

# 3D Computer Vision

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Open Informatics Master's Course

## ► Example: Fitting A Circle To Scattered Points

**Problem:** Fit an origin-centered circle  $\mathcal{C}$ :  $\|\mathbf{x}\|^2 - r^2 = 0$  to a set of 2D points  $Z = \{x_i\}_{i=1}^k$

1. consider radial error as the 'algebraic error'  $\epsilon(\mathbf{x}) = \|\mathbf{x}\|^2 - r^2$  'arbitrary' choice
2. linearize it at  $\hat{\mathbf{x}}$  we are dropping  $i$  in  $\epsilon_i$ ,  $\mathbf{e}_i$  etc for clarity

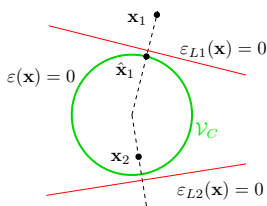
$$\epsilon(\hat{\mathbf{x}}) \approx \epsilon(\mathbf{x}) + \underbrace{\frac{\partial \epsilon(\mathbf{x})}{\partial \mathbf{x}}}_{\mathbf{J}(\mathbf{x})=2\mathbf{x}^\top} \underbrace{(\hat{\mathbf{x}} - \mathbf{x})}_{\mathbf{e}(\hat{\mathbf{x}}, \mathbf{x})} = \dots = 2\mathbf{x}^\top \hat{\mathbf{x}} - (r^2 + \|\mathbf{x}\|^2) \stackrel{\text{def}}{=} \epsilon_L(\hat{\mathbf{x}})$$

$\epsilon_L(\hat{\mathbf{x}}) = 0$  is a line with normal  $\frac{\mathbf{x}}{\|\mathbf{x}\|}$  and intercept  $\frac{r^2 + \|\mathbf{x}\|^2}{2\|\mathbf{x}\|}$  not tangent to  $\mathcal{C}$ , outside!

3. using (18), express error approximation  $\mathbf{e}^*$  as

$$\|\mathbf{e}^*\|^2 = \boldsymbol{\epsilon}^\top (\mathbf{J}\mathbf{J}^\top)^{-1} \boldsymbol{\epsilon} = \frac{(\|\mathbf{x}\|^2 - r^2)^2}{4\|\mathbf{x}\|^2}$$

4. fit circle



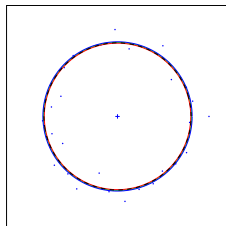
$$r^* = \arg \min_r \sum_{i=1}^k \frac{(\|\mathbf{x}_i\|^2 - r^2)^2}{4\|\mathbf{x}_i\|^2} = \dots = \left( \frac{1}{k} \sum_{i=1}^k \frac{1}{\|\mathbf{x}_i\|^2} \right)^{-\frac{1}{2}}$$

- this example results in a convex quadratic optimization problem
- note that

$$\arg \min_r \sum_{i=1}^k (\|\mathbf{x}_i\|^2 - r^2)^2 = \left( \frac{1}{k} \sum_{i=1}^k \|\mathbf{x}_i\|^2 \right)^{\frac{1}{2}}$$

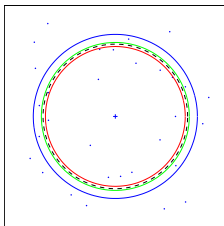
# Circle Fitting: Some Results

medium radial noise



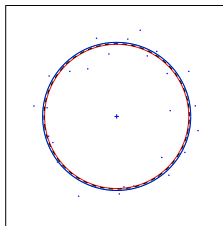
opt: 1.8, Smp: 1.9, dir: 2.3

big radial noise



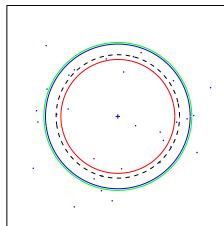
1.6, 1.8, 2.6

medium isotropic noise



1.8, 2.0, 2.2

big isotropic noise



1.6, 2.0, 2.4

mean ranks over 10000 random trials with  $k = 32$  samples

- green – ground truth
- red – Sampson error  $e$  minimizer
- blue – direct radial error  $\epsilon$  minimizer
- black – optimal estimator for isotropic error

optimal estimator for isotropic error (black, dashed):

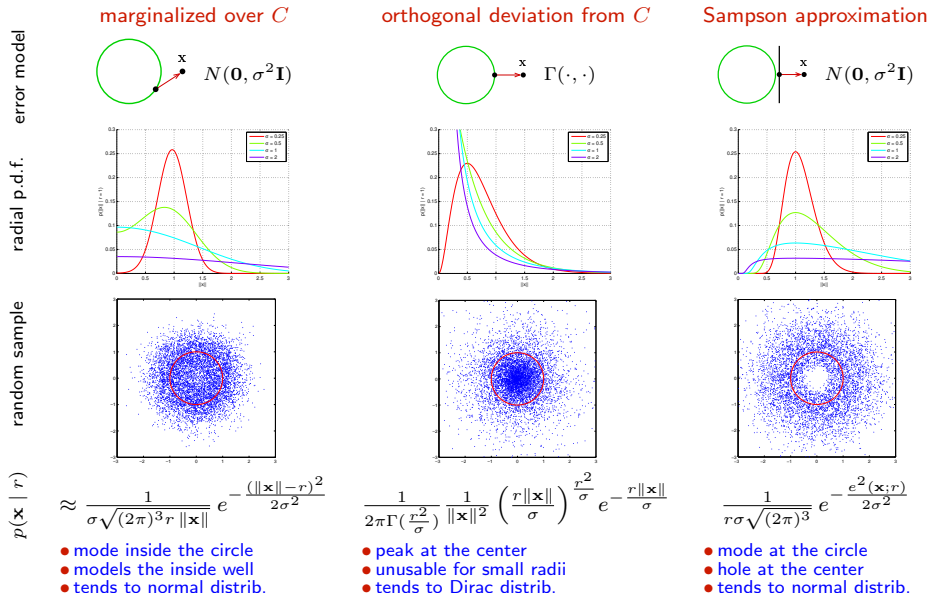
$$r \approx \frac{3}{4k} \sum_{i=1}^k \|\mathbf{x}_i\| + \sqrt{\left( \frac{3}{4k} \sum_{i=1}^k \|\mathbf{x}_i\| \right)^2 - \frac{1}{2k} \sum_{i=1}^k \|\mathbf{x}_i\|^2}$$

## which method is better?

- error should model noise, radial noise and isotropic noise behave differently
- ground truth: Normally distributed isotropic error, Gamma-distributed radial error
- Sampson: better for the radial distribution model; Direct: better for the isotropic model
- no matter how corrected, the algebraic error minimizer is not an unbiased parameter estimator  
Cramér-Rao bound tells us how close one can get with unbiased estimator and given  $k$

# Discussion: On The Art of Probabilistic Model Design...

- a few models for fitting zero-centered circle  $C$  of radius  $r$  to points in  $\mathbb{R}^2$



## ► Sampson Error for Fundamental Matrix Manifold

The epipolar algebraic error is

$$\varepsilon_i(\mathbf{F}) = \underline{\mathbf{y}}_i^\top \mathbf{F} \underline{\mathbf{x}}_i, \quad \mathbf{x}_i = (u_i^1, v_i^1), \quad \mathbf{y}_i = (u_i^2, v_i^2), \quad \varepsilon_i \in \mathbb{R}$$

Let  $\mathbf{F} = [\mathbf{F}_1 \quad \mathbf{F}_2 \quad \mathbf{F}_3]$  (per columns) =  $\begin{bmatrix} (\mathbf{F}^1)^\top \\ (\mathbf{F}^2)^\top \\ (\mathbf{F}^3)^\top \end{bmatrix}$  (per rows),  $\mathbf{S} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \end{bmatrix}$ , then

### Sampson

$$\mathbf{J}_i(\mathbf{F}) = \left[ \frac{\partial \varepsilon_i(\mathbf{F})}{\partial u_i^1}, \frac{\partial \varepsilon_i(\mathbf{F})}{\partial v_i^1}, \frac{\partial \varepsilon_i(\mathbf{F})}{\partial u_i^2}, \frac{\partial \varepsilon_i(\mathbf{F})}{\partial v_i^2} \right] \quad \mathbf{J}_i \in \mathbb{R}^{1,4} \quad \text{derivatives over point coordinates}$$

$$= \left[ (\mathbf{F}_1)^\top \underline{\mathbf{y}}_i, (\mathbf{F}_2)^\top \underline{\mathbf{y}}_i, (\mathbf{F}^1)^\top \underline{\mathbf{x}}_i, (\mathbf{F}^2)^\top \underline{\mathbf{x}}_i \right] = \begin{bmatrix} \mathbf{S}\mathbf{F}^\top \underline{\mathbf{y}}_i \\ \mathbf{S}\mathbf{F} \underline{\mathbf{x}}_i \end{bmatrix}^\top$$

$$\mathbf{e}_i(\mathbf{F}) = -\frac{\mathbf{J}_i(\mathbf{F}) \varepsilon_i(\mathbf{F})}{\|\mathbf{J}_i(\mathbf{F})\|^2} \quad \mathbf{e}_i(\mathbf{F}) \in \mathbb{R}^4 \quad \text{Sampson error vector}$$

$$e_i(\mathbf{F}) \stackrel{\text{def}}{=} \|\mathbf{e}_i(\mathbf{F})\| = \frac{\varepsilon_i(\mathbf{F})}{\|\mathbf{J}_i(\mathbf{F})\|} = \frac{\underline{\mathbf{y}}_i^\top \mathbf{F} \underline{\mathbf{x}}_i}{\sqrt{\|\mathbf{S}\mathbf{F} \underline{\mathbf{x}}_i\|^2 + \|\mathbf{S}\mathbf{F}^\top \underline{\mathbf{y}}_i\|^2}} \quad e_i(\mathbf{F}) \in \mathbb{R} \quad \text{scalar Sampson error}$$

- Sampson error 'normalizes' the algebraic error
- automatically copes with multiplicative factors  $\mathbf{F} \mapsto \lambda \mathbf{F}$
- actual optimization not yet covered →110

## ► Back to Triangulation: The Golden Standard Method

Given  $\mathbf{P}_1$ ,  $\mathbf{P}_2$  and a correspondence  $x \leftrightarrow y$ , look for 3D point  $\mathbf{X}$  projecting to  $x$  and  $y \rightarrow 89$

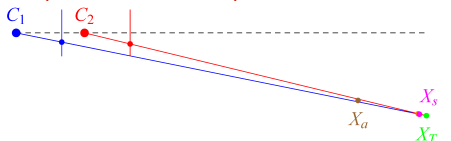
Idea:

1. if not given, compute  $\mathbf{F}$  from  $\mathbf{P}_1$ ,  $\mathbf{P}_2$ , e.g.  $\mathbf{F} = (\mathbf{Q}_1 \mathbf{Q}_2^{-1})^\top [\mathbf{q}_1 - (\mathbf{Q}_1 \mathbf{Q}_2^{-1}) \mathbf{q}_2]_{\times}$
2. correct the measurement by the linear estimate of the correction vector  $\rightarrow 101$

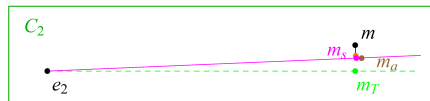
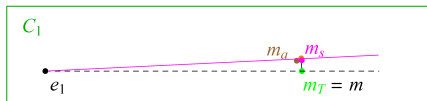
$$\begin{bmatrix} \hat{u}^1 \\ \hat{v}^1 \\ \hat{u}^2 \\ \hat{v}^2 \end{bmatrix} \approx \begin{bmatrix} u^1 \\ v^1 \\ u^2 \\ v^2 \end{bmatrix} - \frac{\varepsilon}{\|\mathbf{J}\|^2} \mathbf{J}^\top = \begin{bmatrix} u^1 \\ v^1 \\ u^2 \\ v^2 \end{bmatrix} - \frac{\underline{\mathbf{y}}^\top \mathbf{F} \underline{\mathbf{x}}}{\|\mathbf{S} \mathbf{F} \underline{\mathbf{x}}\|^2 + \|\mathbf{S} \mathbf{F}^\top \underline{\mathbf{y}}\|^2} \begin{bmatrix} (\mathbf{F}_1)^\top \underline{\mathbf{y}} \\ (\mathbf{F}_2)^\top \underline{\mathbf{y}} \\ (\mathbf{F}^1)^\top \underline{\mathbf{x}} \\ (\mathbf{F}^2)^\top \underline{\mathbf{x}} \end{bmatrix}$$

3. use the SVD triangulation algorithm with numerical conditioning  $\rightarrow 90$

Ex (cont'd from  $\rightarrow 93$ ):



- $X_T$  – noiseless ground truth position
- – reprojection error minimizer
- $X_s$  – Sampson-corrected algebraic error minimizer
- $X_a$  – algebraic error minimizer
- $m$  – measurement ( $m_T$  with noise in  $v^2$ )



## ► Back to Fundamental Matrix Estimation

**Goal:** Given a set  $X = \{(x_i, y_i)\}_{i=1}^k$  of  $k \gg 7$  inlier correspondences, compute a statistically efficient estimate for fundamental matrix  $\mathbf{F}$ .

### What we have so far

- 7-point algorithm for  $\mathbf{F}$  (5-point algorithm for  $\mathbf{E}$ ) → 84
- definition of Sampson error per correspondence  $e_i(\mathbf{F} \mid x_i, y_i)$  → 105
- triangulation requiring an optimal  $\mathbf{F}$

### What we need

- correspondence recognition
- an optimization algorithm for many ( $k \gg 7$ ) correspondences

$$\mathbf{F}^* = \arg \min_{\mathbf{F}} \sum_{i=1}^k e_i^2(\mathbf{F} \mid X)$$

- the 7-point estimate is a good starting point  $\mathbf{F}_0$

# Levenberg-Marquardt (LM) Iterative Optimization in a Nutshell

Consider error function  $\mathbf{e}_i(\boldsymbol{\theta}) = f(\mathbf{x}_i, \mathbf{y}_i, \boldsymbol{\theta}) \in \mathbb{R}^m$ , with  $\mathbf{x}_i, \mathbf{y}_i$  given,  $\boldsymbol{\theta} \in \mathbb{R}^q$  unknown  
 $\boldsymbol{\theta} = \mathbf{F}$ ,  $q = 9$ ,  $m = 1$  for f.m. estimation

**Our goal:**  $\boldsymbol{\theta}^* = \arg \min_{\boldsymbol{\theta}} \sum_{i=1}^k \|\mathbf{e}_i(\boldsymbol{\theta})\|^2$

**Idea 1** (Gauss-Newton approximation): proceed iteratively for  $s = 0, 1, 2, \dots$

$$\boldsymbol{\theta}^{s+1} := \boldsymbol{\theta}^s + \mathbf{d}_s, \quad \text{where } \mathbf{d}_s = \arg \min_{\mathbf{d}} \sum_{i=1}^k \|\mathbf{e}_i(\boldsymbol{\theta}^s + \mathbf{d})\|^2 \quad (19)$$

$$\mathbf{e}_i(\boldsymbol{\theta}^s + \mathbf{d}) \approx \mathbf{e}_i(\boldsymbol{\theta}^s) + \mathbf{L}_i \mathbf{d},$$

$$(\mathbf{L}_i)_{jl} = \frac{\partial (\mathbf{e}_i(\boldsymbol{\theta}))_j}{\partial (\boldsymbol{\theta})_l}, \quad \mathbf{L}_i \in \mathbb{R}^{m,q} \quad \text{typically a long matrix, } m \ll q$$

Then the solution to Problem (19) is a set of 'normal' eqs

$$-\underbrace{\sum_{i=1}^k \mathbf{L}_i^\top \mathbf{e}_i(\boldsymbol{\theta}^s)}_{\mathbf{e} \in \mathbb{R}^{q,1}} = \underbrace{\left( \sum_{i=1}^k \mathbf{L}_i^\top \mathbf{L}_i \right)}_{\mathbf{L} \in \mathbb{R}^{q,q}} \mathbf{d}_s, \quad (20)$$

- $\mathbf{d}_s$  can be solved for by Gaussian elimination using Choleski decomposition of  $\mathbf{L}$   
 $\mathbf{L}$  symmetric PSD  $\Rightarrow$  use Choleski, almost  $2\times$  faster than Gauss-Seidel, see bundle adjustment  $\rightarrow 141$
- beware of rank deficiency in  $\mathbf{L}$  when  $k$  is small
- such updates do not lead to stable convergence  $\rightarrow$  ideas of Levenberg and Marquardt



**Idea 2** (Levenberg): replace  $\sum_i \mathbf{L}_i^\top \mathbf{L}_i$  with  $\sum_i \mathbf{L}_i^\top \mathbf{L}_i + \lambda \mathbf{I}$  for some damping factor  $\lambda \geq 0$

**Idea 3** (Marquardt): replace  $\lambda \mathbf{I}$  with  $\lambda \sum_i \text{diag}(\mathbf{L}_i^\top \mathbf{L}_i)$  to adapt to local curvature:

$$-\sum_{i=1}^k \mathbf{L}_i^\top \mathbf{e}_i(\boldsymbol{\theta}^s) = \left( \sum_{i=1}^k (\mathbf{L}_i^\top \mathbf{L}_i + \lambda \text{diag}(\mathbf{L}_i^\top \mathbf{L}_i)) \right) \mathbf{d}_s$$

**Idea 4** (Marquardt): adaptive  $\lambda$       small  $\lambda \rightarrow$  Gauss-Newton, large  $\lambda \rightarrow$  gradient descend

1. choose  $\lambda \approx 10^{-3}$  and compute  $\mathbf{d}_s$
2. if  $\sum_i \|\mathbf{e}_i(\boldsymbol{\theta}^s + \mathbf{d}_s)\|^2 < \sum_i \|\mathbf{e}_i(\boldsymbol{\theta}^s)\|^2$  then accept  $\mathbf{d}_s$  and set  $\lambda := \lambda/10$ ,  $s := s + 1$
3. otherwise set  $\lambda := 10\lambda$  and recompute  $\mathbf{d}_s$

- sometimes different constants are needed for the 10 and  $10^{-3}$
- note that  $\mathbf{L}_i \in \mathbb{R}^{m,q}$  (long matrix) but each contribution  $\mathbf{L}_i^\top \mathbf{L}_i$  is a square singular  $q \times q$  matrix (always singular for  $k < q$ )
- $\lambda$  helps avoid the consequences of gauge freedom  $\rightarrow 143$
- error can be made robust to outliers  $\rightarrow 113$
- we have approximated the least squares Hessian by ignoring second derivatives of the error function (Gauss-Newton approximation) See [Triggs et al. 1999, Sec. 4.3]
- modern variants of LM are Trust Region methods

**Sampson** (derived by linearization over point coordinates  $u^1, v^1, u^2, v^2$ )

$$e_i(\mathbf{F}) = \frac{\varepsilon_i}{\|\mathbf{J}_i\|} = \frac{\underline{\mathbf{y}}_i^\top \mathbf{F} \underline{\mathbf{x}}_i}{\sqrt{\|\mathbf{S} \mathbf{F} \underline{\mathbf{x}}_i\|^2 + \|\mathbf{S} \mathbf{F}^\top \underline{\mathbf{y}}_i\|^2}} \quad \text{where} \quad \mathbf{S} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \end{bmatrix}$$

**LM** (by linearization over parameters  $\mathbf{F}$ )

$$\mathbf{L}_i = \frac{\partial e_i(\mathbf{F})}{\partial \mathbf{F}} = \dots = \frac{1}{2\|\mathbf{J}_i\|} \left[ \left( \underline{\mathbf{y}}_i - \frac{2e_i}{\|\mathbf{J}_i\|} \mathbf{S} \mathbf{F} \underline{\mathbf{x}}_i \right) \underline{\mathbf{x}}_i^\top + \underline{\mathbf{y}}_i \left( \underline{\mathbf{x}}_i - \frac{2e_i}{\|\mathbf{J}_i\|} \mathbf{S} \mathbf{F}^\top \underline{\mathbf{y}}_i \right)^\top \right] \quad (21)$$

- $\mathbf{L}_i$  in (21) is a  $3 \times 3$  matrix, must be reshaped to dimension-9 vector  $\text{vec}(\mathbf{L}_i)$  to be used in LM
- $\underline{\mathbf{x}}_i$  and  $\underline{\mathbf{y}}_i$  in Sampson error are normalized to unit homogeneous coordinate (21) relies on this
- reinforce rank  $\mathbf{F} = 2$  after each LM update to stay on the fundamental matrix manifold and  $\|\mathbf{F}\| = 1$  to avoid gauge freedom by SVD  $\rightarrow$  111
- LM linearization could be done by numerical differentiation (we have a small dimension here)

## ► Local Optimization for Fundamental Matrix Estimation

### Summary so far

- Given a set  $X = \{(x_i, y_i)\}_{i=1}^k$  of  $k \gg 7$  inlier correspondences, compute a statistically efficient estimate for fundamental matrix  $\mathbf{F}$ .
  - Find the conditioned ( $\rightarrow 92$ ) 7-point  $\mathbf{F}_0$  ( $\rightarrow 84$ ) from a suitable 7-tuple
  - Improve the  $\mathbf{F}_0^*$  using the LM optimization ( $\rightarrow 108-109$ ) and the Sampson error ( $\rightarrow 110$ ) on all inliers, reinforce rank-2, unit-norm  $\mathbf{F}_k^*$  after each LM iteration using SVD

### Partial conceptualization

- inlier = correspondence
- outlier = non-correspondence
- binary inlier/outlier labels are hidden
- we can get their likely estimate only, with respect to a model

### We are not yet done

- if there are no wrong correspondences (mismatches, outliers), this gives a local optimum given the 7-point initial estimate
- the algorithm breaks under contamination of (inlier) correspondences by outliers
- the full problem involves finding the inliers!
- in addition, we need a mechanism for jumping out of local minima (and exploring the space of all fundamental matrices)

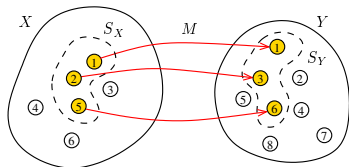
# ► The Full Problem of Matching and Fundamental Matrix Estimation

**Problem:** Given image keypoint sets  $X = \{x_i\}_{i=1}^m$  and  $Y = \{y_j\}_{j=1}^n$  and their descriptors  $D$ , find the most probable

1. inlier keypoints  $S_X \subseteq X$ ,  $S_Y \subseteq Y$
2. one-to-one perfect matching  $M: S_X \rightarrow S_Y$
3. fundamental matrix  $\mathbf{F}$  such that  $\text{rank } \mathbf{F} = 2$
4. such that for each  $x_i \in S_X$  and  $y_j = M(x_i)$  it is probable that
  - a) the image descriptor  $D(x_i)$  is similar to  $D(y_j)$ , and
  - b) the total reprojection error  $E = \sum_{ij} e_{ij}^2(\mathbf{F})$  is small
5. inlier-outlier and outlier-outlier matches are improbable

perfect matching: 1-factor of the bipartite graph

note a slight change in notation:  $e_{ij}$



$M:$

	Y							
X	1	2	3	4	5	6	7	8
1	1							
2			1					
3								
4								
5						1		
6								

□ = 0  
 ■ = 1 (matched)

$$(M^*, \mathbf{F}^*) = \arg \max_{M, \mathbf{F}} p(E, D, \mathbf{F} \mid M) P(M) \quad (22)$$

- probabilistic model: an efficient language for problem formulation it also unifies 4.a and 4.b
- the (22) is a Bayesian probabilistic model there is a constant number of random variables!
- binary matching table  $M_{ij} \in \{0, 1\}$  of fixed size  $m \times n$ 
  - each row/column contains at most one unity
  - zero rows/columns correspond to unmatched point  $x_i/y_j$

# Deriving A Robust Matching Model by Approximate Marginalization

For algorithmic efficiency, instead of  $(M^*, \mathbf{F}^*) = \arg \max_{M, \mathbf{F}} p(E, D, \mathbf{F} | M) P(M)$  solve

$$\mathbf{F}^* = \arg \max_{\mathbf{F}} p(E, D, \mathbf{F}) \quad (23)$$

by marginalization of  $p(E, D, \mathbf{F} | M) P(M)$  over  $\mathcal{M}$  s.t.  $M \in \mathcal{M}$  this changes the problem!  
drop the assumption that  $M$  is a 1:1 matching, assume correspondence-wise independence:

$$p(E, D, \mathbf{F} | M) P(M) = \prod_{i=1}^m \prod_{j=1}^n p_e(e_{ij}, d_{ij}, \mathbf{F} | m_{ij}) P(m_{ij})$$

- $e_{ij}$  represents (reprojection) error for match  $x_i \leftrightarrow y_j$ : e.g.  $e_{ij}(x_i, y_j, \mathbf{F})$
- $d_{ij}$  represents descriptor similarity for match  $x_i \leftrightarrow y_j$ : e.g.  $d_{ij} = \|\mathbf{d}(x_i) - \mathbf{d}(y_j)\|$

Approximate marginalization:

take all the  $2^{mn}$  terms in place of  $M$

$$\begin{aligned} p(E, D, \mathbf{F}) &\approx \sum_{m_{11} \in \{0,1\}} \sum_{m_{12}} \cdots \sum_{m_{mn}} p(E, D, \mathbf{F} | M) P(M) = \\ &= \sum_{m_{11}} \cdots \sum_{m_{mn}} \prod_{i=1}^m \prod_{j=1}^n p_e(e_{ij}, d_{ij}, \mathbf{F} | m_{ij}) P(m_{ij}) = \overset{*}{\dots} \overset{!}{=} \\ &= \prod_{i=1}^m \prod_{j=1}^n \underbrace{\sum_{m_{ij} \in \{0,1\}} p_e(e_{ij}, d_{ij}, \mathbf{F} | m_{ij}) P(m_{ij})}_{\text{we will continue with this term}} \end{aligned}$$

Thank You

