

Non-Bayesian Methods

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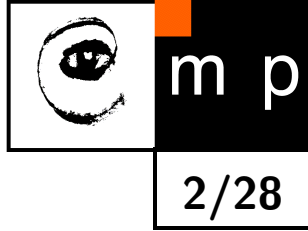
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Lecture Outline



1. Limitations of Bayesian Decision Theory
2. Neyman Pearson Task
3. Minimax Task
4. Wald Task

Bayesian Decision Theory



Recall:

X set of observations

K set of hidden states

D set of decisions

p_{XK} : $X \times K \rightarrow \mathbb{R}$: joint probability

W : $K \times D \rightarrow \mathbb{R}$: *loss function*,

q : $X \rightarrow D$: strategy

$R(q)$: risk:

$$R(q) = \sum_{x \in X} \sum_{k \in K} p_{XK}(x, k) W(k, q(x)) \quad (1)$$

Bayesian strategy q^* :

$$q^* = \operatorname{argmin}_{q \in X \rightarrow D} R(q) \quad (2)$$

Limitations of the Bayesian Decision Theory

The limitations follow from the very ingredients of the Bayesian Decision Theory — the necessity to know all the probabilities and the loss function.

- ◆ The loss function W must make sense, but in many tasks it wouldn't
 - medical diagnosis task (W : price of medicines, staff labor, etc. but what penalty in case of patient's death?) Uncomparable penalties on different axes of X .
 - nuclear plant
 - judicial error
- ◆ The prior probabilities $p_K(k)$: must exist and be known. But in some cases it does not make sense to talk about probabilities because the events are not random.
 - $K = \{1, 2\} \equiv \{\text{own army plane, enemy plane}\}$;
 $p(x|1)$, $p(x|2)$ do exist and can be estimated, but $p(1)$ and $p(2)$ don't.
- ◆ The conditionals may be subject to non-random intervention; $p(x | k, z)$ where $z \in Z = \{1, 2, 3\}$ are different interventions.
 - a system for handwriting recognition: The training set has been prepared by 3 different persons. But the test set has been constructed by one of the 3 persons only. This **cannot** be done:

$$(!) \quad p(x | k) = \sum_z p(z)p(x | k, z) \quad (3)$$

Neyman Pearson Task

- ◆ $K = \{1, 2\}$ (two classes, sometimes called 1='dangerous', 2='normal')
- ◆ X set of observations
- ◆ Conditionals $p(x | 1)$, $p(x | 2)$ are given
- ◆ The priors $p(1)$ and $p(2)$ are unknown or do not exist
- ◆ $q: X \rightarrow K$ strategy

The Neyman Pearson Task looks for the optimal strategy q^* for which

- i) the error of classification for class 1 is lower than a predefined threshold $\bar{\epsilon}_1$ ($0 < \bar{\epsilon}_1 < 1$), while
- ii) the classification error for class 2 is as low as possible.

This is formulated as an optimization task with an inequality constraint:

$$q^* = \operatorname{argmin}_{q: X \rightarrow K} \sum_{x: q(x) \neq 2} p(x | 2) \quad (4)$$

$$\text{subject to: } \sum_{x: q(x) \neq 1} p(x | 1) \leq \bar{\epsilon}_1. \quad (5)$$

Neyman Pearson Task

(copied from the previous slide:)

$$q^* = \operatorname{argmin}_{q: X \rightarrow K} \sum_{x: q(x) \neq 2} p(x | 2) \quad (4)$$

$$\text{subject to: } \sum_{x: q(x) \neq 1} p(x | 1) \leq \bar{\epsilon}_1. \quad (5)$$

A strategy is characterized by the classification error values ϵ_2 and ϵ_1 :

$$\epsilon_1 = \sum_{x: q(x) \neq 1} p(x | 1) \quad (6)$$

$$\epsilon_2 = \sum_{x: q(x) \neq 2} p(x | 2) \quad (7)$$

Example: Male/Female Recognition (Neyman Pearson) (1)

A hotel has an advertising screen in an elevator. Based on recognition of gender, it wants to display a relevant advert for a shopping mall located at the ground floor. The shopping mall is primarily designed to be interesting for female customers. For this reason, the female classification error threshold is set to $\bar{\epsilon}_1 = 0.2$. At the same time, the objective is to minimize mis-classification of male customers.

- ◆ $K = \{1, 2\} \equiv \{F, M\}$ (female, male)
- ◆ measurements $X = \text{height} \times \text{weight}$ (height sensor = simple optical sensor, weight sensor = standard component of elevators)
- ◆ height $\in \{h_1, h_2, h_3\}$, weight $\in \{w_1, w_2, w_3, w_4\}$ ($h_1 < h_2 < h_3$), ($w_1 < w_2 < w_3 < w_4$)
- ◆ Prior probabilities do not exist.
- ◆ Conditionals are given as follows:

$$p(x|F)$$

h_1	.197	.145	.094	.017
h_2	.077	.299	.145	.017
h_3	.001	.008	.000	.000
	w_1	w_2	w_3	w_4

$$p(x|M)$$

h_1	.011	.005	.011	.011
h_2	.005	.071	.408	.038
h_3	.002	.014	.255	.169
	w_1	w_2	w_3	w_4

(8)

Neyman Pearson : Solution

The optimal strategy q^* for a given $x \in X$ is constructed using the likelihood ratio $\frac{p(x|2)}{p(x|1)}$.

Let there be a constant $\mu \geq 0$. Given this μ , a strategy q is constructed as follows:

$$\frac{p(x|2)}{p(x|1)} > \mu \quad \Rightarrow \quad q(x) = 2, \quad (9)$$

$$\frac{p(x|2)}{p(x|1)} \leq \mu \quad \Rightarrow \quad q(x) = 1. \quad (10)$$

The optimal strategy q^* is obtained by selecting the minimal μ for which there still holds that $\epsilon_1 \leq \bar{\epsilon}_1$.

Let us show this on an example.

Example: Male/Female Recognition (Neyman Pearson) (2)

$p(x|1)$

h_1	.197	.145	.094	.017
h_2	.077	.299	.145	.017
h_3	.001	.008	.000	.000
	w_1	w_2	w_3	w_4

$p(x|2)$

h_1	.011	.005	.011	.011
h_2	.005	.071	.408	.038
h_3	.002	.014	.255	.169
	w_1	w_2	w_3	w_4

$r(x) = p(x|2)/p(x|1)$

h_1	0.056	0.034	0.117	0.647
h_2	0.065	0.237	2.814	2.235
h_3	2.000	1.750	∞	∞
	w_1	w_2	w_3	w_4

rank order of $p(x|2)/p(x|1)$

h_1	2	1	4	6
h_2	3	5	10	9
h_3	8	7	11	12
	w_1	w_2	w_3	w_4

Here, different μ 's can produce 11 different strategies.

First, let us take $2.814 < \mu < \infty$, e.g. $\mu = 3$. This produces a strategy $q^*(x) = 1$ everywhere except where $p(x|1) = 0$. Obviously, classification error $\epsilon_1 = 0$, and $\epsilon_2 = 1 - .255 - .169 = .576$.

Example: Male/Female Recognition (Neyman Pearson) (3)

$p(x|1)$

h_1	.197	.145	.094	.017
h_2	.077	.299	.145	.017
h_3	.001	.008	.000	.000
	w_1	w_2	w_3	w_4

$p(x|2)$

h_1	.011	.005	.011	.011
h_2	.005	.071	.408	.038
h_3	.002	.014	.255	.169
	w_1	w_2	w_3	w_4

$r(x) = p(x|2)/p(x|1)$

h_1	0.056	0.034	0.117	0.647
h_2	0.065	0.237	2.814	2.235
h_3	2.000	1.750	∞	∞
	w_1	w_2	w_3	w_4

rank, and $q^*(x) = \{1, 2\}$ for $\mu = 2.5$

h_1	2	1	4	6
h_2	3	5	10	9
h_3	8	7	11	12
	w_1	w_2	w_3	w_4

Next, take μ which satisfies

$$r_9 < \mu < r_{10} \quad (\text{e.g. } \mu = 2.5) \tag{11}$$

(where r_i is the likelihood ratios indexed by its rank.)

Here, $\epsilon_1 = .145$, and $\epsilon_2 = 1 - .255 - .169 - .408 = .168$.

Example: Male/Female Recognition (Neyman Pearson) (4)

 $p(x|1)$

h_1	.197	.145	.094	.017
h_2	.077	.299	.145	.017
h_3	.001	.008	.000	.000
	w_1	w_2	w_3	w_4

 $p(x|2)$

h_1	.011	.005	.011	.011
h_2	.005	.071	.408	.038
h_3	.002	.014	.255	.169
	w_1	w_2	w_3	w_4

 $r(x) = p(x|2)/p(x|1)$

h_1	0.056	0.034	0.117	0.647
h_2	0.065	0.237	2.814	2.235
h_3	2.000	1.750	∞	∞
	w_1	w_2	w_3	w_4

 rank, and $q^*(x) = \{1, 2\}$ for $\mu = 2.1$

h_1	2	1	4	6
h_2	3	5	10	9
h_3	8	7	11	12
	w_1	w_2	w_3	w_4

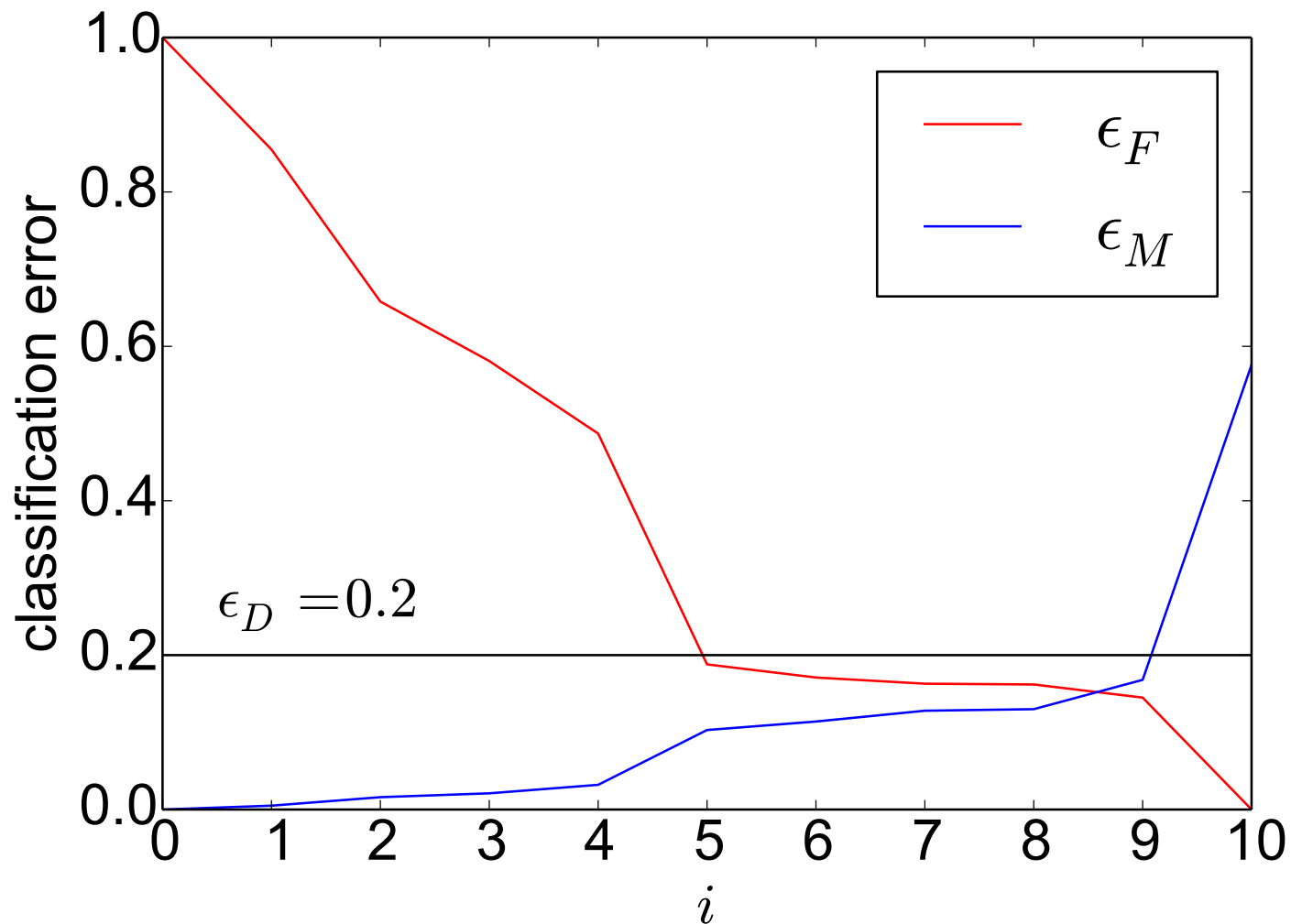
Do the same for μ satisfying

$$r_8 < \mu < r_9 \quad (\text{e.g. } \mu = 2.1) \tag{12}$$

$\Rightarrow \epsilon_1 = .162$, and $\epsilon_2 = 0.13$.

Example: Male/Female Recognition (Neyman Pearson) (5)

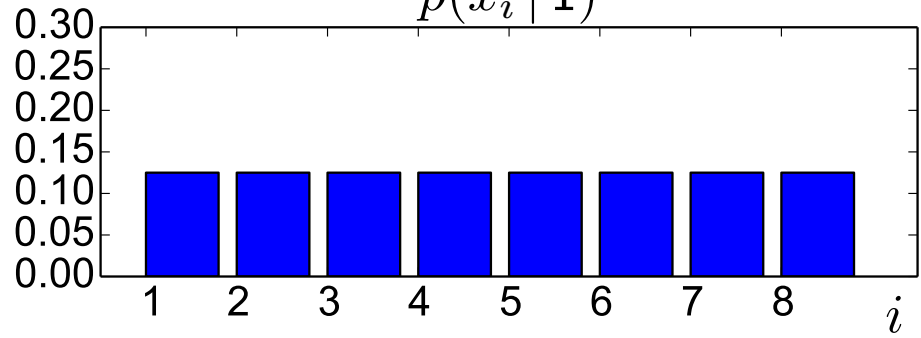
Classification errors for 1 and 2, for $\mu_i = \frac{r_i+r_{i+1}}{2}$ and $\mu_0 = 0$.



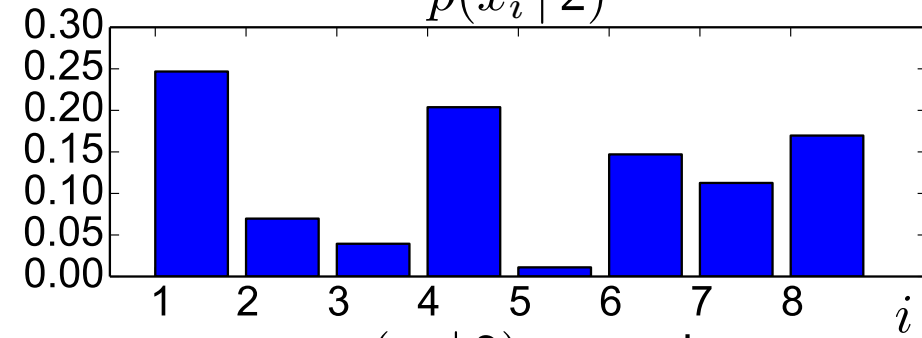
The optimum is reached for $r_5 < \mu < r_6$; $\epsilon_1 = .188$, $\epsilon_2 = .103$

Neyman Pearson : Simple Case (1)

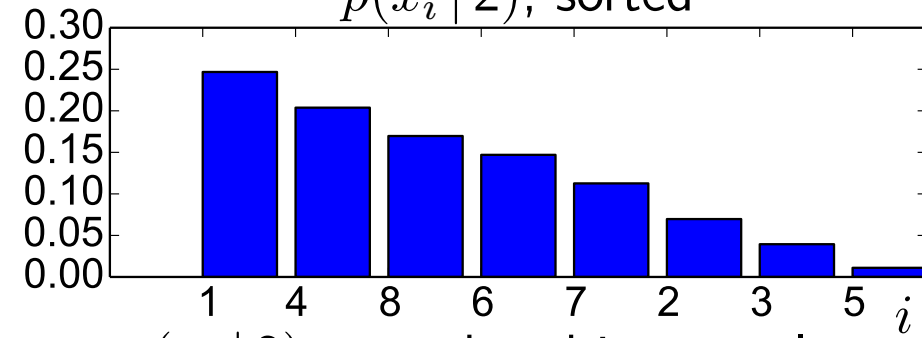
$p(x_i | 1)$



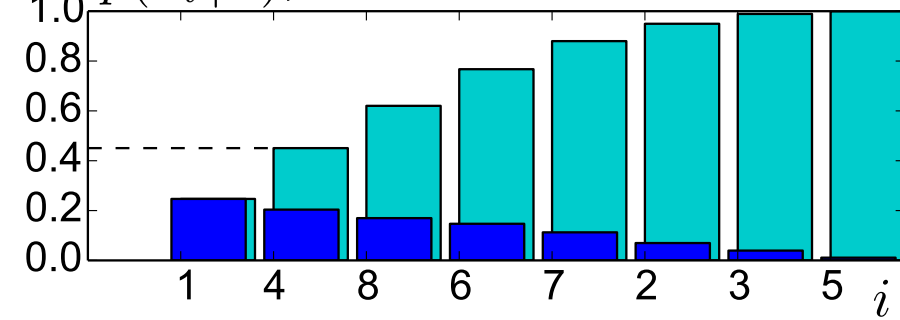
$p(x_i | 2)$



$p(x_i | 2), \text{ sorted}$



$p(x_i | 2), \text{ sorted and its cumul. sum } [\bullet]$



Consider a simple case when $p(x_i | 1) = \text{const}$. Possible values for ϵ_1 are $0, \frac{1}{8}, \frac{2}{8}, \dots, 1$. If a strategy q classifies P observations as normal then $\epsilon_1 = \frac{P}{8}$.
 If $P = 1$ then $\epsilon_1 = \frac{1}{8}$ and it is clear that ϵ_2 will attain minimum if the (one) observation which is classified as normal is the one with the highest $p(x_i | 2)$. Similarly, if $P = 2$ then the two observations to be classified as normal are the one with the first two highest $p(x_i | 2)$. Etc.

↑ cumulative sum of sorted $p(x_i | 2)$ shows the classification success rate for 2, that is, $1 - \epsilon_2$, for $\epsilon_1 = \frac{1}{8}, \frac{2}{8}, \dots, 1$. For example, for $\epsilon_1 = \frac{2}{8}$ ($P = 2$), $\epsilon_2 = 1 - 0.45 = 0.55$ (as shown, dashed.)

Neyman Pearson : Towards General Case (2)

In general, $p(x_i | 1) \neq \text{const.}$ Consider the following example:

$p(x_i 1)$		
x_1	x_2	x_3
0.5	0.25	0.25

$p(x_i 2)$		
x_1	x_2	x_3
0.6	0.35	0.05

But this can easily be converted to the previous special case by (only formally) splitting x_1 to two observations x'_1 and x''_1 :

$p(x_i 1)$			
x'_1	x''_1	x_2	x_3
0.25	0.25	0.25	0.25

$p(x_i 2)$			
x'_1	x''_1	x_2	x_3
0.3	0.3	0.35	0.05

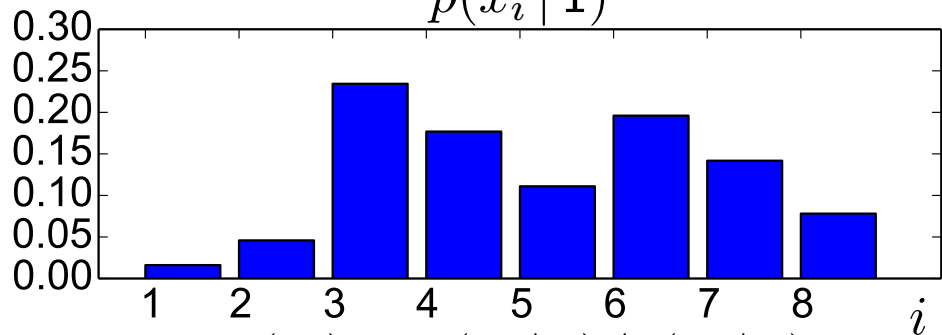
which would result in ordering the observations by decreasing $p(x_i | 2)$ as: x_2, x_1, x_3 .

Obviously, the same ordering is obtained when $p(x_i | 2)$ is 'normalized' by $p(x_i | 1)$, that is, using the likelihood ratio

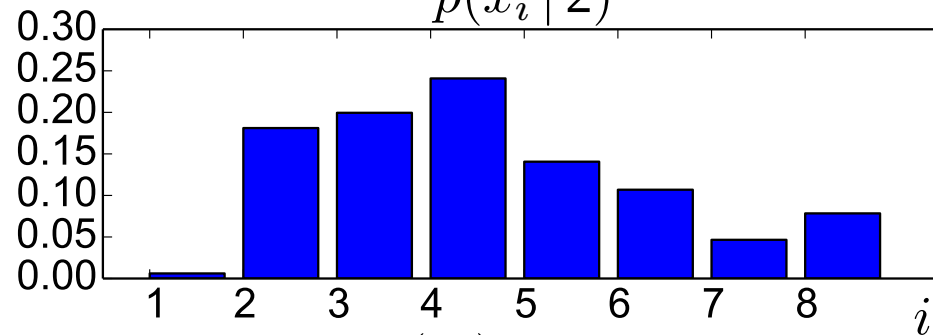
$$r(x_i) = \frac{p(x_i | 2)}{p(x_i | 1)}. \tag{13}$$

Neyman Pearson : General Case Example (3)

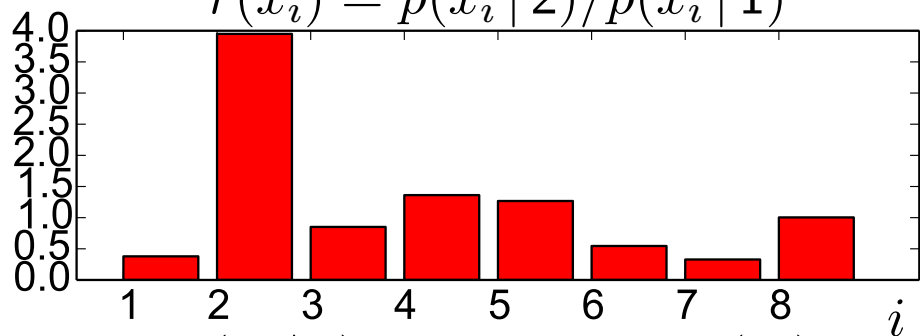
$p(x_i | 1)$



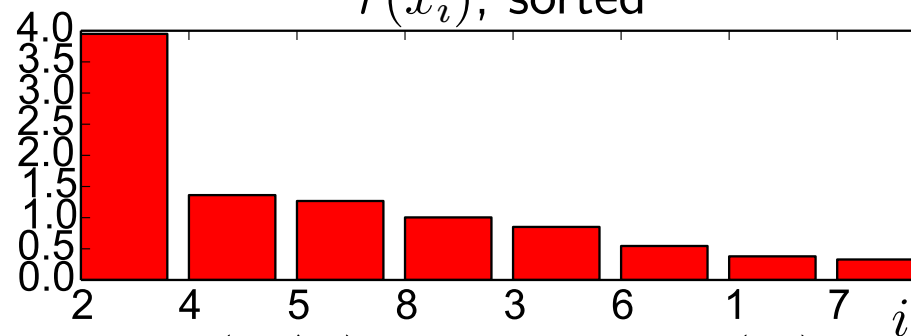
$p(x_i | 2)$



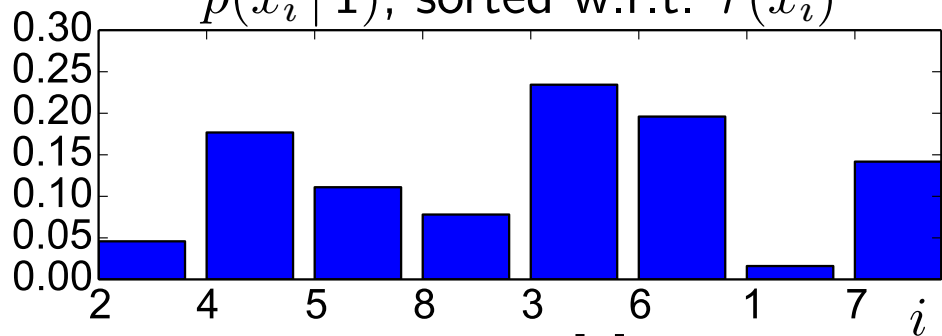
$r(x_i) = p(x_i | 2) / p(x_i | 1)$



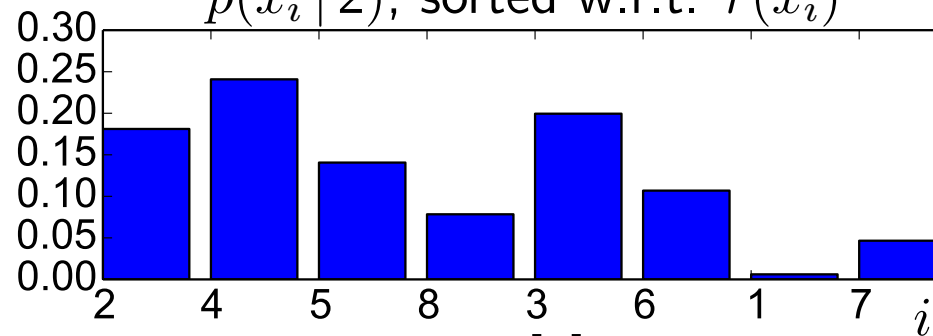
$r(x_i)$, sorted



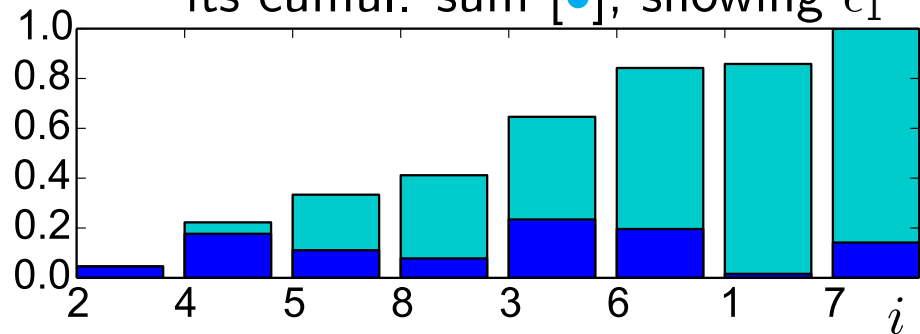
$p(x_i | 1)$, sorted w.r.t. $r(x_i)$



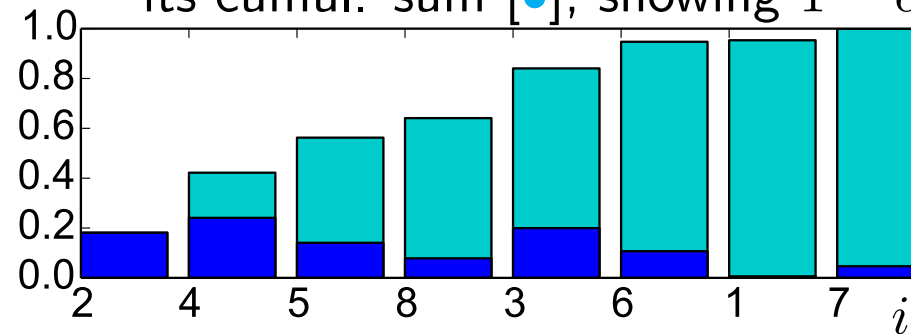
$p(x_i | 2)$, sorted w.r.t. $r(x_i)$



its cumul. sum [●], showing ϵ_1



its cumul. sum [●], showing $1 - \epsilon_2$



Neyman Pearson Solution : Illustration of Principle

Lagrangian of the Neyman Pearson Task is

$$L(q) = \underbrace{\sum_{x: q(x)=1} p(x|2)}_{=} + \mu \left(\sum_{x: q(x)=2} p(x|1) - \bar{\epsilon}_D \right) \quad (14)$$

$$= 1 - \sum_{x: q(x)=2} p(x|2) + \mu \left(\sum_{x: q(x)=2} p(x|1) \right) - \mu \bar{\epsilon}_1 \quad (15)$$

$$= 1 - \mu \bar{\epsilon}_1 + \sum_{x: q(x)=2} \underbrace{\{\mu p(x|1) - p(x|2)\}}_{T(x)} \quad (16)$$

If $T(x)$ is negative for an x then it will decrease the objective function and the optimal strategy q^* will decide $q^*(x) = 2$. This illustrates why the solution to the Neyman Pearson Task has the form

$$\frac{p(x|2)}{p(x|1)} > \mu \quad \Rightarrow \quad q(x) = 2, \quad (9)$$

$$\frac{p(x|2)}{p(x|1)} \leq \mu \quad \Rightarrow \quad q(x) = 1. \quad (10)$$

Neyman Pearson : Derivation (1)

$$q^* = \min_{q: X \rightarrow K} \sum_{x: q(x) \neq 2} p(x | 2) \quad \text{subject to:} \quad \sum_{x: q(x) \neq 1} p(x | 1) \leq \bar{\epsilon}_1. \quad (17)$$

Let us rewrite this as

$$q^* = \min_{q: X \rightarrow K} \sum_{x \in X} \alpha(x) p(x | 2) \quad \text{subject to:} \quad \sum_{x \in X} [1 - \alpha(x)] p(x | 1) \leq \bar{\epsilon}_1. \quad (18)$$

$$\text{and:} \quad \alpha(x) \in \{0, 1\} \quad \forall x \in X \quad (19)$$

This is a combinatorial optimization problem. If the relaxation is done from $\alpha(x) \in \{0, 1\}$ to $0 \leq \alpha(x) \leq 1$, this can be solved by **linear programming** (LP). The Lagrangian of this problem with inequality constraints is:

$$L(\alpha(x_1), \alpha(x_2), \dots, \alpha(x_N)) = \sum_{x \in X} \alpha(x) p(x | 2) + \mu \left(\sum_{x \in X} [1 - \alpha(x)] p(x | 1) - \bar{\epsilon}_1 \right) \quad (20)$$

$$- \sum_{x \in X} \mu_0(x) \alpha(x) + \sum_{x \in X} \mu_1(x) (\alpha(x) - 1) \quad (21)$$

Neyman Pearson : Derivation (2)

$$L(\alpha(x_1), \alpha(x_2), \dots, \alpha(x_N)) = \sum_{x \in X} \alpha(x)p(x | 2) + \mu \left(\sum_{x \in X} [1 - \alpha(x)]p(x | 1) - \bar{\epsilon}_1 \right) \quad (20)$$

$$- \sum_{x \in X} \mu_0(x)\alpha(x) + \sum_{x \in X} \mu_1(x)(\alpha(x) - 1) \quad (21)$$

The conditions for optimality are ($\forall x \in X$):

$$\frac{\partial L}{\partial \alpha(x)} = p(x | 2) - \mu p(x | 1) - \mu_0(x) + \mu_1(x) = 0, \quad (22)$$

$$\mu \geq 0, \mu_0(x) \geq 0, \mu_1(x) \geq 0, \quad 0 \leq \alpha(x) \leq 1, \quad (23)$$

$$\mu_0(x)\alpha(x) = 0, \mu_1(x)(\alpha(x) - 1) = 0, \mu \left(\sum_{x \in X} [1 - \alpha(x)]p(x | 1) - \bar{\epsilon}_1 \right) = 0. \quad (24)$$

Case-by-case analysis:

case	implications
$\mu = 0$	L minimized by $\alpha(x) = 0 \quad \forall x$
$\mu \neq 0, \alpha(x) = 0$	$\mu_1(x) = 0 \Rightarrow \mu_0(x) = p(x 2) - \mu p(x 1) \Rightarrow p(x 2)/p(x 1) \leq \mu$
$\mu \neq 0, \alpha(x) = 1$	$\mu_0(x) = 0 \Rightarrow \mu_1(x) = -[p(x 2) - \mu p(x 1)] \Rightarrow p(x 2)/p(x 1) \geq \mu$
$\mu \neq 0,$ $0 < \alpha(x) < 1$	$\mu_0(x) = \mu_1(x) = 0 \Rightarrow p(x 2)/p(x 1) = \mu$

Neyman Pearson : Derivation (3)

Case-by-case analysis:

case	implications
$\mu = 0$	L minimized by $\alpha(x) = 0 \quad \forall x$
$\mu \neq 0, \alpha(x) = 0$	$\mu_1(x) = 0 \Rightarrow \mu_0(x) = p(x 2) - \mu p(x 1) \Rightarrow p(x 2)/p(x 1) \leq \mu$
$\mu \neq 0, \alpha(x) = 1$	$\mu_0(x) = 0 \Rightarrow \mu_1(x) = -[p(x 2) - \mu p(x 1)] \Rightarrow p(x 2)/p(x 1) \geq \mu$
$\mu \neq 0,$ $0 < \alpha(x) < 1$	$\mu_0(x) = \mu_1(x) = 0 \Rightarrow p(x 2)/p(x 1) = \mu$

Optimal Strategy for a given $\mu \geq 0$ and particular $x \in X$:

$$\frac{p(x | 2)}{p(x | 1)} \begin{cases} < \mu & \Rightarrow q(x) = 1 \text{ (as } \alpha(x) = 0) \\ > \mu & \Rightarrow q(x) = 2 \text{ (as } \alpha(x) = 1) \\ = \mu & \Rightarrow \text{LP relaxation does not give the desired solution, as } \alpha \notin \{0, 1\} \end{cases} \quad (25)$$

Neyman Pearson : Note on Randomized Strategies (1)

Consider:

$p(x 1)$		
x_1	x_2	x_3
0.9	0.09	0.01

$p(x 2)$		
x_1	x_2	x_3
0.09	0.9	0.01

$r(x) = p(x 2)/p(x 1)$		
x_1	x_2	x_3
0.1	10	1

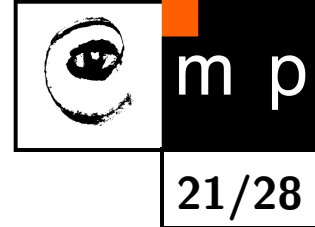
and $\bar{\epsilon}_1 = 0.03$.

- ◆ $q_1 : (x_1, x_2, x_3) \rightarrow (1, 1, 1) \Rightarrow \epsilon_1 = 0.00, \epsilon_2 = 1.00$
- ◆ $q_2 : (x_1, x_2, x_3) \rightarrow (1, 1, 2) \Rightarrow \epsilon_1 = 0.01, \epsilon_2 = 0.99$
- ◆ no other deterministic strategy q is feasible, that is all other ones have $\epsilon_1 > \bar{\epsilon}_1$
- ◆ q_2 is the best deterministic strategy but it does not comply with the previous basic result of constructing the optimal strategy because it decides for 2 for likelihood ratio 1 but decides for 1 for likelihood ratios 0.01 and 10. Why is that?
- ◆ we can construct a randomized strategy which attains $\bar{\epsilon}_1$ and reaches lower ϵ_2 :

$$q(x_1) = q(x_3) = 1, \quad q(x_2) = \begin{cases} 2 & 1/3 \text{ of the time} \\ 1 & 2/3 \text{ of the time} \end{cases} \quad (26)$$

For such strategy, $\epsilon_1 = 0.03, \epsilon_2 = 0.7$.

Neyman Pearson : Note on Randomized Strategies (2)



- ◆ This is not a problem but a feature which is caused by discrete nature of X (does not happen when X is continuous).
- ◆ This is exactly what the case of $\mu = p(x | 2)/p(x | 1)$ is on slide 18.

Neyman Pearson : Notes (1)

- ◆ The task can be generalized to 3 hidden states, of which 2 are dangerous, $K = \{2, D_1, D_2\}$. It is formulated as an analogous problem with two inequality constraints and minimization of classification error for 2.
- ◆ Neyman's and Pearson's work dates to 1928 and 1933.
- ◆ A particular strength of the approach lies in that the likelihood ratio $r(x)$ or even $p(x | 2)$ need not be known. For the task to be solved, it is enough to know the $p(x | 1)$ and the **rank order** of the likelihood ratio (to be demonstrated on the next page)

Minimax Task

- ◆ $K = \{1, 2, \dots, N\}$
- ◆ X set of observations
- ◆ Conditionals $p(x | k)$ are known $\forall k \in K$
- ◆ The priors $p(k)$ are unknown or do not exist
- ◆ $q: X \rightarrow K$ strategy

The Minimax Task looks for the optimum strategy q^* which minimizes the classification error of the worst classified class:

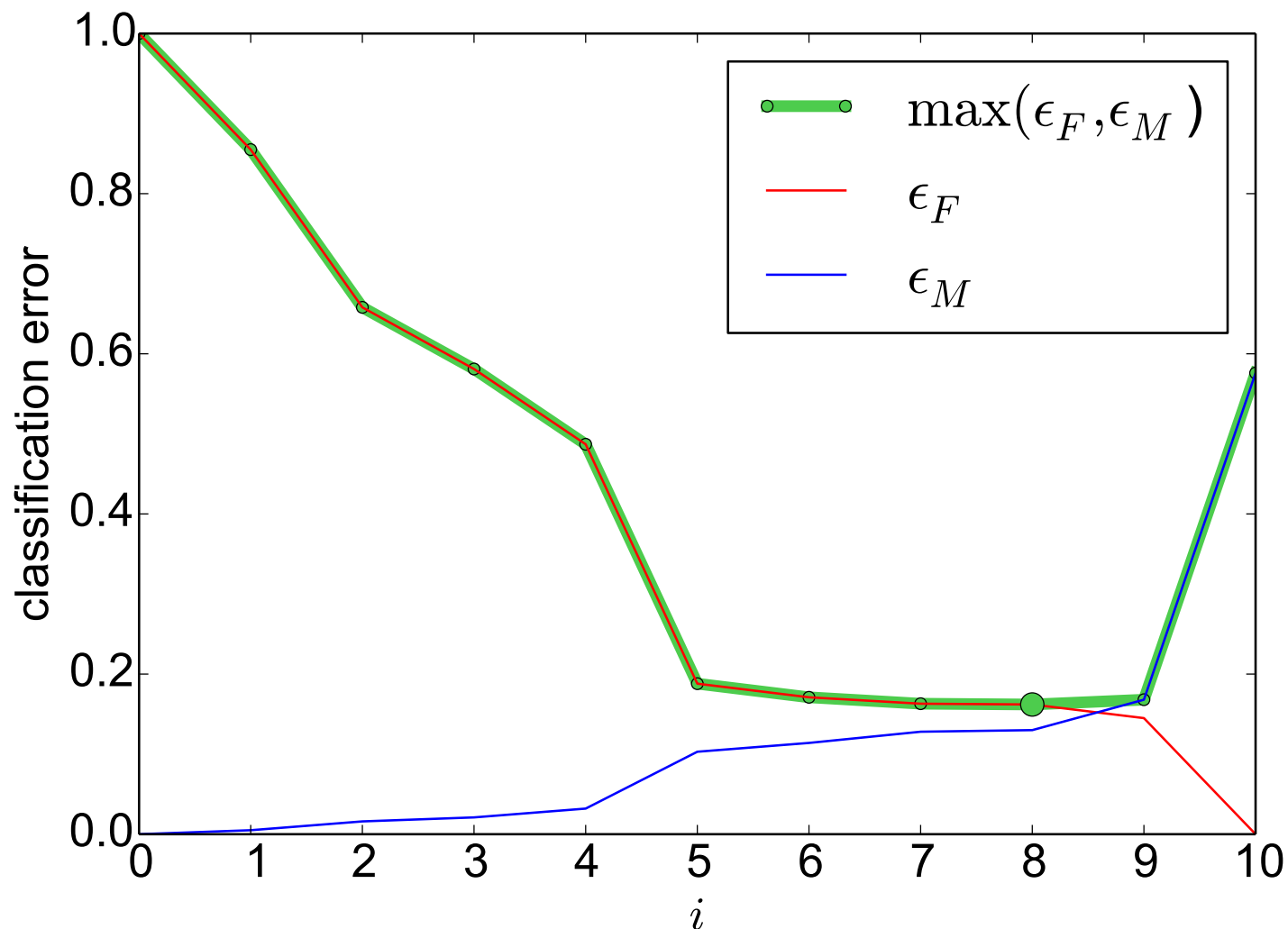
$$q^* = \operatorname{argmin}_{q: X \rightarrow K} \max_{k \in K} \epsilon(k), \quad \text{where} \quad (27)$$

$$\epsilon(k) = \sum_{x: q(x) \neq k} p(x | k) \quad (28)$$

- ◆ Example: A recognition algorithm qualifies for a competition using preliminary tests. During the final competition, only objects from the hardest-to-classify class are used.
- ◆ For a 2-class problem, the strategy is again constructed using the likelihood ratio.
- ◆ In the case of continuous observations space X , equality of classification errors is attained: $\epsilon_1 = \epsilon_2$
- ◆ The derivation can again be done using Linear Programming.

Example: Male/Female Recognition (Minimax)

Classification errors for 1 and 2, for $\mu_i = \frac{r_i+r_{i+1}}{2}$ and $\mu_0 = 0$.



The optimum is attained for $i = 8$, $\epsilon_1 = .162$, $\epsilon_2 = .13$. The corresponding strategy is as shown on slide 11.

Minimax: Comparison with Bayesian Decision with Unknown Priors

- ◆ Consider the same setting as in the Minimax task, but let the priors $p(k)$ exist but be unknown.
- ◆ The Bayesian error ϵ for strategy q is

$$\epsilon = \sum_k \sum_{x: q(x) \neq k} p(x, k) = \sum_k p(k) \underbrace{\sum_{x: q(x) \neq k} p(x | k)}_{\epsilon(k)} \quad (29)$$

- ◆ We want to minimize ϵ but we do not know $p(k)$'s. What is the maximum it can attain? Obviously, the $p(k)$'s do the convex combination of the class errors $\epsilon(k)$; the maximum Bayesian error will be attained when $p(k) = 1$ for the class k with the highest class error $\epsilon(k)$.
- ◆ Thus, to minimize the Bayesian error ϵ under this setting, the solution is to minimize the error of the hardest-to-classify class.
- ◆ Therefore, Minimax formulation and the Bayesian formulation with Unknown Priors lead to the same solution.

Wald Task (1)

- ◆ Let us consider classification with two states, $K = \{1, 2\}$.
- ◆ We want to set a threshold ϵ on the classification error of both of the classes: $\epsilon_1 \leq \epsilon$, $\epsilon_2 \leq \epsilon$.
- ◆ It is clear that there may be **no** feasible solution if ϵ is set too low.
- ◆ That is why the possibility of decision “do not know” is introduced. Thus $D = K \cup \{?\}$
- ◆ A strategy $q : X \rightarrow D$ is characterized by:

$$\epsilon_1 = \sum_{x: q(x)=2} p(x | 1) \quad (\text{classification error for 1}) \quad (30)$$

$$\epsilon_2 = \sum_{x: q(x)=1} p(x | 2) \quad (\text{classification error for 2}) \quad (31)$$

$$\kappa_1 = \sum_{x: q(x)=?} p(x | 1) \quad (\text{undecided rate for 1}) \quad (32)$$

$$\kappa_2 = \sum_{x: q(x)=?} p(x | 2) \quad (\text{undecided rate for 2}) \quad (33)$$

Wald Task (2)

- ◆ The optimal strategy q^* :

$$q^* = \operatorname{argmin}_{q: X \rightarrow D} \max_{i=\{1,2\}} \kappa_i \quad (34)$$

$$\text{subject to: } \epsilon_1 \leq \epsilon, \epsilon_2 \leq \epsilon \quad (35)$$

- ◆ The task is again solvable using LP (even for more than 2 classes)
- ◆ The optimal solution is again based on the likelihood ratio

$$r(x) = \frac{p(x | 1)}{p(x | 2)} \quad (36)$$

- ◆ The optimal strategy is constructed using suitably chosen thresholds μ_l and μ_h such that:

$$q(x) = \begin{cases} 2 & \text{for } r(x) < \mu_l \\ 1 & \text{for } r(x) > \mu_h \\ ? & \text{for } \mu_l \leq r(x) \leq \mu_h \end{cases} \quad (37)$$

Example: Male/Female Recognition (Wald)

Solve the Wald task for $\epsilon = 0.05$.

$p(x|F)$

h_1	.197	.145	.094	.017
h_2	.077	.299	.145	.017
h_3	.001	.008	.000	.000
	w_1	w_2	w_3	w_4

$p(x|M)$

h_1	.011	.005	.011	.011
h_2	.005	.071	.408	.038
h_3	.002	.014	.255	.169
	w_1	w_2	w_3	w_4

$r(x) = p(x|2)/p(x|1)$

h_1	0.056	0.034	0.117	0.647
h_2	0.065	0.237	2.814	2.235
h_3	2.000	1.750	∞	∞
	w_1	w_2	w_3	w_4

rank, and $q^*(x) = \{1, 2, ?\}$

h_1	2	1	4	6
h_2	3	5	10	9
h_3	8	7	11	12
	w_1	w_2	w_3	w_4

Result: $\epsilon_2 = 0.032$, $\epsilon_1 = 0$, $\kappa_2 = 0.544$, $\kappa_1 = 0.487$

$(r_4 < \mu_l < r_5, r_{10} < \mu_h < \infty)$