

Ch. I Markov Models on chains and acyclic graphs

1. Markov models on chains

1A. Definitions and basic properties

- Sequence $S = (S_1, \dots, S_n)$ of K -valued random variables $S_i \in K$
- K is a finite set, its elements are called states
- $p(s) = p(s_1, \dots, s_n)$ - joint probability distribution on K^n

w.l.o.g. we may write

$$\begin{aligned}
 p(s_1, \dots, s_n) &= p(s_n | s_1, \dots, s_{n-1}) p(s_1, \dots, s_{n-1}) \\
 &= \dots \\
 &= p(s_n | s_1, \dots, s_{n-1}) p(s_{n-1} | s_1, \dots, s_{n-2}) \dots p(s_2 | s_1) p(s_1)
 \end{aligned}$$

Definition 1a A p.d. on K^n is a Markov chain if

$$p(s) = p(s_1) \prod_{i=2}^n p(s_i | s_{i-1})$$

holds for $\forall s \in K^n$.

Definition 1b A p.d. on K^n is a Markov chain if

$$p(s) = \prod_{i=2}^n g_i(s_{i-1}, s_i)$$

holds for $\forall s \in K^n$, where $g_i: K^2 \rightarrow \mathbb{R}_+$ are some functions.

Equivalence:

a) \rightarrow b) trivial

b) \rightarrow a) recursively apply

$$P(S_{n-1}, S_n) = \left\{ \sum_{S_1, \dots, S_{n-2}} \prod_{i=2}^{n-1} g_i(S_{i-1}, S_i) \right\} g_n(S_{n-1}, S_n)$$

$\hookrightarrow g_n(S_{n-1}, S_n) = p(S_n | S_{n-1}) \cdot b_{n-1}(S_{n-1})$ with some b_{n-1}

Therefore

$$P(S_1, \dots, S_n) = \underbrace{\left[\prod_{i=2}^{n-1} g_i(S_{i-1}, S_i) \right]}_{P(S_1, \dots, S_{n-1})} b_{n-1}(S_{n-1}) \cdot p(S_n | S_{n-1})$$

Another useful formula

$$P(S_1, \dots, S_n) = \frac{P(S_1, S_2) \cdot p(S_2, S_3) \cdot \dots \cdot p(S_{n-1}, S_n)}{p(S_2) \cdot p(S_3) \cdot \dots \cdot p(S_{n-1})}$$

Example 1 (Ehrenfest model)

The model considers N particles in two containers. At each discrete time $t = 1, 2, \dots$, independently of the past, a particle is selected at random and moved to the other container. Let S_t denote the number of particles in the first container at time t . Then we have

$$P(S_t = k | S_{t-1} = l) = \begin{cases} \frac{N-l}{N} & \text{if } k = l+1 \\ \frac{l}{N} & \text{if } k = l-1 \\ 0 & \text{otherwise} \end{cases}$$

Q: How does $p(S_t = k)$, $k = 0, 1, \dots, N$ behaves for $t \rightarrow \infty$?

Example 2 (random walk on a graph)

Consider a random walk on an undirected graph V, E

- $K = V$ states, $S_t \in V$ - position of the walker at time t .
- $p(S_1)$ - some p.d. for the start vertex

$$p(s_t = i | s_{t-1} = j) = \begin{cases} W_{ij} & \text{if } \{i, j\} \in E \\ 0 & \text{otherwise} \end{cases}$$

where $W_{ij} \geq 0$ fulfill $\sum_{i \in \mathcal{N}(j)} W_{ij} = 1 \quad \forall j \in V$ \square

B. Homogeneous Markov chains, stationary p.d.s

A Markov chain is homogeneous if the conditional prob's $p(s_i | s_{i-1})$ do not depend on the position i , i.e.

$$p(s_i = k | s_{i-1} = k') = q(k, k') \quad \forall i = 2, \dots, n.$$

We know that

$$p(s_i = k) = \sum_{k' \in K} p(s_i = k | s_{i-1} = k') p(s_{i-1} = k').$$

Consider $p(s_i = k)$, $k \in K$ as components of an vector $\vec{\pi}_i \in \mathbb{R}_+^K$ and $p(s_i = k | s_{i-1} = k')$, $k, k' \in K$ as elements of a $K \times K$ matrix P . The previous eq. reads

$$\vec{\pi}_i = P \cdot \vec{\pi}_{i-1},$$

and, more general, we have $\vec{\pi}_i = P^{i-1} \vec{\pi}_1$. It may happen that there exists a p.d. $\vec{\pi}^*$ on K s.t. $P \cdot \vec{\pi}^* = \vec{\pi}^*$.

We call it a stationary p.d.

Definition 2 A homogeneous Markov chain is irreducible if for each pair $k, k' \in K$ there is an $m > 0$ s.t. $P_{kk'}^{(m)} > 0$. i.e., there is a non-zero probability to reach state k starting from state k' (after m transitions). \square

A condition somewhat stronger than irreducibility ensures existence and uniqueness of a stationary p.d.

Theorem 1 (w/o proof) If for some $m > 0$ all elements of the matrix P^m are strictly positive, then the Markov chain has a unique stationary distribution $\vec{\pi}^*$, which is a fixpoint

$$P^n \vec{\pi} \xrightarrow{n \rightarrow \infty} \vec{\pi}^* \quad \forall \vec{\pi}.$$

Moreover

$$P^n = \vec{\pi}^* \otimes \vec{e} + E(n),$$

where $\vec{e} = (1, \dots, 1)$ and $E_{kk'}(n) = O(h^n)$ with some $0 < h < 1$. ■

Remark 1 (infinite Markov chains)

- Consider infinite sequences $S = (s_1, s_2, \dots)$, $s_i \in K$. $K^{\mathbb{N}}$ is uncountable infinite. Any probability on $K^{\mathbb{N}}$ will assign zero probability to almost every sequence $S \in K^{\mathbb{N}}$
- A finite sequence $(k_1, k_2, \dots, k_n) \in K^n$ can be seen as a set of infinite sequences

$$(k_1, \dots, k_n) \mapsto \{S \in K^{\mathbb{N}} \mid s_1 = k_1, \dots, s_n = k_n\}.$$

A Markov model on $K^{\mathbb{N}}$ assigns probabilities to such sets in the same way as described for finite sequences.

C. Hidden Markov models on chains

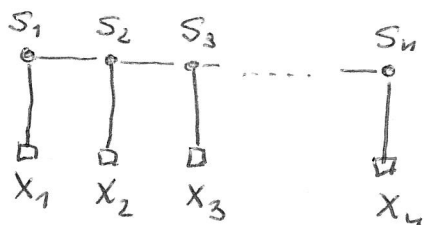
Common situation in pattern recognition:

$X = (X_1, \dots, X_n)$ sequence of features (observable)

$S = (S_1, \dots, S_n)$ sequence of states (hidden)

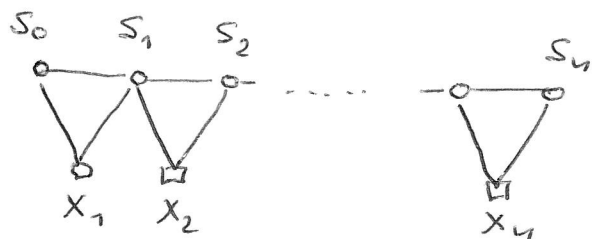
Hidden Markov model (HMM): a p.d. on pairs (x, s) s.t.

$$a) \quad p(x, s) = \underbrace{\prod_{i=1}^n p(x_i | s_i)}_{p(x|s)} \cdot p(s_1) \cdot \underbrace{\prod_{i=2}^n p(s_i | s_{i-1})}_{p(s) - \text{Markov model}}$$



b) or, slightly more general

$$p(x, s) = p(s_0) \prod_{i=1}^n p(x_i, s_i | s_{i-1})$$



(stochastic regular language!)