:
$$\alpha(x) \in \{0,1\} \ \forall x \in X$$
 (18)

This is a combinatorial optimization problem. If the relaxation is done from $\alpha(x) \in \{0,1\}$ to $0 \le \alpha(x) \le 1$, this can be solved by **linear programming** (LP). The Lagrangian of this problem with inequality constraints is:

$$L(\alpha(x_1), \alpha(x_2), ..., \alpha(x_N)) = \sum_{x \in X} \alpha(x) p(x \mid N) + \mu \left(\sum_{x \in X} [1 - \alpha(x)] p(x \mid D) - \bar{\epsilon}_D \right)$$
(19)

$$-\sum_{x \in X} \mu_0(x)\alpha(x) + \sum_{x \in X} \mu_1(x)(\alpha(x) - 1)$$
 (20)

e m

Neyman Pearson: Derivation (2)

$$L(\alpha(x_1), \alpha(x_2), ..., \alpha(x_N)) = \sum_{x \in X} \alpha(x) p(x \mid N) + \mu \left(\sum_{x \in X} [1 - \alpha(x)] p(x \mid D) - \bar{\epsilon}_D \right)$$
(19)

$$-\sum_{x \in X} \mu_0(x)\alpha(x) + \sum_{x \in X} \mu_1(x)(\alpha(x) - 1)$$
 (20)

The conditions for optimality are $(\forall x \in X)$:

$$\frac{\partial L}{\partial \alpha(x)} = p(x \mid \mathsf{N}) - \mu p(x \mid \mathsf{D}) - \mu_0(x) + \mu_1(x) = 0, \tag{21}$$

$$\mu \ge 0, \, \mu_0(x) \ge 0, \, \mu_1(x) \ge 0, \quad 0 \le \alpha(x) \le 1,$$
 (22)

$$\mu_0(x)\alpha(x) = 0, \ \mu_1(x)(\alpha(x) - 1) = 0, \ \mu\left(\sum_{x \in X} [1 - \alpha(x)]p(x \mid \mathsf{D}) - \bar{\epsilon}_\mathsf{D}\right) = 0.$$
 (23)

Case-by-case analysis:

case	implications
$\mu = 0$	L minimized by $\alpha(x) = 0 \forall x$
$\mu \neq 0, \alpha(x) = 0$	$\mu_1(x) = 0 \Rightarrow \mu_0(x) = p(x \mid N) - \mu p(x \mid D) \Rightarrow p(x \mid N)/p(x \mid D) \leq \mu$
$\mu \neq 0, \alpha(x) = 1$	$\mu_0(x) = 0 \Rightarrow \mu_1(x) = -[p(x \mid N) - \mu p(x \mid D)] \Rightarrow p(x \mid N)/p(x \mid D) \ge \mu$
$\mu \neq 0,$ $0 < \alpha(x) < 1$	$\mu_0(x) = \mu_1(x) = 0 \Rightarrow p(x \mid \mathbf{N})/p(x \mid \mathbf{D}) = \mu$

Neyman Pearson: Derivation (3)



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Case-by-case analysis:

y
implications
$\int L$ minimized by $\alpha(x) = 0 \forall x$
$\mu_1(x) = 0 \Rightarrow \mu_0(x) = p(x \mid N) - \mu p(x \mid D) \Rightarrow p(x \mid N)/p(x \mid D) \leq \mu$
$\mu_0(x) = 0 \Rightarrow \mu_1(x) = -[p(x \mid N) - \mu p(x \mid D)] \Rightarrow p(x \mid N)/p(x \mid D) \ge \mu$
$\mu_0(x) = \mu_1(x) = 0 \Rightarrow p(x \mid N)/p(x \mid D) = \mu$

Optimal Strategy for a given $\mu \geq 0$ and particular $x \in X$:

$$\frac{p(x \mid \mathsf{N})}{p(x \mid \mathsf{D})} \begin{cases} < \mu & \Rightarrow q(x) = \mathsf{D} \text{ (as } \alpha(x) = 0) \\ > \mu & \Rightarrow q(x) = \mathsf{N} \text{ (as } \alpha(x) = 1) \\ = \mu & \Rightarrow \mathsf{LP} \text{ relaxation does not give the desired solution, as } \alpha \notin \{0, 1\} \end{cases} \tag{24}$$

Neyman Pearson: Note on Randomized Strategies (1)

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Consider:

	p(x D)				
x_1	x_2	x_3			
0.9	0.09	0.01			

	p(x N)				
x_1	x_2	x_3			
0.09	0.9	0.01			

r(x)	= p(x N)/p(x D)
x_1	x_2	x_3
0.1	10	1

and $\bar{\epsilon}_D = 0.03$.

- $q_1:(x_1,x_2,x_3)\to (\mathsf{D},\mathsf{D},\mathsf{D}) \Rightarrow \epsilon_\mathsf{D}=0.00, \, \epsilon_\mathsf{N}=1.00$
- $q_2: (x_1, x_2, x_3) \to (\mathsf{D}, \mathsf{D}, \mathsf{N}) \quad \Rightarrow \quad \epsilon_{\mathsf{D}} = 0.01, \ \epsilon_{\mathsf{N}} = 0.99$
- lacktriangle no other deterministic strategy q is feasible, that is all other ones have $\epsilon_{\mathsf{D}} > \overline{\epsilon}_{\mathsf{D}}$
- ullet q_2 is the best deterministic strategy but it does not comply with the previous basic result of constructing the optimal strategy because it decides for N for likelihood ratio 1 but decides for D for likelihood ratios 0.01 and 10. Why is that?
- lacktriangle we can construct a randomized strategy which attains $\overline{\epsilon}_{\mathsf{D}}$ and reaches lower ϵ_{N} :

$$q(x_1) = q(x_3) = D$$
, $q(x_2) = \begin{cases} N & 1/3 \text{ of the time} \\ D & 2/3 \text{ of the time} \end{cases}$ (25)

For such strategy, $\epsilon_D = 0.03$, $\epsilon_N = 0.7$.

Neyman Pearson: Note on Randomized Strategies (2)



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- lacktriangle This is not a problem but a feature which is caused by discrete nature of X (does not happen when X is continuous).
- This is exactly what the case of $\mu = p(x \mid N)/p(x \mid D)$ is on slide 15.

Neyman Pearson: Notes (1)



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- The task can be generalized to 3 hidden states, of which 2 are dangerous, $K = \{N, D_1, D_2\}$. It is formulated as an analogous problem with two inequality constraints and minimization of classification error for N.
- Neyman's and Pearson's work dates to 1928 and 1933.
- A particular strength of the approach lies in that the likelihood ratio r(x) or even $p(x \mid N)$ need not be known. For the task to be solved, it is enough to know the $p(x \mid D)$ and the **rank order** of the likelihood ratio (to be demonstrated on the next page)

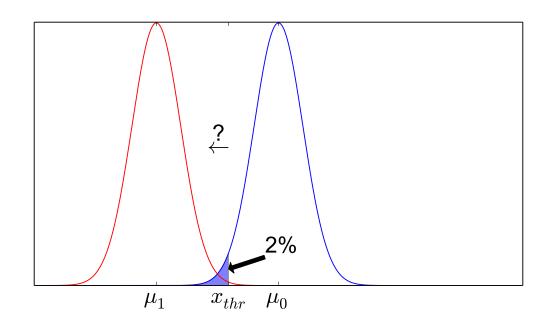
(8)

| | |

Neyman Pearson : Notes (2)

The likelihood ratio is

- Consider a medicine for reducing weight. The normal population has a distribution of weight $p(x \mid D)$ as shown in blue. Let it be normal, $p(x \mid D) = \mathcal{N}(x \mid \mu_0, \sigma)$. The distribution of weights after 1 month of taking the medicine is assumed to be normal as well, with the same variance but uknown shift of mean to the left, $p(x \mid N) = \mathcal{N}(x \mid \mu_1, \sigma)$, with $\mu_1 < \mu_0$ but otherwise unknown (shown in red).
 - $r(x) = \exp \frac{1}{2\sigma^2} \left(-(x \mu_1)^2 + (x \mu_0)^2 \right) = \exp \left(\frac{1}{\sigma^2} (\mu_1 \mu_0) x + \text{const} \right)$. It is thus decreasing (monotone) with x (irrespective of μ_1 , $\mu_1 < \mu_0$).
- Setting $\bar{\epsilon}_D = 0.02$, we go along the decreasing r(x) and find the point x_{thr} for which $\int_{-\infty}^{x_{thr}} p(x \mid D) = \bar{\epsilon}_D = 0.02$ (0.02-quantile). Note that the threshold μ on r(x) is still uknown as $p(x \mid N)$ is unknown.



- \bullet $K = \{1, 2, ..., N\}$
- X set of observations
- lacktriangle Conditionals $p(x \mid k)$ are known $\forall k \in K$
- lacktriangle The priors p(k) are unknown or do not exist
- $lack q \colon X \to K$ strategy

The Minimax Task looks for the optimum strategy q^* which minimizes the classification error of the worst classified class:

$$q^* = \underset{q:X \to K}{\operatorname{argmin}} \max_{k \in K} \epsilon(k), \quad \text{where}$$
(26)

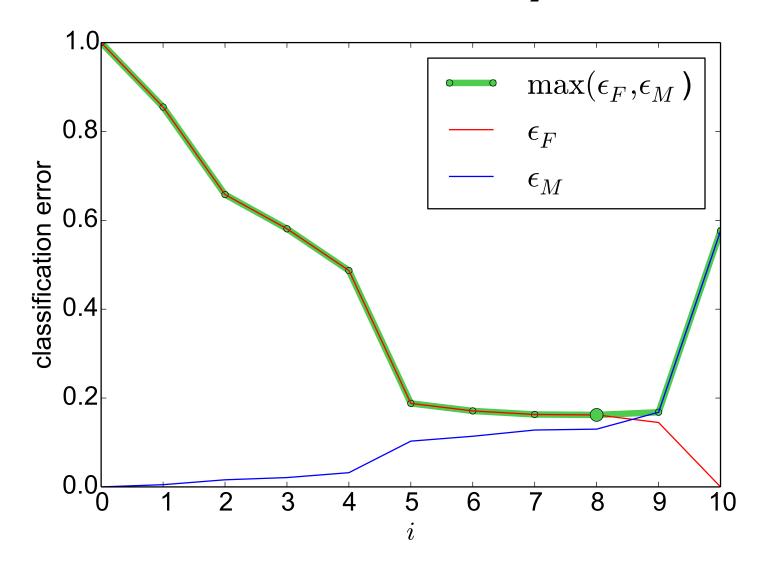
$$\epsilon(k) = \sum_{x: \, q(x) \neq k} p(x \mid k) \tag{27}$$

- Example: A recognition algorithm qualifies for a competition using preliminary tests.
 During the final competition, only objects from the hardest-to-classify class are used.
- For a 2-class problem, the strategy is again constructed using the likelihood ratio.
- In the case of continuous observations space X, equality of classification errors is attained: $\epsilon_1 = \epsilon_2$
- The derivation can again be done using Linear Programming.

Example: Male/Female Recognition (Minimax)



Classification errors for F and M, for $\mu_i = \frac{r_i + r_{i+1}}{2}$ and $\mu_0 = 0$.



The optimum is attained for i=8, $\epsilon_{\rm F}=.162$, $\epsilon_{\rm M}=.13$. The corresponding strategy is as shown on slide 11.

Minimax: Comparison with Bayesian Decision with Unknown Priors

- Consider the same setting as in the Minimax task, but let the priors p(k) exist but be unknown.
- lacktriangle The Bayesian error ϵ for strategy q is

$$\epsilon = \sum_{k} \sum_{x: q(x) \neq k} p(x, k) = \sum_{k} p(k) \underbrace{\sum_{x: q(x) \neq k} p(x \mid k)}_{\epsilon(k)}$$
(28)

- We want to minimize ϵ but we do not know p(k)'s. What is the maximum it can attain? Obviously, the p(k)'s do the convex combination of the class errors $\epsilon(k)$; the maximum Bayesian error will be attained when p(k)=1 for the class k with the highest class error $\epsilon(k)$.
- ullet Thus, to minimize the Bayesian error ϵ under this setting, the solution is to minimize the error of the hardest-to-classify class.
- Therefore, Minimax formulation and the Bayesian formulation with Unknown Priors lead to the same solution.

Wald Task (1)



- Let us consider classification with two states, $K = \{1, 2\}$.
- We want to set a threshold ϵ on the classification error of both of the classes: $\epsilon_1 \leq \epsilon$, $\epsilon_2 \leq \epsilon$.
- ullet As the previous analysis shows (Neyman Pearson, Minimax), there may be **no** feasible solution if ϵ is set too low.
- That is why the possibility of decision "do not know" is introduced. Thus $D = K \cup \{?\}$
- lack A strategy q:X o D is characterized by:

$$\epsilon_1 = \sum_{x: q(x)=2} p(x \mid 1)$$
 (classification error for 1) (29)

$$\epsilon_2 = \sum_{x: q(x)=1} p(x \mid 2)$$
 (classification error for 2) (30)

$$\kappa_1 = \sum_{x: q(x)=?} p(x \mid 1) \quad \text{(undecided rate for 1)} \tag{31}$$

$$\kappa_2 = \sum_{x: q(x)=?} p(x \mid 2) \quad \text{(undecided rate for 2)} \tag{32}$$

Wald Task (2)

lacktriangle The optimal strategy q^* :

$$q^* = \underset{q:X \to D}{\operatorname{argmin}} \max_{i = \{1,2\}} \kappa_i \tag{33}$$

subject to:
$$\epsilon_1 \le \epsilon, \ \epsilon_2 \le \epsilon$$
 (34)

- The task is again solvable using LP (even for more than 2 classes)
- The optimal solution is again based on the likelihood ratio

$$r(x) = \frac{p(x \mid 1)}{p(x \mid 2)} \tag{35}$$

lacktriangle The optimal strategy is constructed using suitably chosen thresholds μ_l and μ_h such that:

$$q(x) = \begin{cases} 2 & \text{for } r(x) < \mu_l \\ 1 & \text{for } r(x) > \mu_h \\ ? & \text{for } \mu_l \le r(x) \le \mu_h \end{cases}$$

$$(36)$$





Solve the Wald task for $\epsilon = 0.05$.

(୮ \			
	p(x F)		

short	.197	.145	.094	.017
normal	.077	.299	.145	.017
tall	.001	.008	.000	.000
	u-light	light	avg	heavy

$$p(x|\mathsf{M})$$

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short	.011	.005	.011	.011
normal	.005	.071	.408	.038
tall	.002	.014	.255	.169
	u-light	light	avg	heavy

$$r(x) = p(x|\mathsf{M})/p(x|\mathsf{F})$$

short	0.056	0.034	0.117	0.647
normal	0.065	0.237	2.814	2.235
tall	2.000	1.750	∞	∞
	u-light	light	avg	heavy

rank, and
$$q^*(x) = \{F, M, ?\}$$

	u-light	light	avg	heavy
tall	8	7	11	12
normal	3	5	10	9
short	2	1	4	6

Result: $\epsilon_{\rm M}=0.032$, $\epsilon_{\rm F}=0$, $\kappa_{\rm M}=0.544$, $\kappa_{\rm F}=0.487$

$$(r_4 < \mu_l < r_5, r_{10} < \mu_h < \infty)$$