

RANSAC

Robust Model Estimation

From Data Contaminated By Outliers

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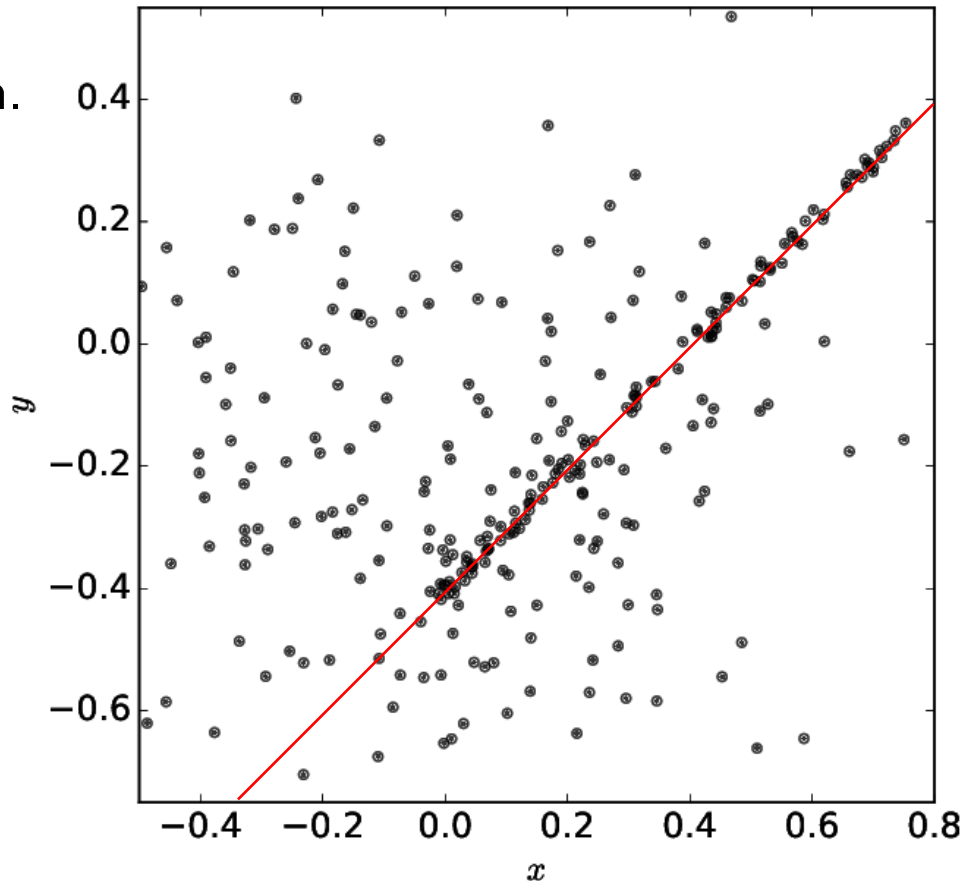
<http://cmp.felk.cvut.cz>

What is RANSAC?



- RANSAC = RANdom SAMple Consensus
- M.A. Fischler and R.C. Bolles. Random sample consensus: A paradigm for model fitting with applications to image analysis and automated cartography. CACM, 24(6):381–395, June 1981.

- **Example:** Finding a line in 2D data.
 - Not all input points are on a line.
 - Finding a line also implicitly divides points to **inliers** (=those on a line) and **outliers** (=those not on a line).



- Line parametrization (homogeneous):

$$ax + by + c = 0, \quad (a \neq 0 \vee b \neq 0) \quad (1)$$

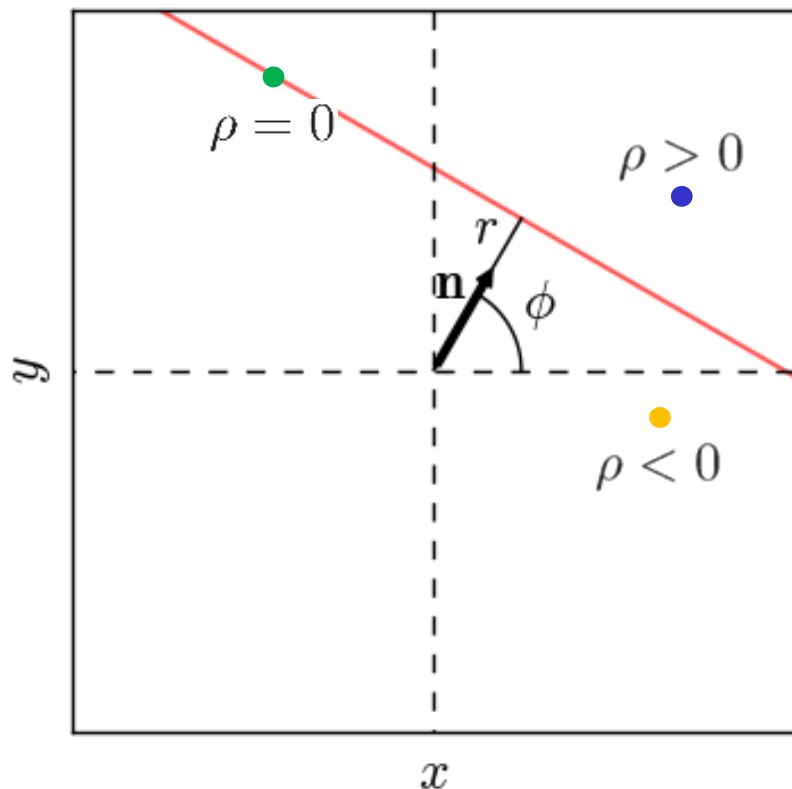
$$a, b, c \in \mathbb{R} : \text{line parameters} \quad (2)$$

$$(x, y) : \text{point coordinates} \quad (3)$$

- Line parametrization (radial):

$$x \cos \phi + y \sin \phi = r, \quad (4)$$

$$\phi \in [0, \pi[, r \in \mathbb{R} : \text{line parameters} \quad (5)$$



- Line parameters: $\phi \in [0, \pi[$, $r \in \mathbb{R}$

- Point $\mathbf{x} = (x, y)$ on the line:

$$x \cos \phi + y \sin \phi = r$$

$$\Leftrightarrow \mathbf{x} \cdot (\cos \phi, \sin \phi) = r$$

- Point $\mathbf{x} = (x, y)$ not on the line:

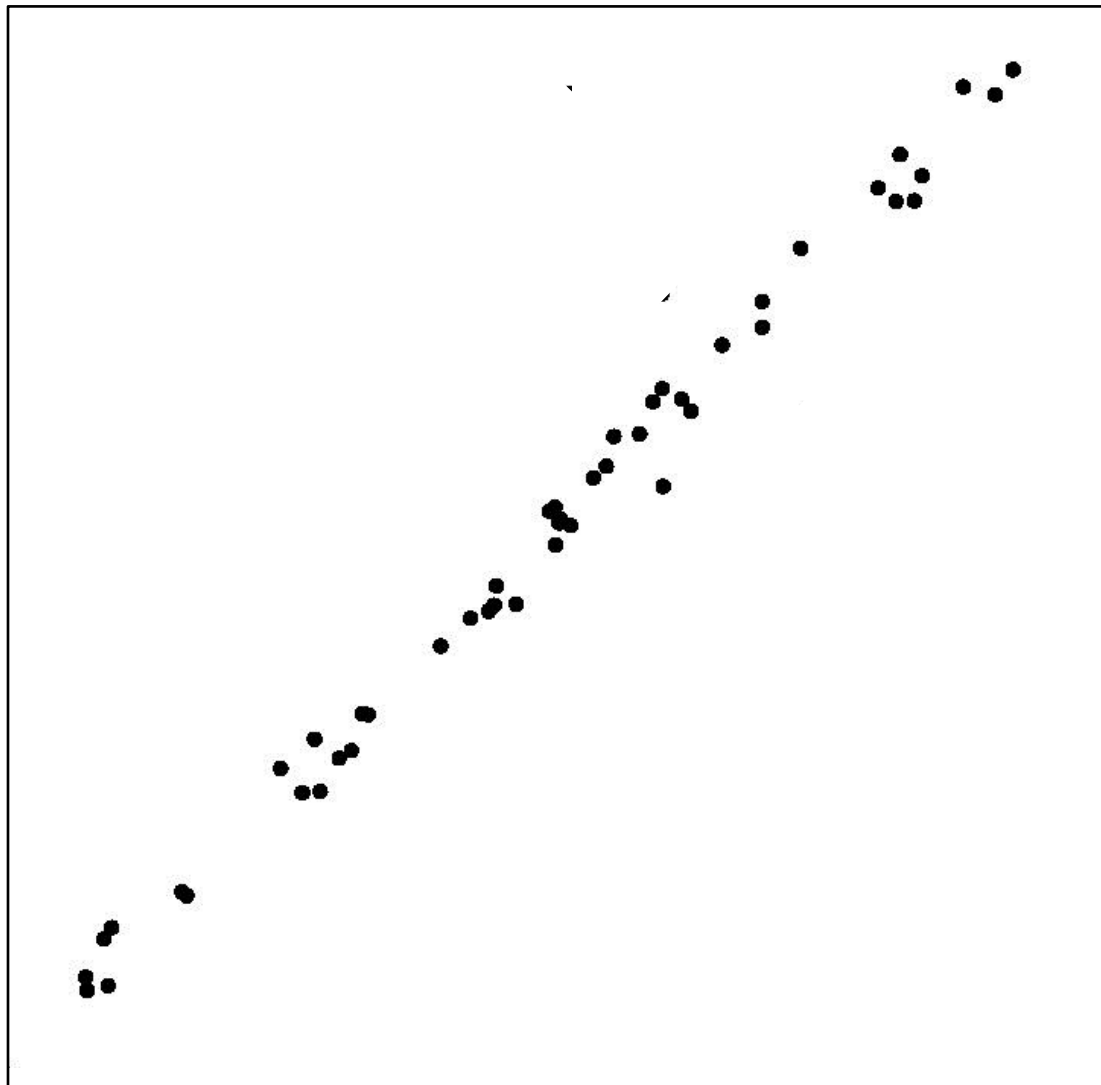
$$\mathbf{x} \cdot (\cos \phi, \sin \phi) \neq r$$

- Signed distance $\rho(\mathbf{x})$ from line:

$$\rho(\mathbf{x}) = \mathbf{x} \cdot (\cos \phi, \sin \phi) - r$$

Note: $\mathbf{n} = (\cos \phi, \sin \phi)$ (thus $\|\mathbf{n}\| = 1$)

Line Fitting, Inliers Only: Easy!



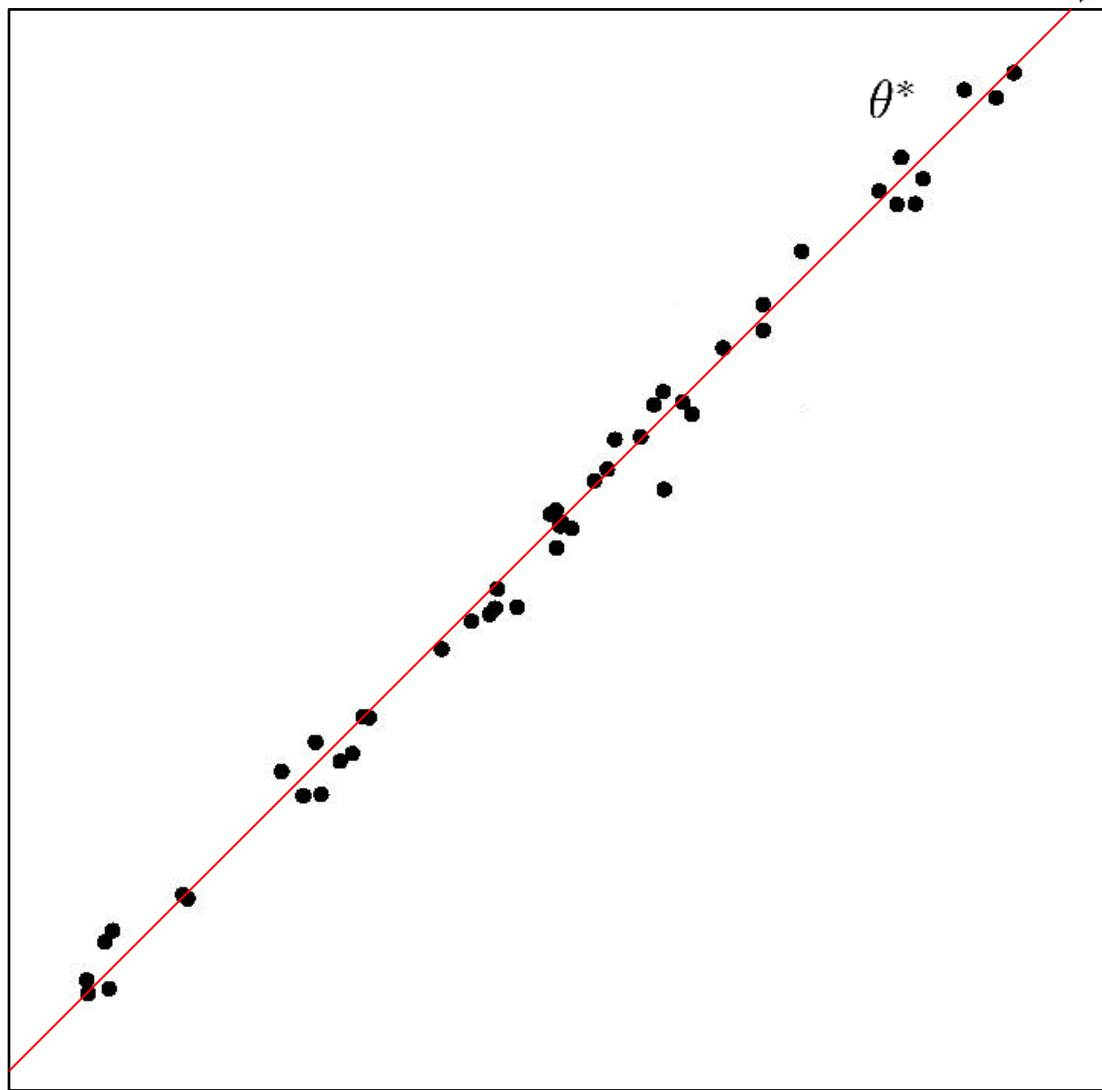
Data points

$$\mathcal{X} = \{\mathbf{x}_j, j = 1, 2, \dots, N_p\}$$
$$(\mathbf{x}_j \in \mathbb{R}^2)$$

Find the line which

“best fits” these points.

Line Fitting, Inliers Only: Easy!



Data points

$$\mathcal{X} = \{\mathbf{x}_j, j = 1, 2, \dots, N_p\}$$
$$(\mathbf{x}_j \in \mathbb{R}^2)$$

Find the line which
“best fits” the points.

As optimization: Find best
line with parameters θ^* as

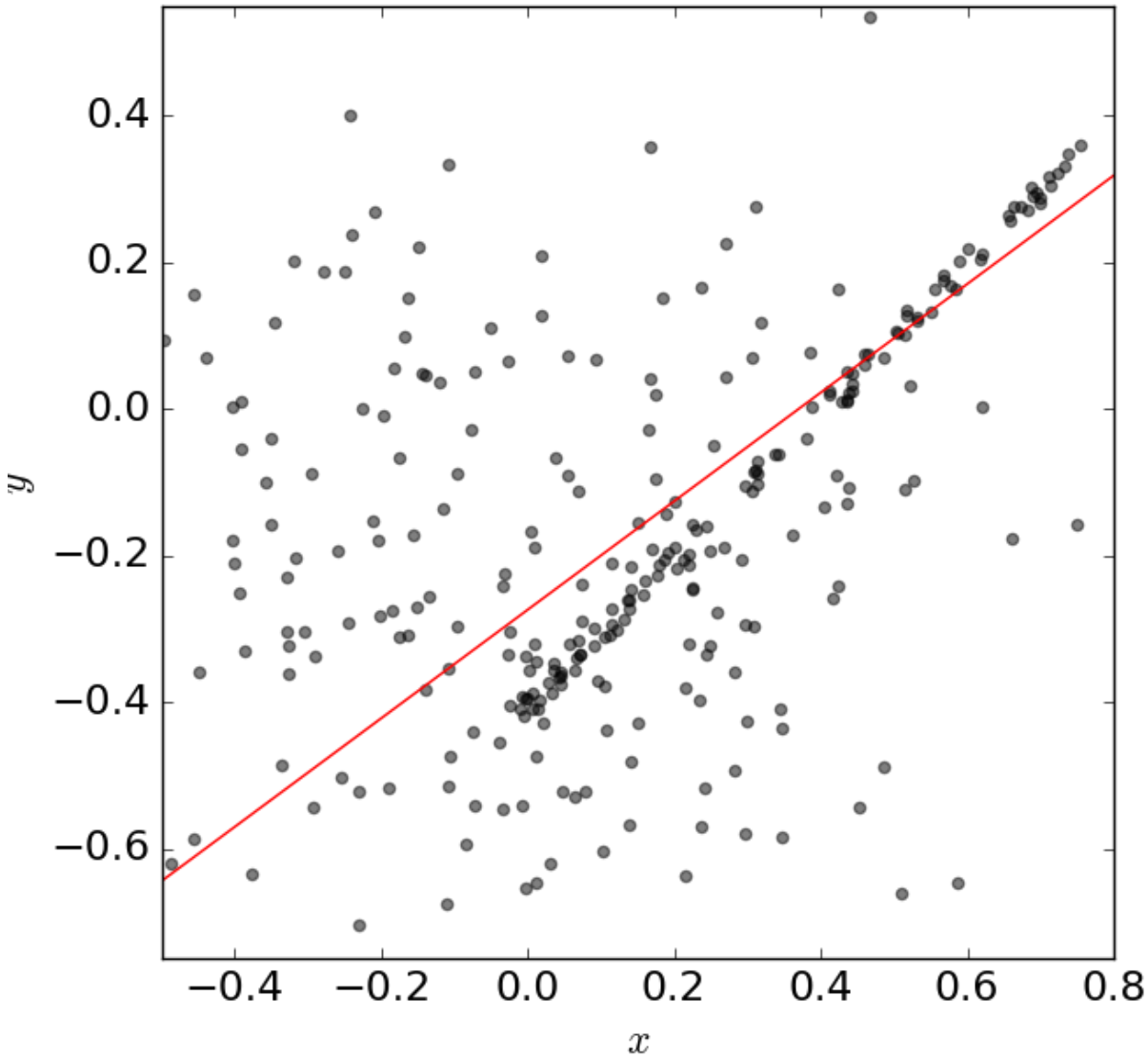
$$\theta^* = \operatorname{argmin}_{\theta} \sum_{\mathbf{x} \in \mathcal{X}} f(\mathbf{x}, \theta)$$

For $f_{LSQ}(\mathbf{x}, \theta) = [\rho(\mathbf{x})]^2$
this is easily solvable by
Singular Value
Decomposition (SVD).

General Case with Outliers, Example 1



Example 1

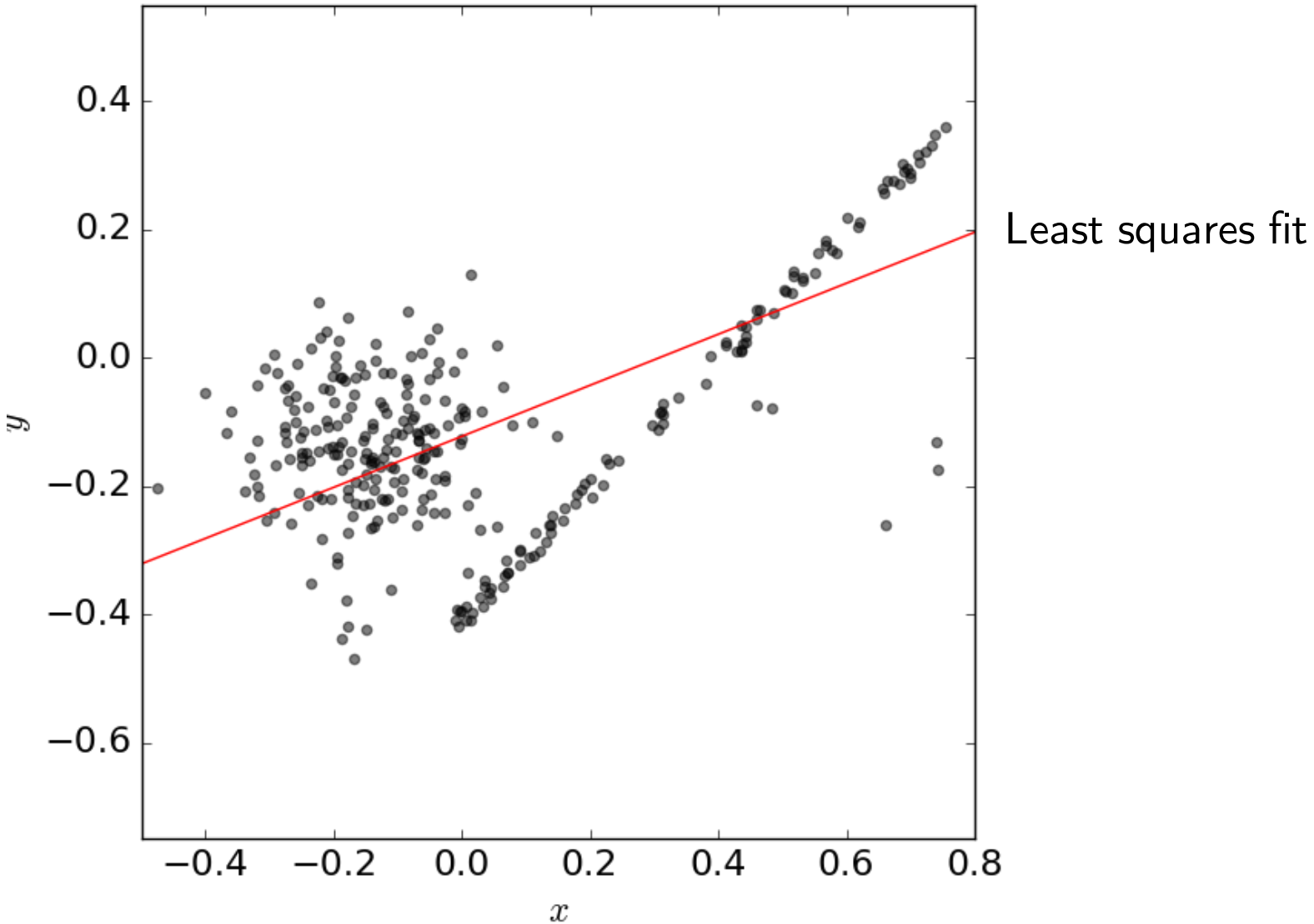


Least squares fit

General Case with Outliers, Example 2



Example 2



- $\mathcal{X} = \{\mathbf{x}_j\}_{j=1}^{N_p}$ set of data points

Find:

$$\theta^* = \arg \min_{\theta} \sum_{\mathbf{x} \in \mathcal{X}} f(\mathbf{x}, \theta)$$

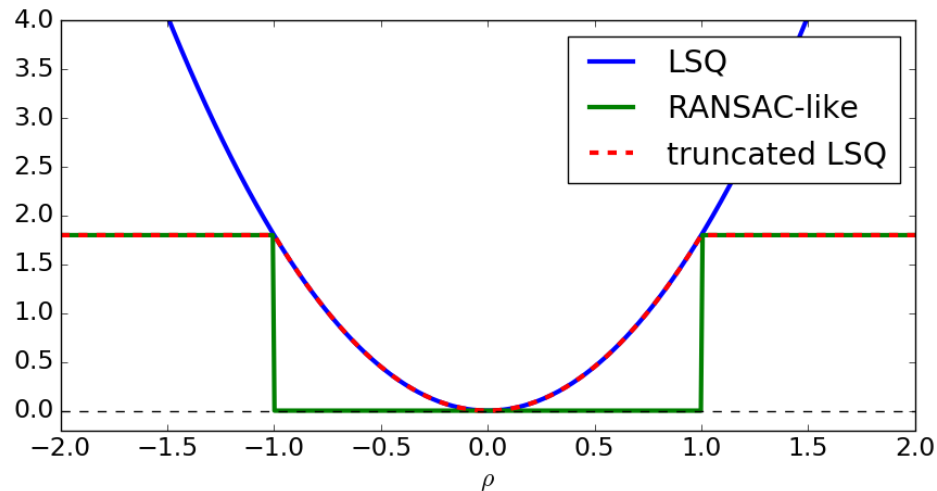
$$\theta = (r, \phi)$$

- No outliers: $f_{LSQ}(\mathbf{x}, \theta) = [\rho(\mathbf{x})]^2$

- **Use instead:**

$$f_{\text{RANSAC}}(\mathbf{x}, \theta) = \begin{cases} 0, & \text{if } \rho(\mathbf{x}) \leq \text{threshold } \sigma \\ \text{const}, & \text{otherwise} \end{cases}$$

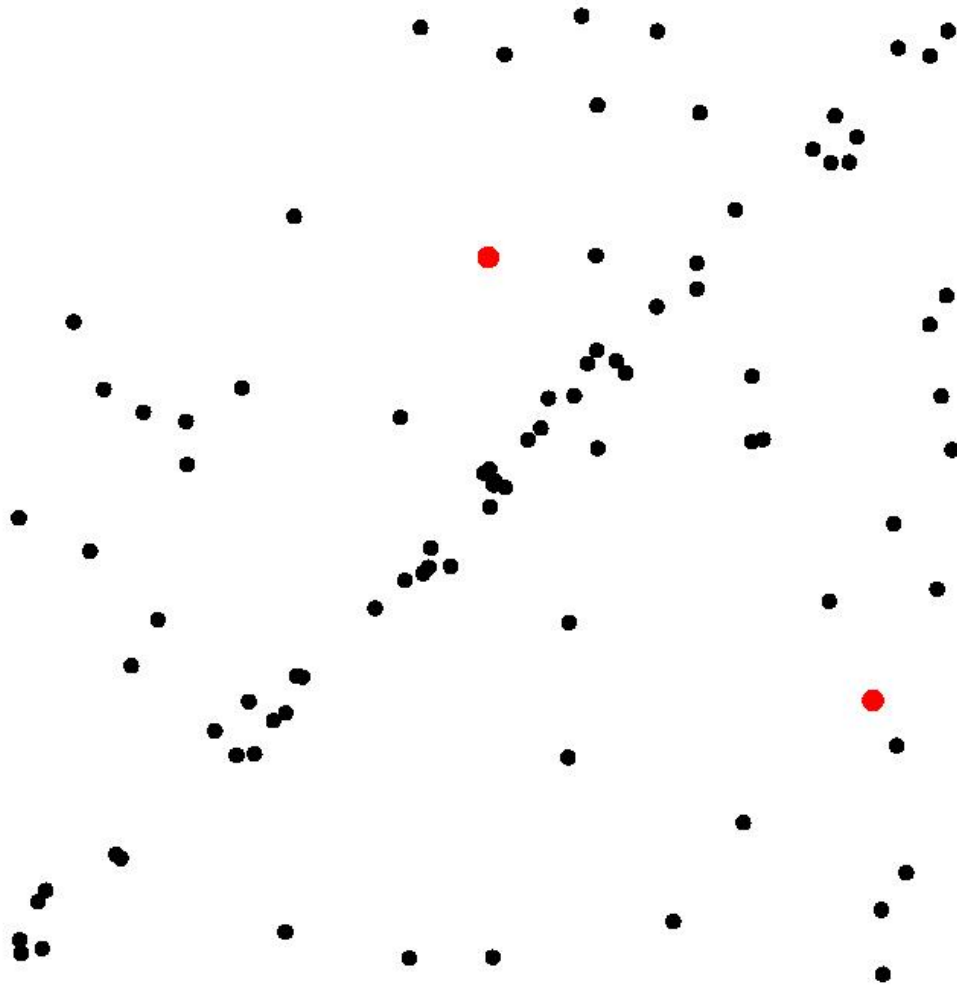
- Such cost function is non-convex
- How to find optimal line parameters?

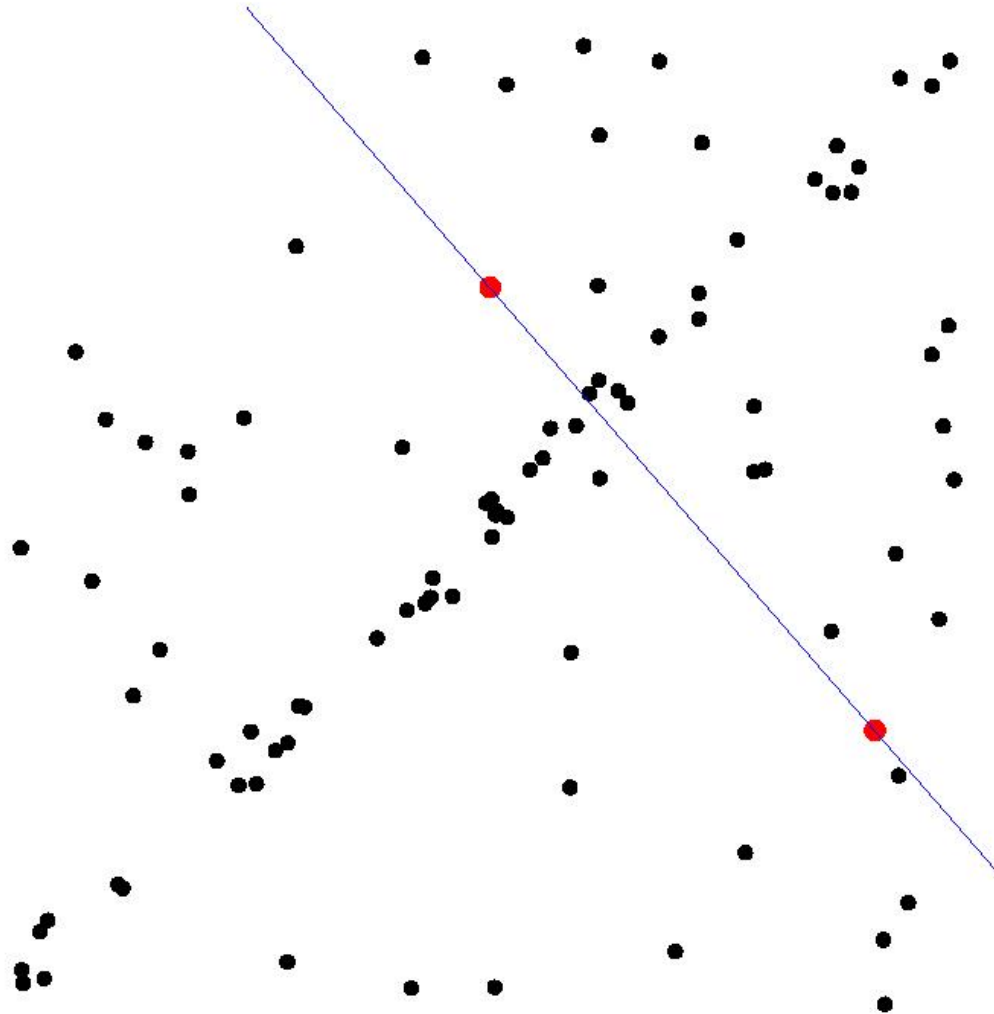


Random Sample Consensus - RANSAC



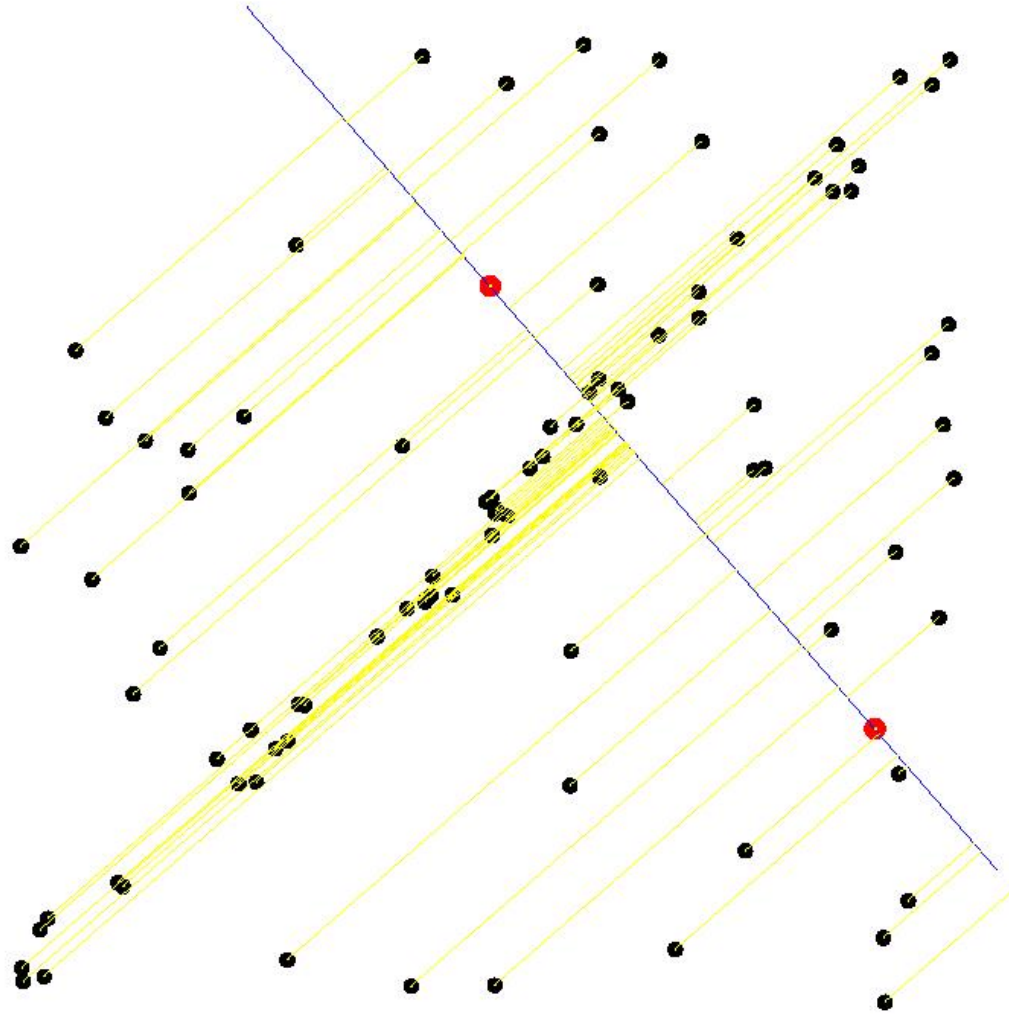
Select sample of m points
at random (here $m=2$)





Select sample of m points
at random

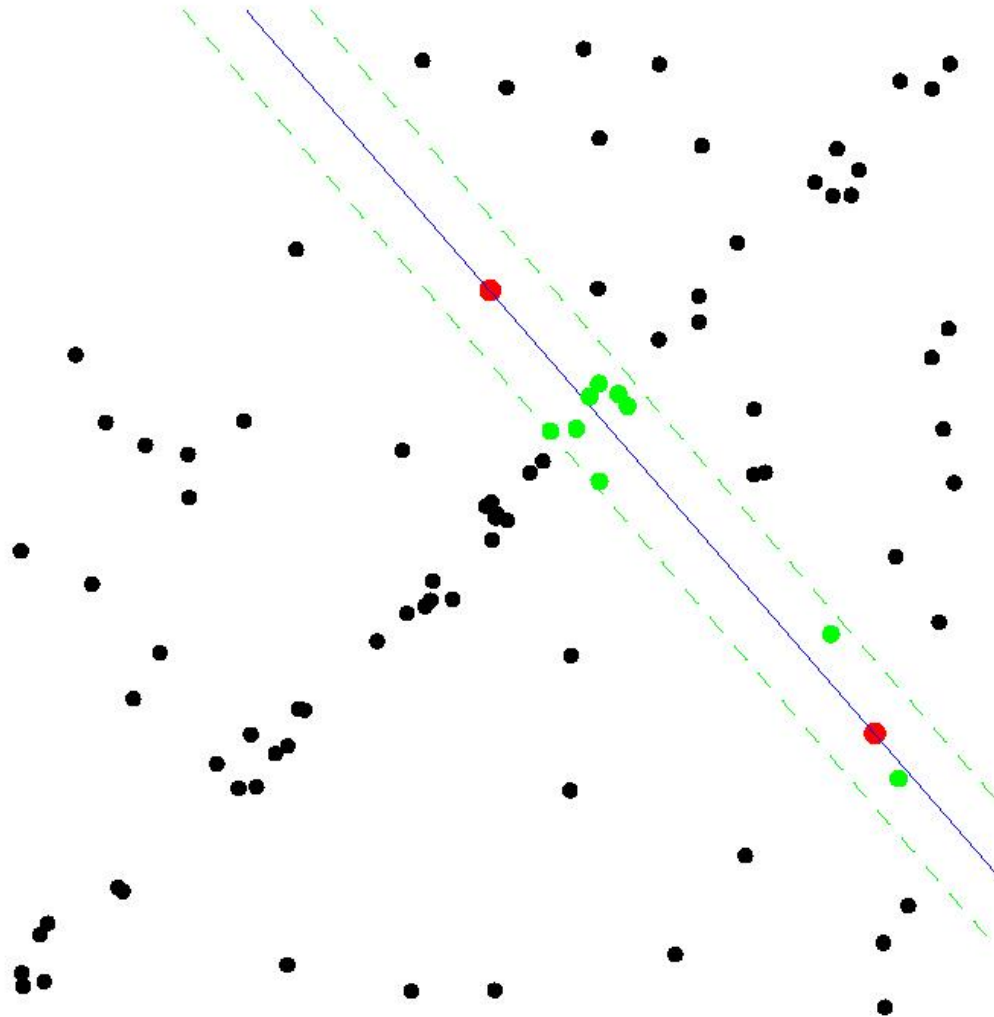
Estimate model parameters
from the data in the sample



Select sample of m points at random

Estimate model parameters from the data in the sample

Evaluate the error (residual) for each data point

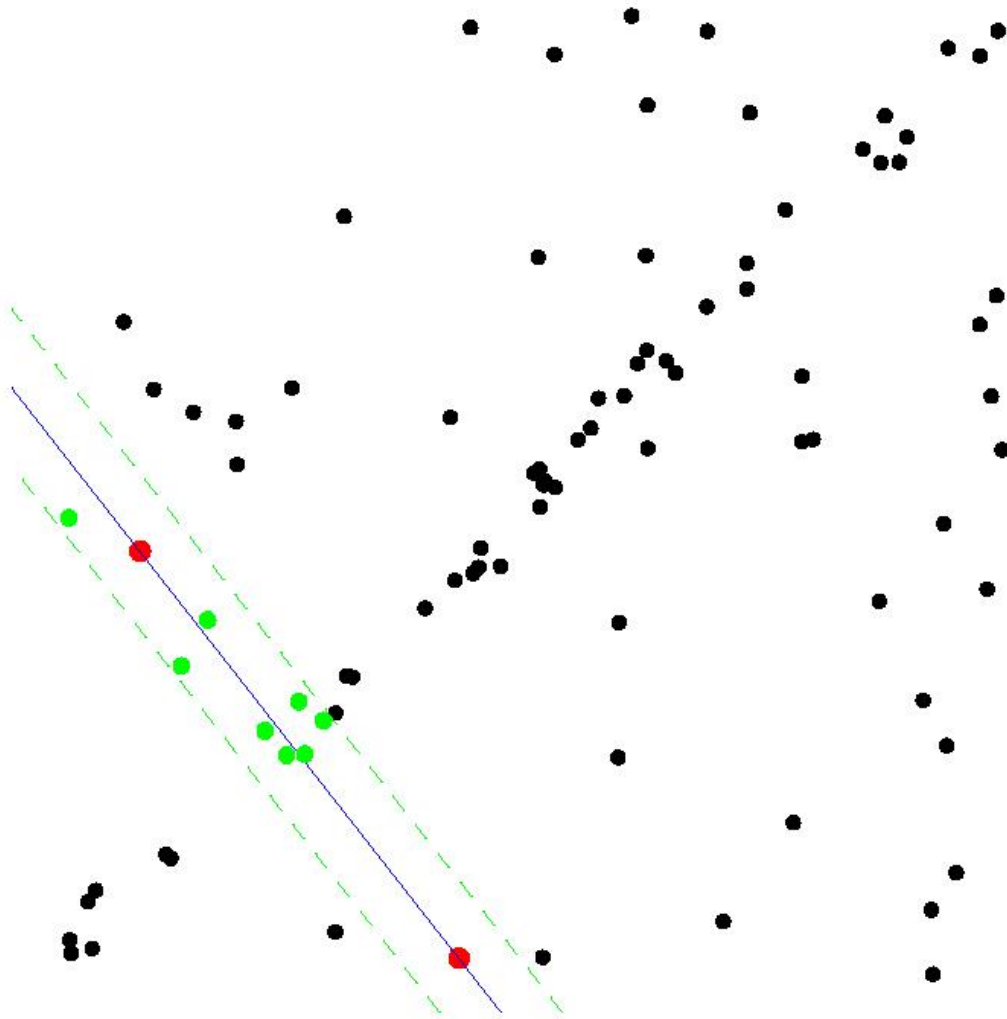


Select sample of m points at random

Estimate model parameters from the data in the sample

Evaluate the error (residual) for each data point

Select data that support the current hypothesis



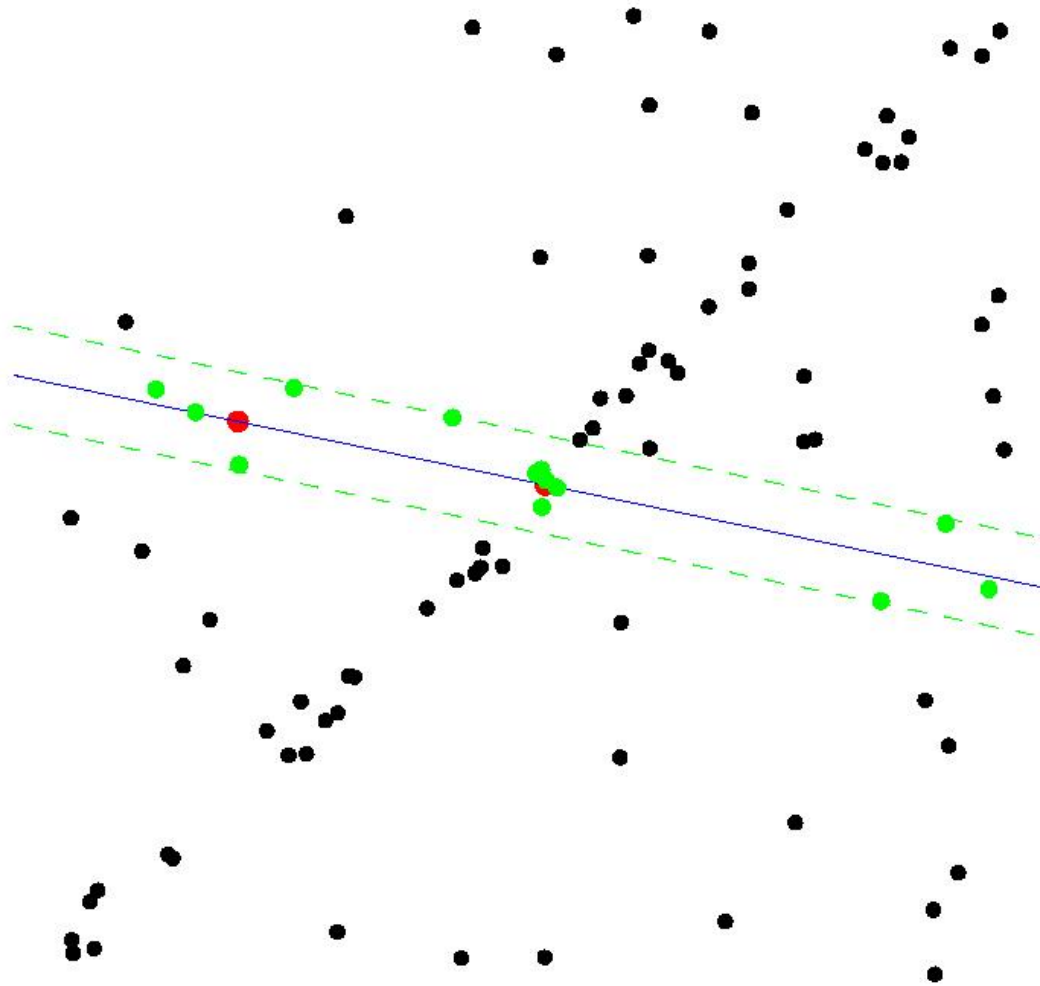
Select sample of m points at random

Estimate model parameters from the data in the sample

Evaluate the error (residual) for each data point

Select data that support the current hypothesis

Repeat sampling



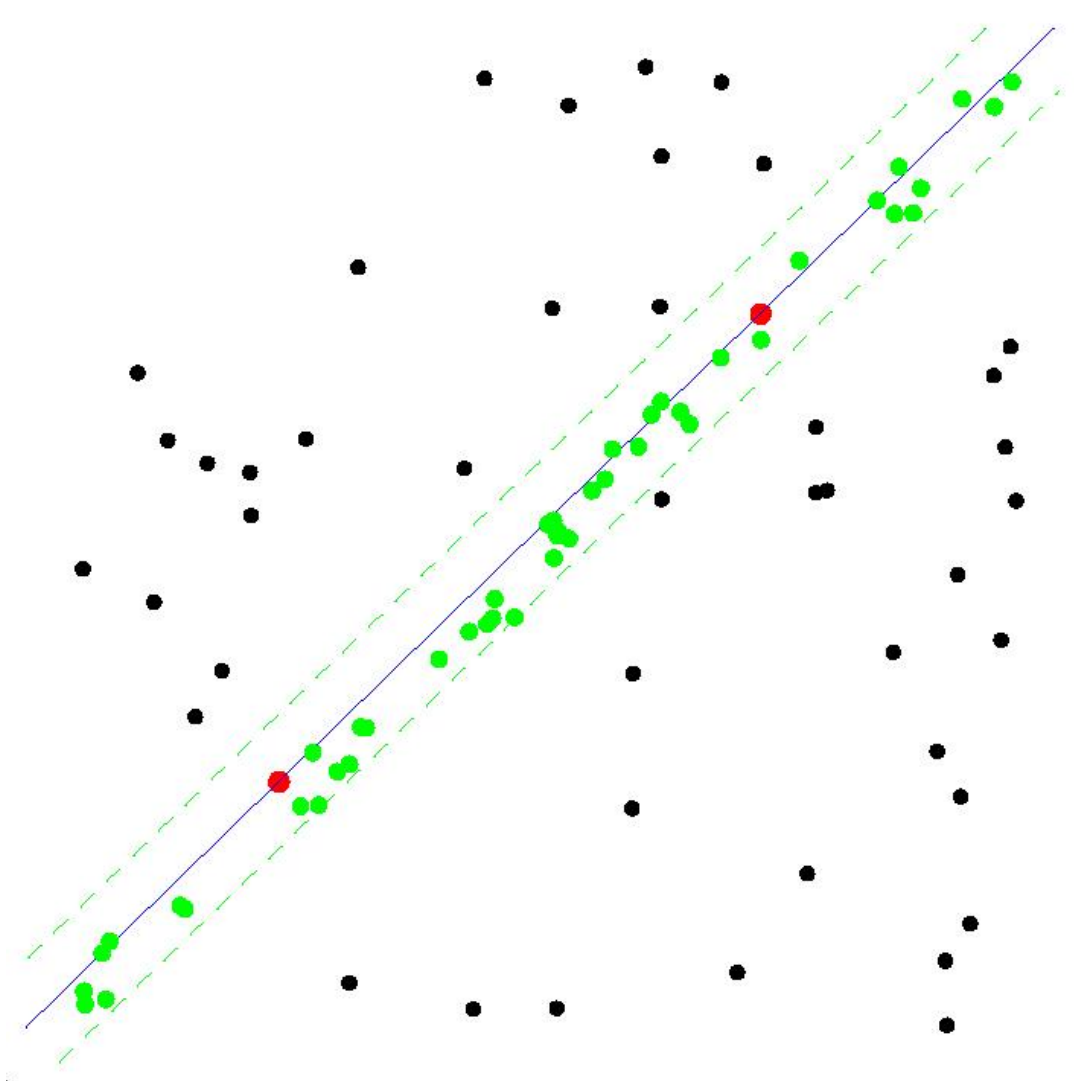
Select sample of m points at random

Estimate model parameters from the data in the sample

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Select sample of m points at random

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Select data that support the current hypothesis

Repeat sampling

Input: $\mathcal{X} = \{\mathbf{x}_j\}_{j=1}^N$ data points

$e(S) = \theta$ estimates *model parameters* θ given sample $S \subseteq \mathcal{X}$

$f(\mathbf{x}, \theta) = \begin{cases} 0, & \text{if distance to model} \leq \text{threshold } \sigma \\ 1, & \text{otherwise} \end{cases}$ Cost function for single data point \mathbf{x}

$\Rightarrow J(\theta) = \sum_{\mathbf{x} \in \mathcal{X}} f(\mathbf{x}, \theta)$ is #outliers

η – required confidence in the solution, σ – outlier threshold

Output: θ^* parameter of the model minimizing the cost function

1: $iter \leftarrow 0, J^* \leftarrow \infty$

2: **repeat**

3: Select *random* $S \subseteq \mathcal{X}$ (sample size $m = |S|$)

SAMPLING

4: Estimate parameters $\theta = e(S)$

5: Evaluate $J(\theta) = \sum_{\mathbf{x} \in \mathcal{X}} f(\mathbf{x}, \theta)$

VERIFICATION

6: If $J(\theta) < J^*$ then

SO-FAR-THE-BEST

$\theta^* \leftarrow \theta, J^* \leftarrow J(\theta)$

7: $iter \leftarrow iter + 1$

8: **until** $P(\text{better solution exists}) = f(|\mathcal{X}|, J^*, iter) < \eta$

9: Compute θ^* from all inliers \mathcal{X}_{in} : $\theta^* \leftarrow \text{LocalOptimization}(\mathcal{X}_{in}, \theta^*)$

RANSAC – how many samples?



- N Number of points
- Q Number of inliers, $Q = N - J^*$
- m Size of sample
- $\epsilon = Q/N$ Inlier ratio

Probability of all-inlier (uncontaminated) sample:

$$P(\text{inlier sample}) = \frac{\binom{Q}{m}}{\binom{N}{m}} = \frac{Q(Q-1)\dots(Q-m+1)}{N(N-1)\dots(N-m+1)} \approx \epsilon^m$$

Mean time for hitting all-inliers sample is proportional to $1/P$.

- How about this formulation:
 - Set the number of samples k such that **at least one** pair of points from the line has been hit with probability larger than η
 - Equivalently ... such that **no** pair of points from the line has been hit with probability lower than $1 - \eta$
- Q Number of inliers, $Q = N - J^*$
- m Size of sample
- $\epsilon = Q/N$ Inlier ratio

Probability of all-inlier (uncontaminated) sample:

$$P(\text{inlier sample}) = \frac{\binom{Q}{m}}{\binom{N}{m}} = \frac{Q(Q-1)\dots(Q-m+1)}{N(N-1)\dots(N-m+1)} \approx \epsilon^m$$

We require:

$$P(\text{bad pair } k \text{ times}) = (1 - P(\text{inlier sample}))^k < 1 - \eta$$

Finding the solution with confidence η therefore requires at least:

$$k \geq \log(1 - \eta) / \log(1 - \epsilon^m)$$

RANSAC termination - How many samples?



Inlier ratio $\epsilon = Q/N$ [%]

Size of the sample m

	15%	20%	30%	40%	50%	70%
2	132	73	32	17	10	4
4	5916	1871	368	116	46	11
7	$1.75 \cdot 10^6$	$2.34 \cdot 10^5$	$1.37 \cdot 10^4$	1827	382	35
8	$1.17 \cdot 10^7$	$1.17 \cdot 10^6$	$4.57 \cdot 10^4$	4570	765	50
12	$2.31 \cdot 10^{10}$	$7.31 \cdot 10^8$	$5.64 \cdot 10^6$	$1.79 \cdot 10^5$	$1.23 \cdot 10^4$	215
18	$2.08 \cdot 10^{15}$	$1.14 \cdot 10^{13}$	$7.73 \cdot 10^9$	$4.36 \cdot 10^7$	$7.85 \cdot 10^5$	1838
30	∞	∞	$1.35 \cdot 10^{16}$	$2.60 \cdot 10^{12}$	$3.22 \cdot 10^9$	$1.33 \cdot 10^5$
40	∞	∞	∞	$2.70 \cdot 10^{16}$	$3.29 \cdot 10^{12}$	$4.71 \cdot 10^6$

computed for $\eta = 0.95$

Pros:

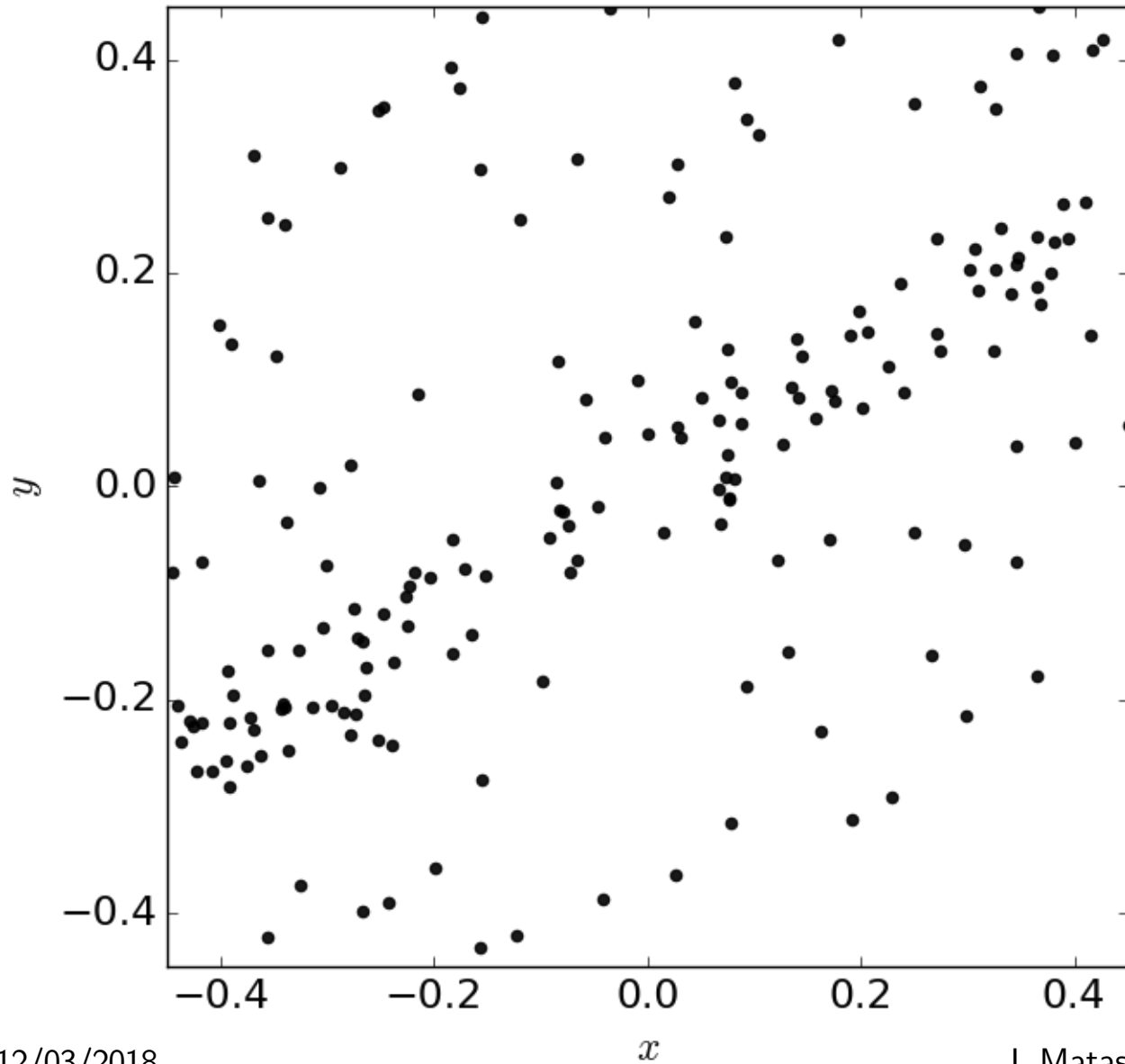
- extremely popular (>17000 citations in Google Scholar)
- used in many applications
- percentage of inliers not needed and not limited
- a probabilistic guarantee for the solution
- mild assumptions: σ known

Cons:

- slow if inlier ratio low
- It was observed experimentally that RANSAC takes several times longer than theoretically expected. This is due to noise – not every all-inlier sample generates a good hypothesis:

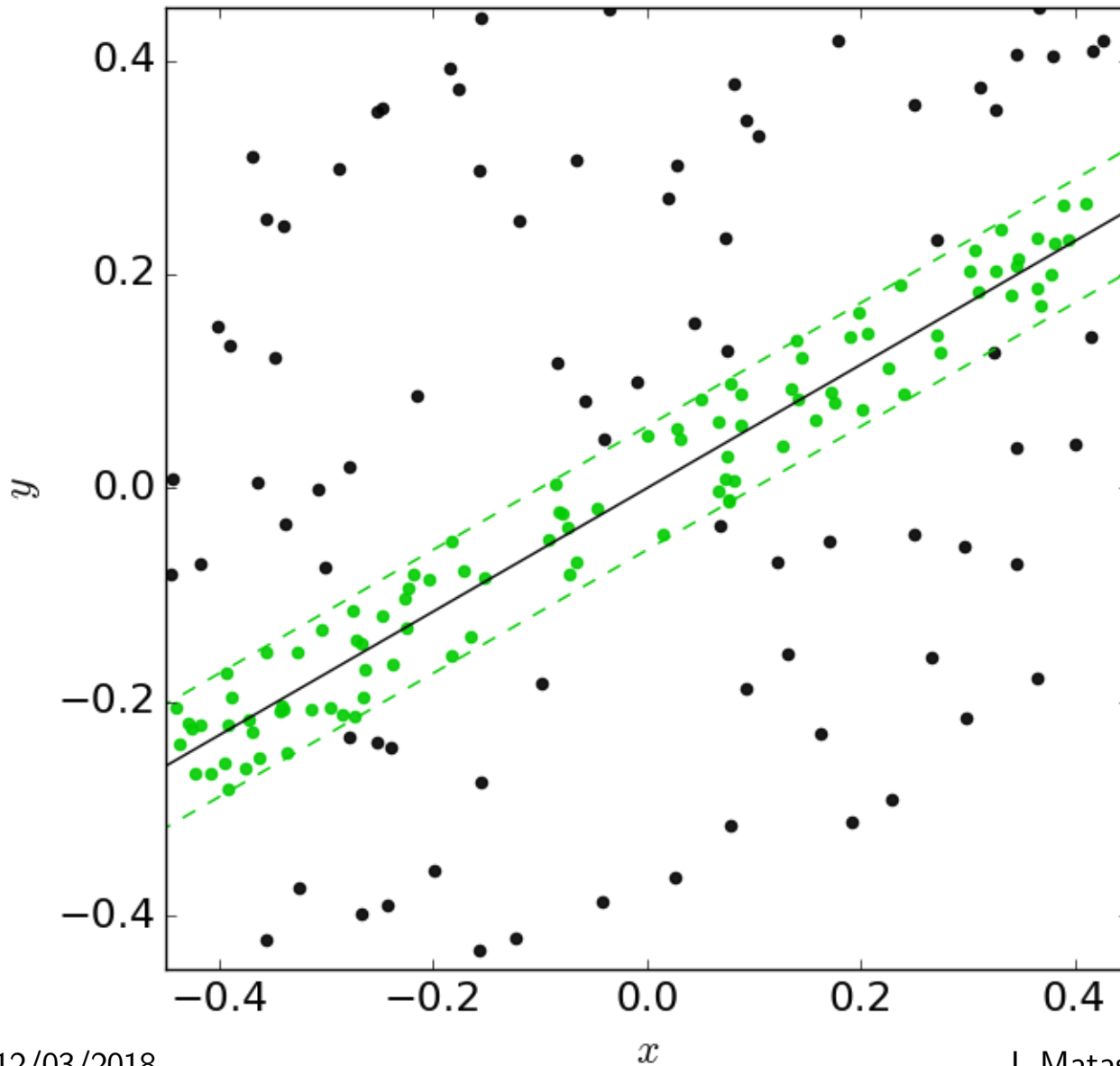
$$P(\text{inlier sample}) \neq P(\text{good model estimate})$$

- **Cost function:** MLESAC, Huber loss, ...
- **Outlier threshold σ :** Least median of Squares, MINPRAN, ...
- **Correctness of the results. Degeneracy.**
Solution: DegenSAC.
- **Accuracy** (parameters are estimated from minimal samples).
Solution: Locally Optimized RANSAC
- **Speed:** Running time grows with
 1. number of data points,
 2. number of iterations (polynomial in the inlier ratio)Addressing the problem:
RANSAC with SPRT (WaldSAC), PROSAC



Data: 200 points

LO-RANSAC: Problem Introduction

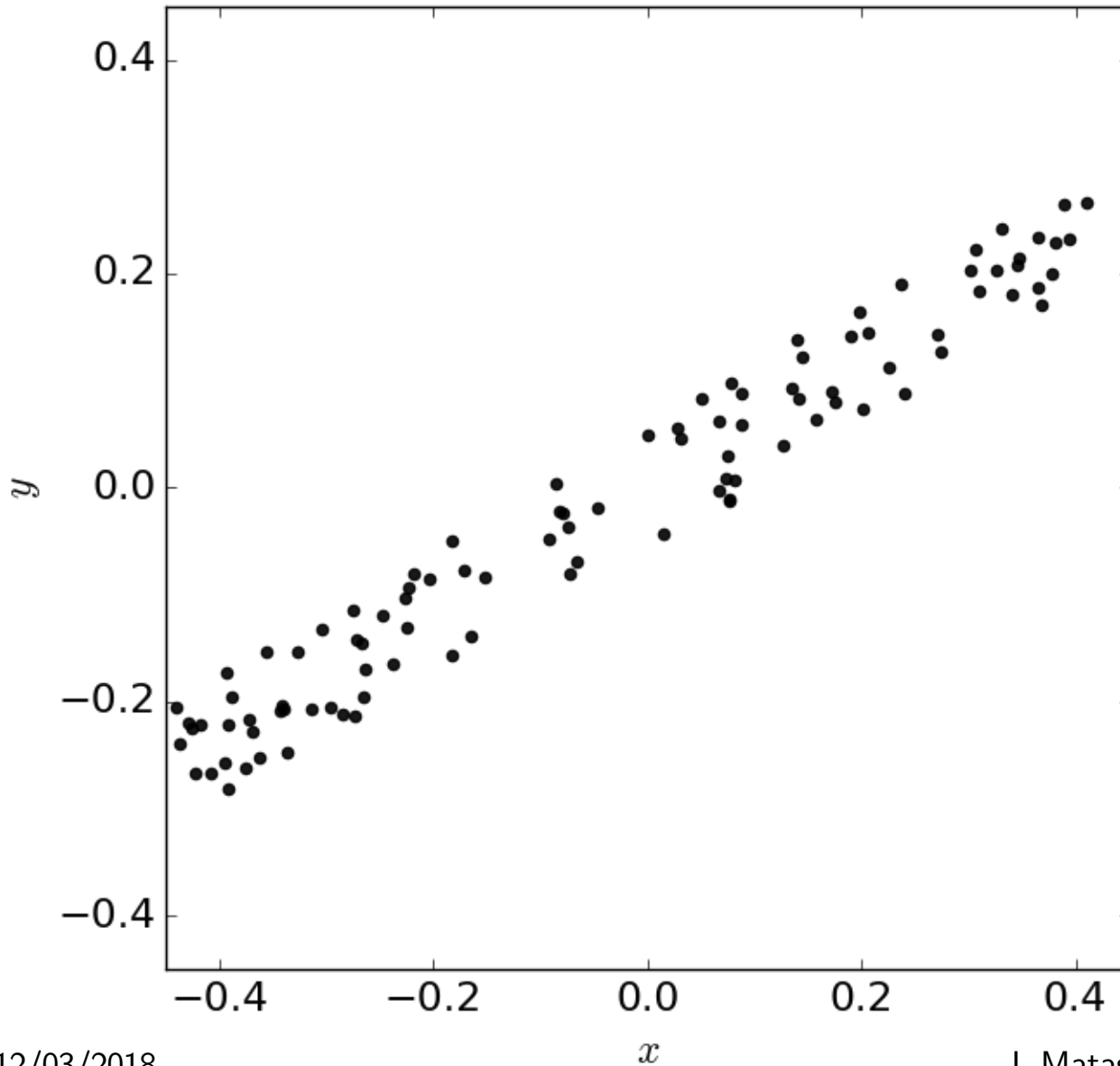


Data: 200 points
Model, 100 inliers

LO-RANSAC: Problem Introduction



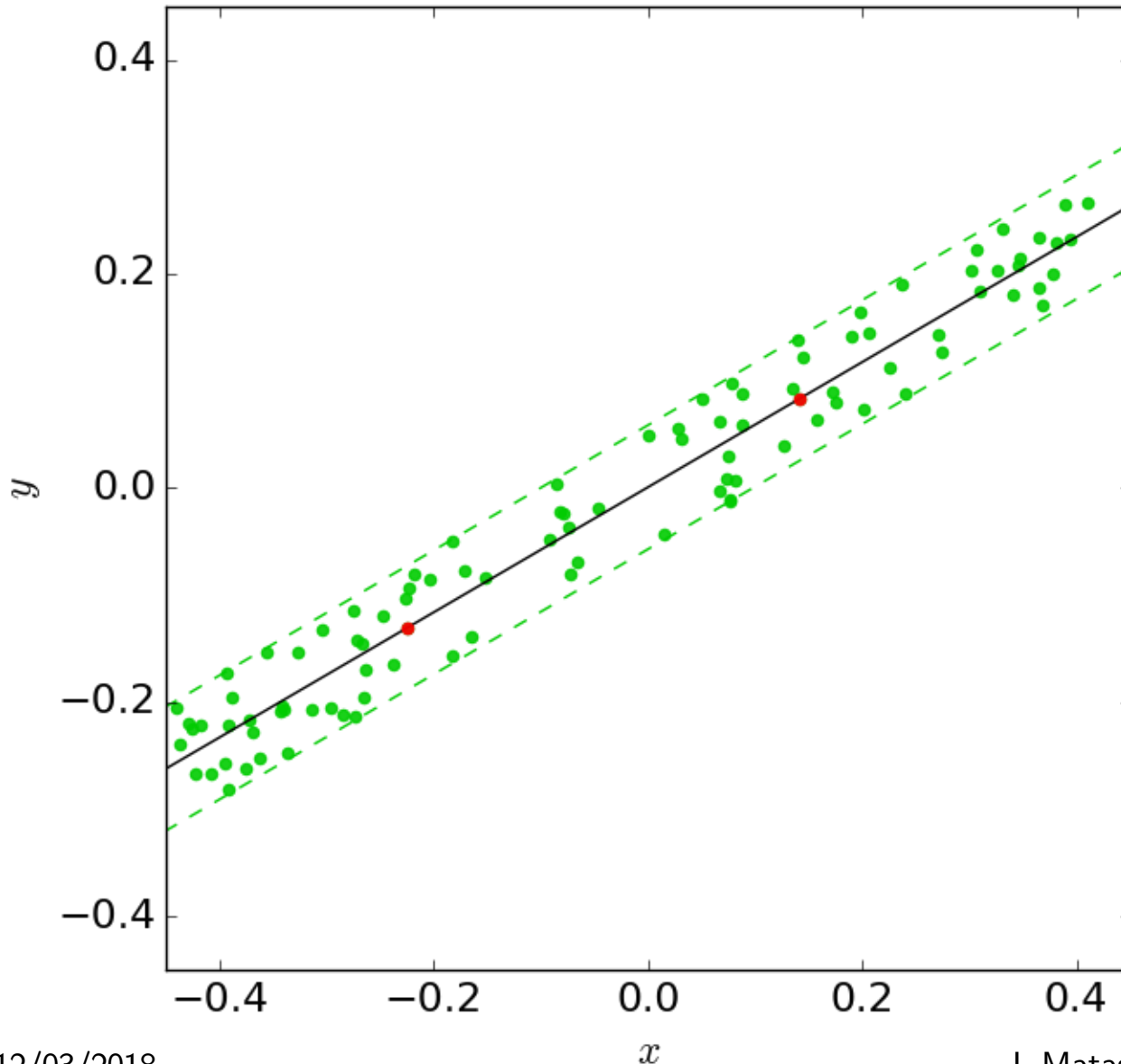
For simplicity, consider only points belonging to the model (100 points)



LO-RANSAC: Problem Introduction



For simplicity, consider only points belonging to the model (100 points)



RANSAC

Hypothesis generation
from 2 points

**Will every two
points generate the
whole inlier set?**

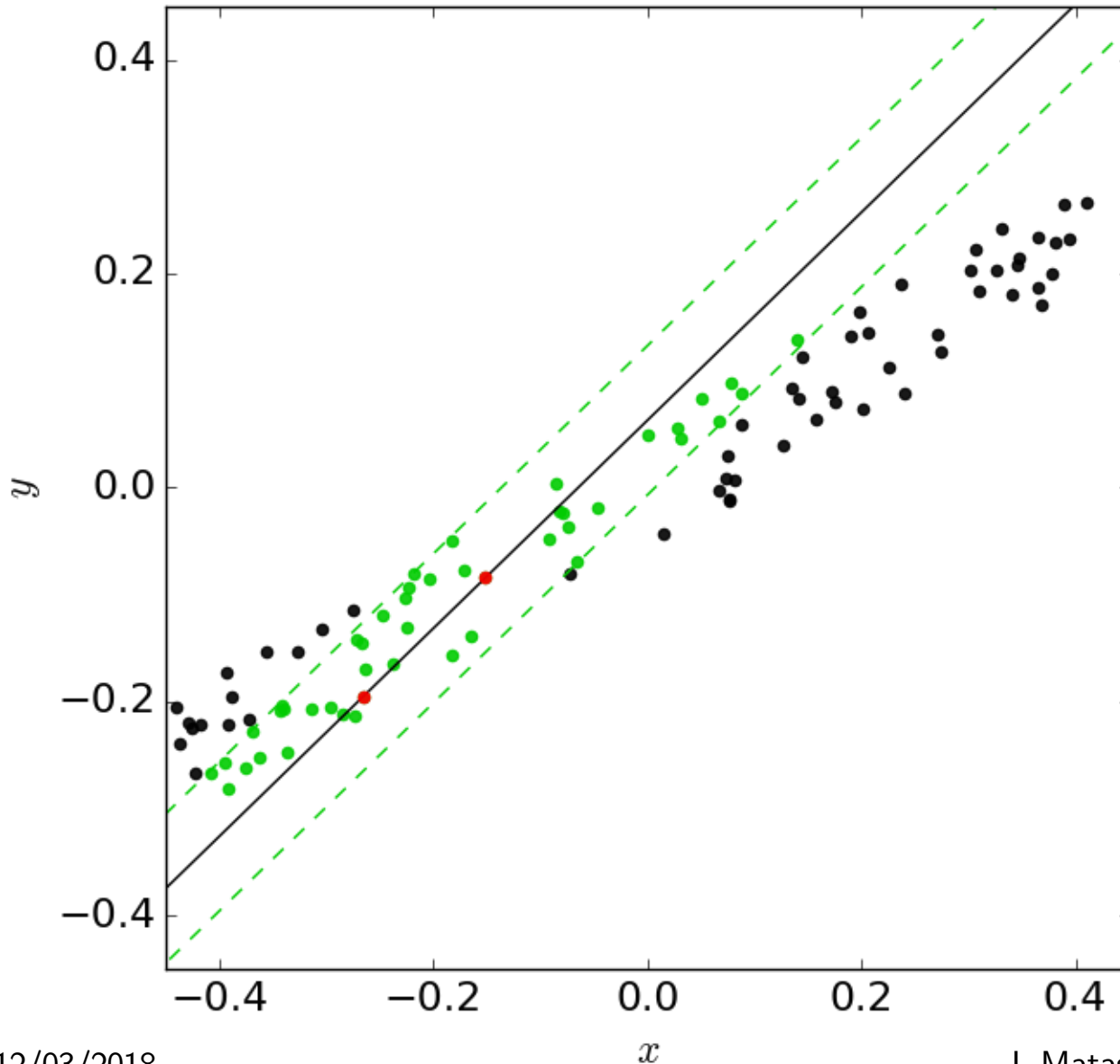
This sample:

YES. 100 inliers.

LO-RANSAC: Problem Introduction



For simplicity, consider only points belonging to the model (100 points)



RANSAC

Hypothesis generation
from 2 points

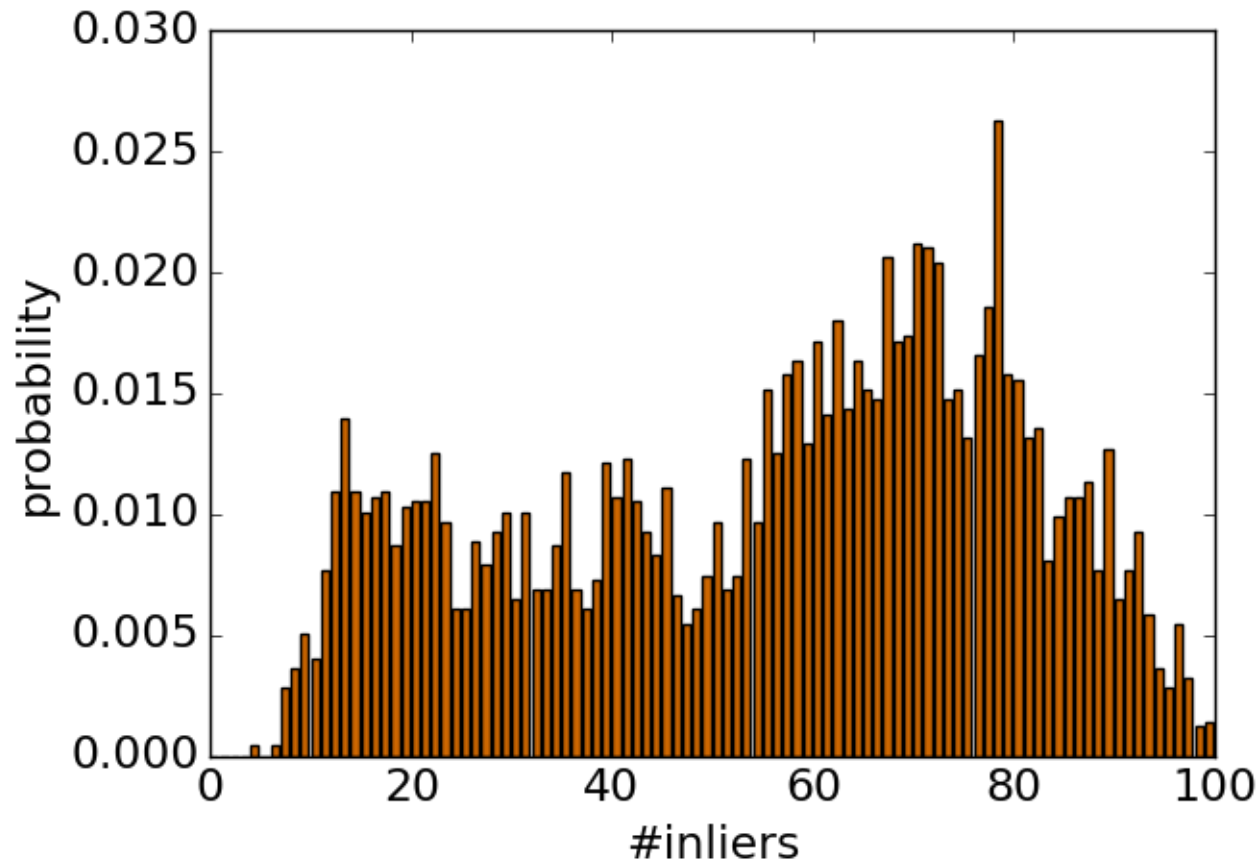
**Will every two
points generate the
whole inlier set?**

This sample:
NO. 45 inliers.

LO-RANSAC: Problem Introduction



For simplicity, consider only points belonging to model (100 points)



RANSAC

Hypothesis generation
from 2 points

**Will every two
points generate the
whole inlier set?**

The distribution of the number of inliers
obtained while randomly sampling points pairs

Input: $\mathcal{X} = \{\mathbf{x}_j\}_{j=1}^N$ data points

$e(S) = \theta$ estimates *model parameters* θ given sample $S \subseteq \mathcal{X}$

$f(\mathbf{x}, \theta) = \begin{cases} 0, & \text{if distance to model} \leq \text{threshold } \sigma \\ 1, & \text{otherwise} \end{cases}$ Cost function for single data point \mathbf{x}

$\Rightarrow J(\theta) = \sum_{\mathbf{x} \in \mathcal{X}} f(\mathbf{x}, \theta)$ is #outliers

η – required confidence in the solution, σ – outlier threshold

Output: θ^* parameter of the model minimizing the cost function

1: $iter \leftarrow 0, J^* \leftarrow \infty$

2: **repeat**

3: Select *random* $S \subseteq \mathcal{X}$ (sample size $m = |S|$)

SAMPLING

4: Estimate parameters $\theta = e(S)$

5: Evaluate $J(\theta) = \sum_{\mathbf{x} \in \mathcal{X}} f(\mathbf{x}, \theta)$

VERIFICATION
SO-FAR-THE-BEST

6: If $J(\theta) < J^*$ then
 $\theta^* \leftarrow \theta, J^* \leftarrow J(\theta)$

7: $iter \leftarrow iter + 1$

8: **until** $P(\text{better solution exists}) = f(|\mathcal{X}|, J^*, iter) < \eta$

9: Compute θ^* from all inliers \mathcal{X}_{in} : $\theta^* \leftarrow \text{LocalOptimization}(\mathcal{X}_{in}, \theta^*)$

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η – required confidence in the solution, σ – outlier threshold

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1: $iter \leftarrow 0, J^* \leftarrow \infty$

2: **repeat**

3: Select *random* $S \subseteq \mathcal{X}$ (sample size $m = |S|$)

SAMPLING

4: Estimate parameters $\theta = e(S)$

5: Evaluate $J(\theta) = \sum_{\mathbf{x} \in \mathcal{X}} f(\mathbf{x}, \theta)$

VERIFICATION

6: If $J(\theta) < J^*$ then

SO-FAR-THE-BEST

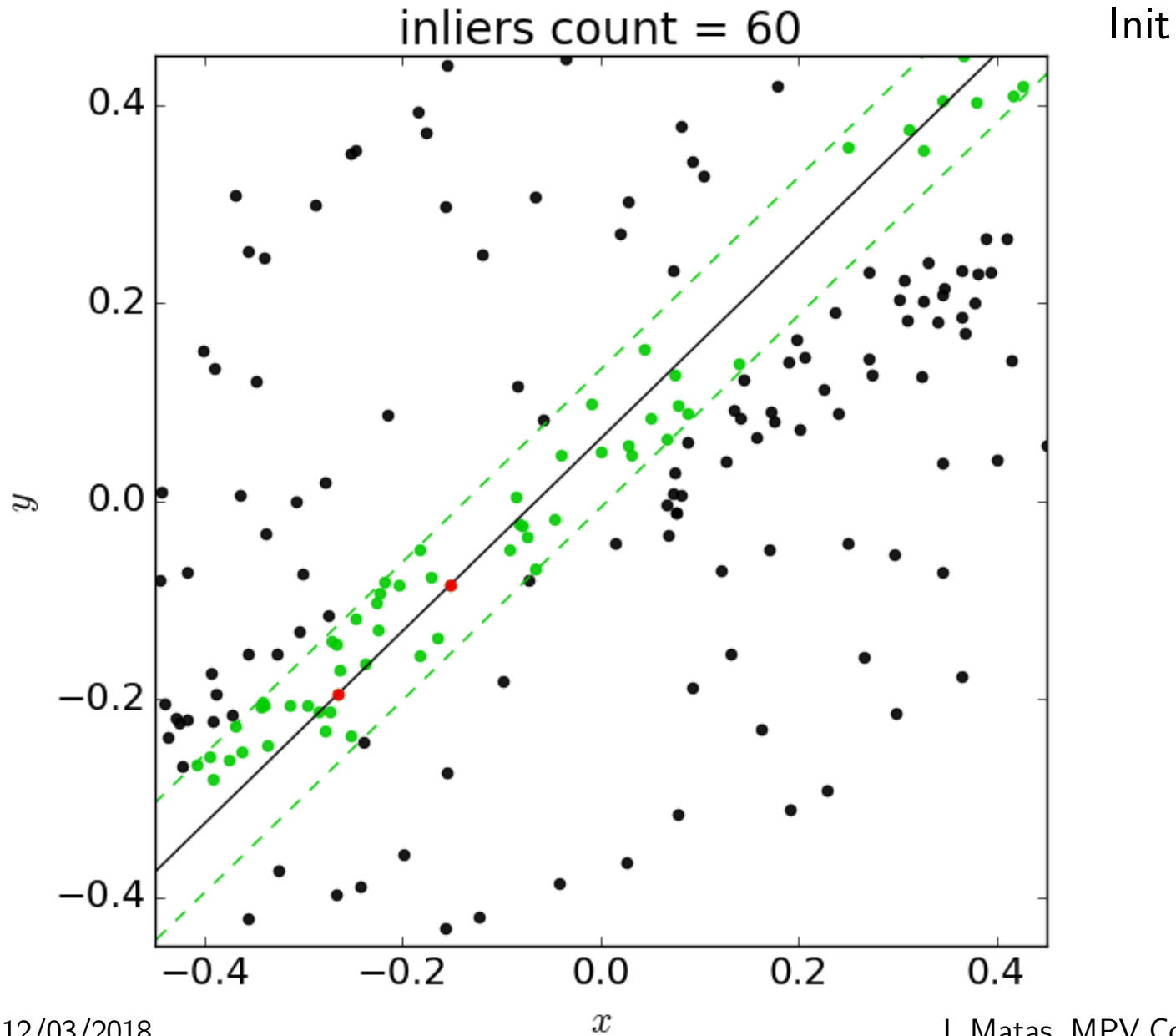
$\theta^* \leftarrow \text{LocalOptimization}(\mathcal{X}_{in}, \theta), J^* \leftarrow J(\theta^*)$

7: $iter \leftarrow iter + 1$

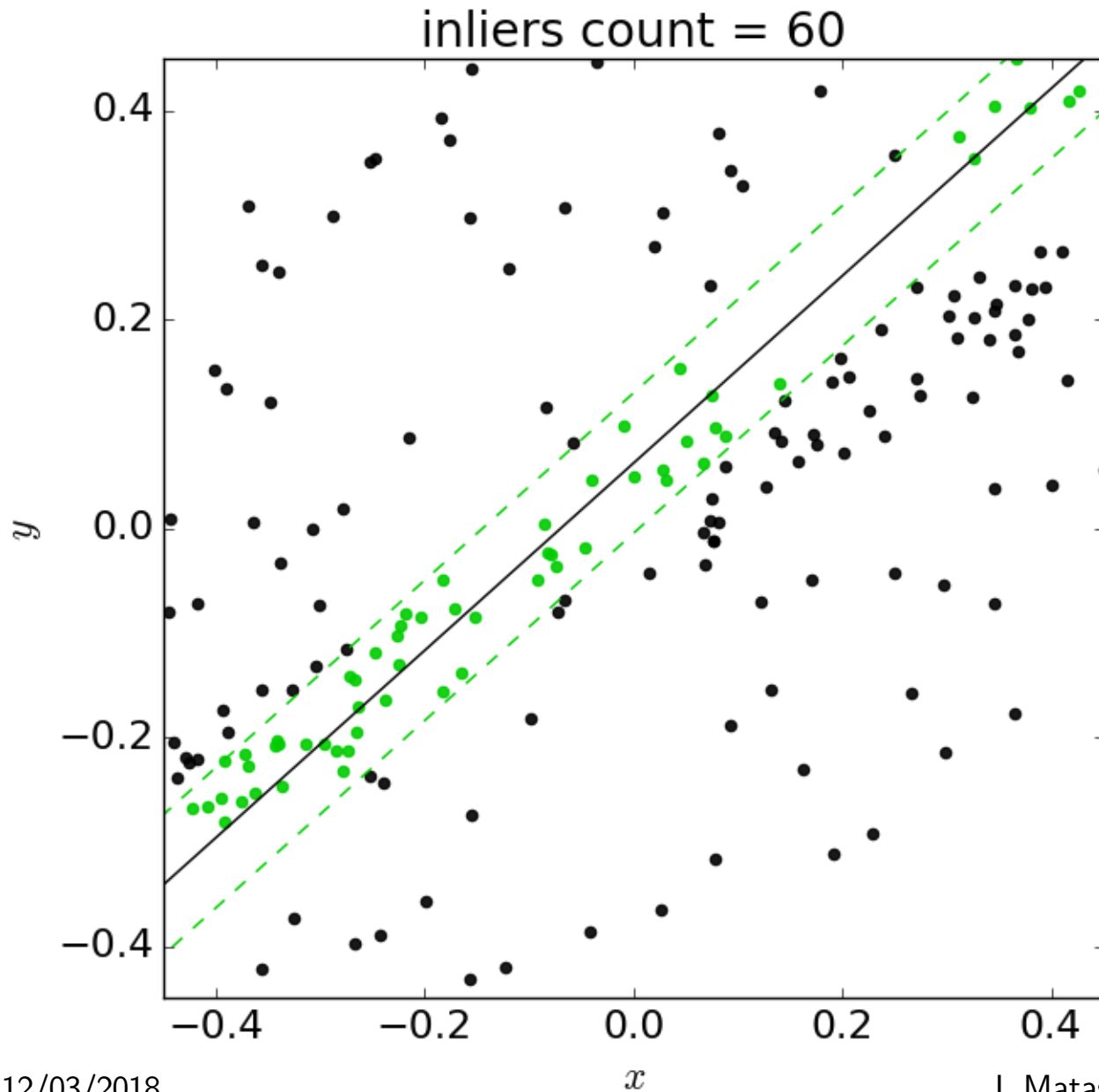
8: **until** $P(\text{better solution exists}) = f(|\mathcal{X}|, J^*, iter) < \eta$

9: **gone**

LO-RANSAC: Example

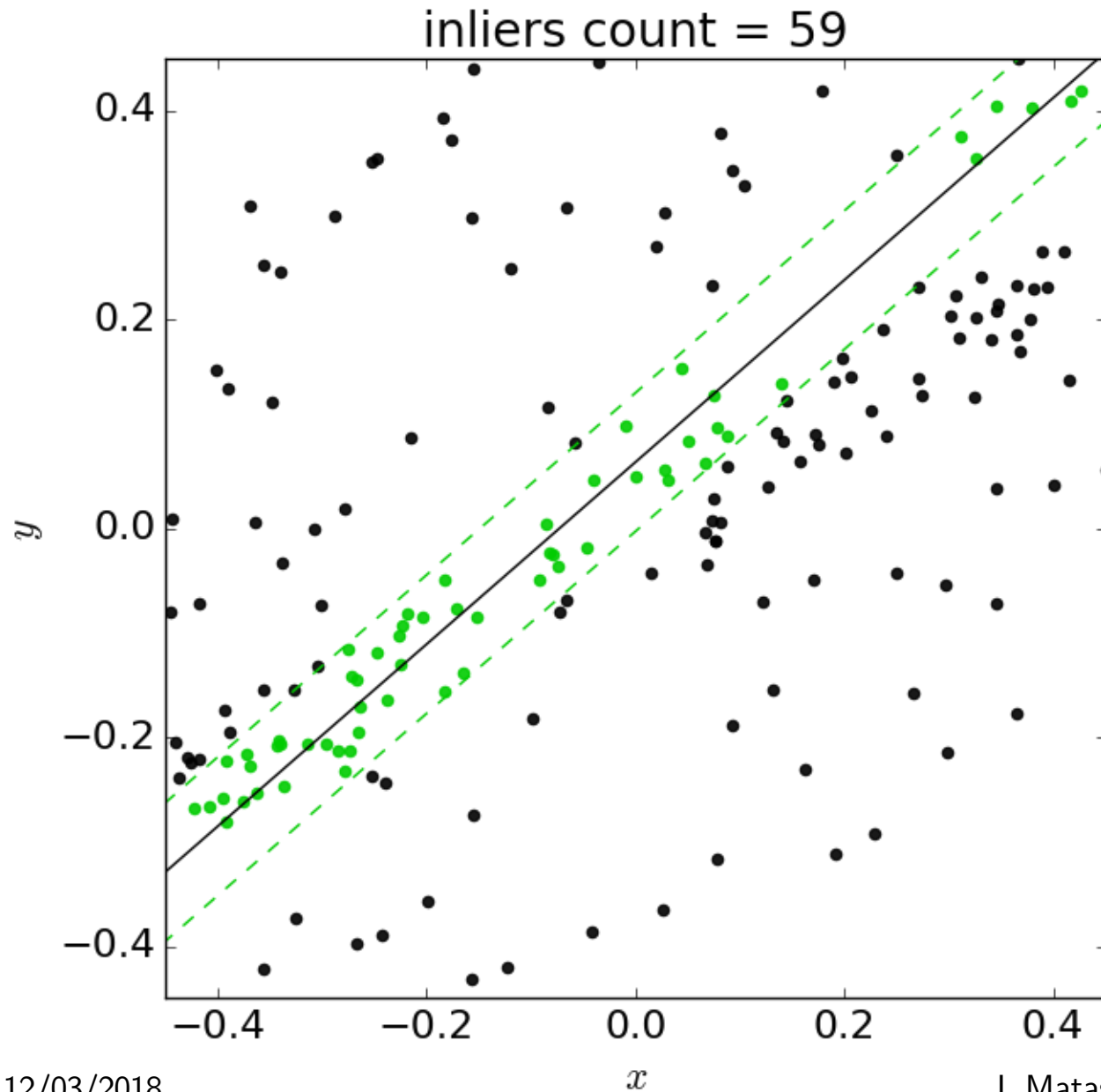


LO-RANSAC: Example



Init
Iteration 1

LO-RANSAC: Example

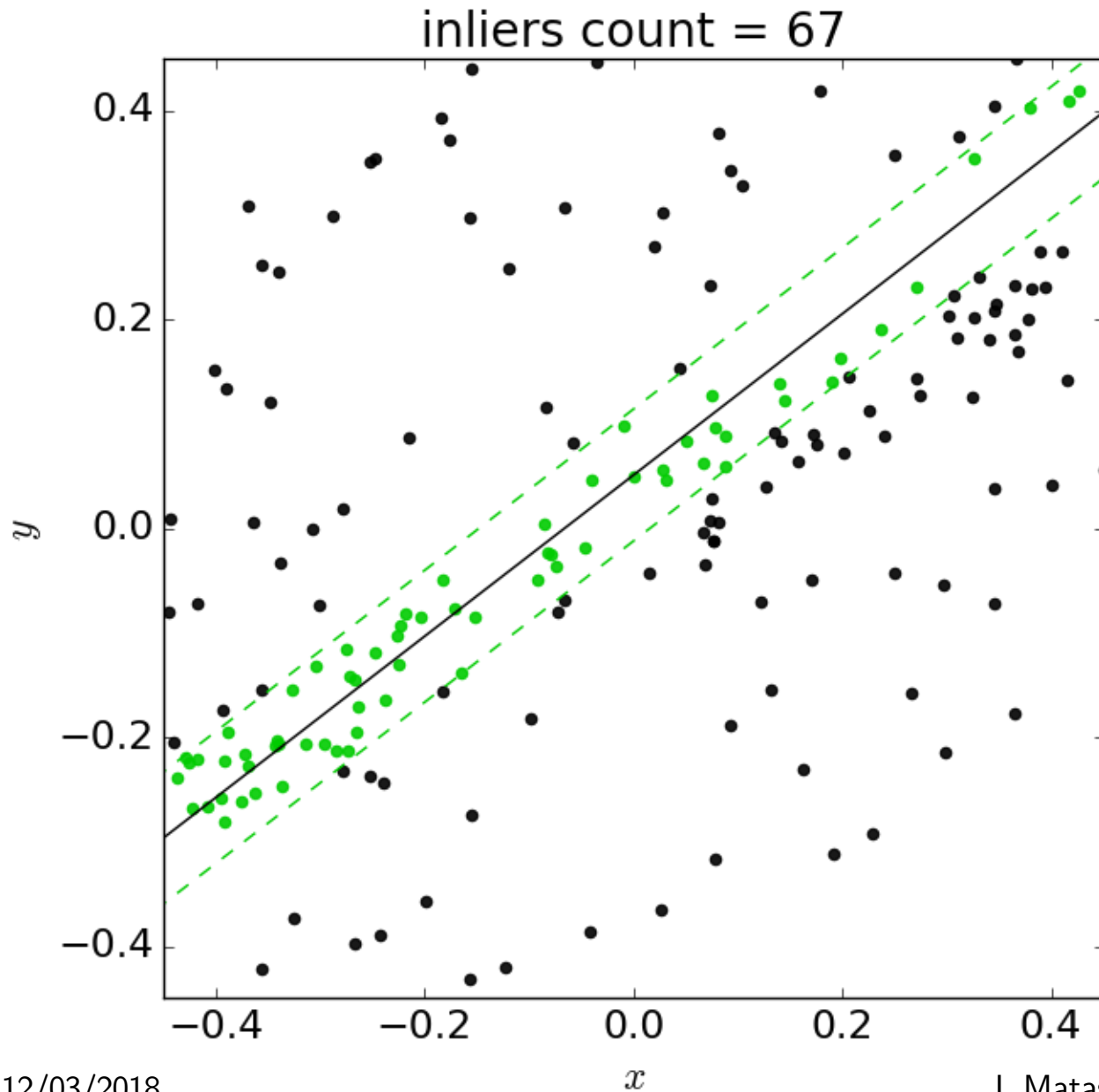


Init

Iteration 1

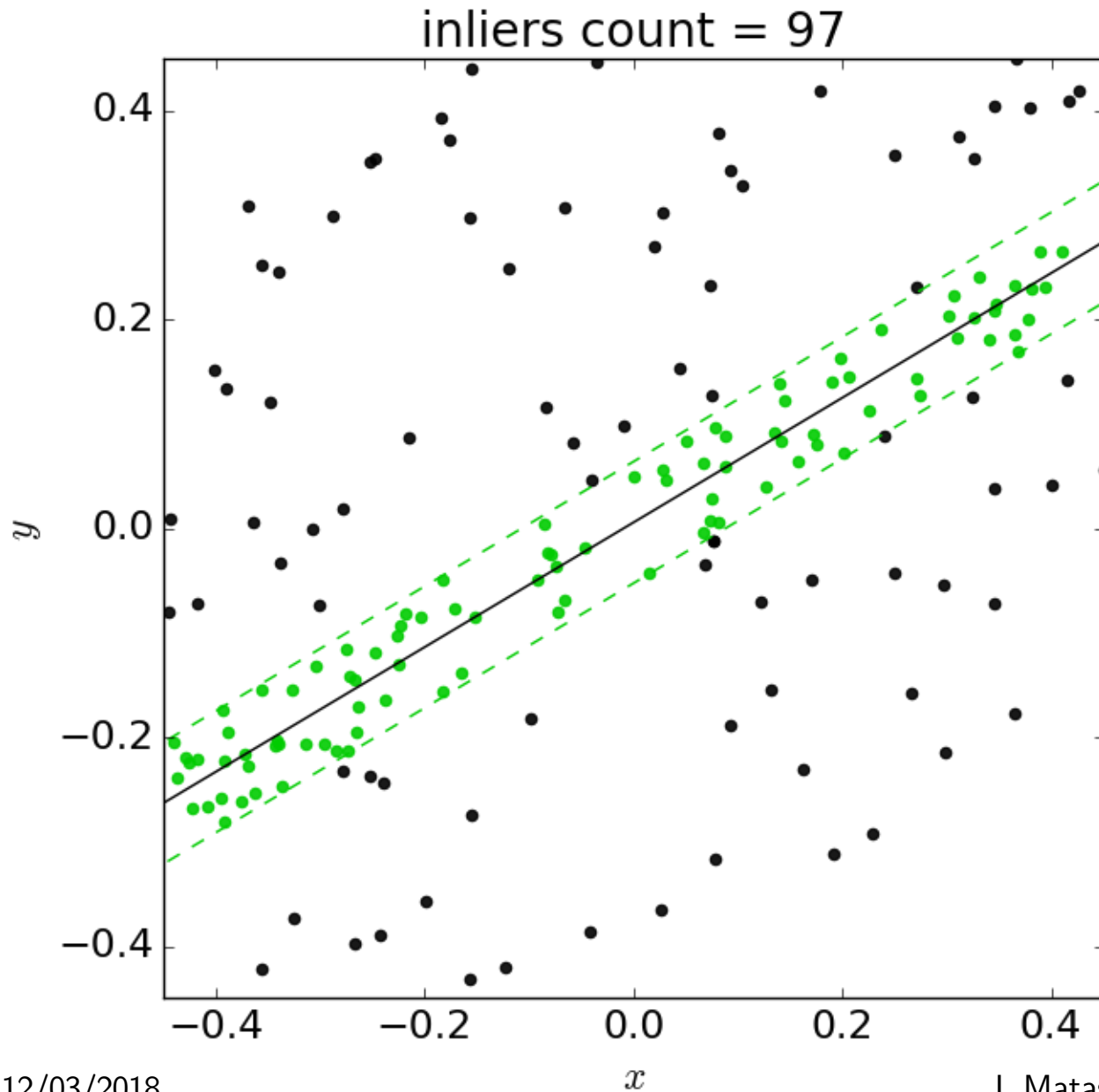
Iteration 2

LO-RANSAC: Example



- Init
- Iteration 1
- Iteration 2
- ...
- Iteration 7

LO-RANSAC: Example

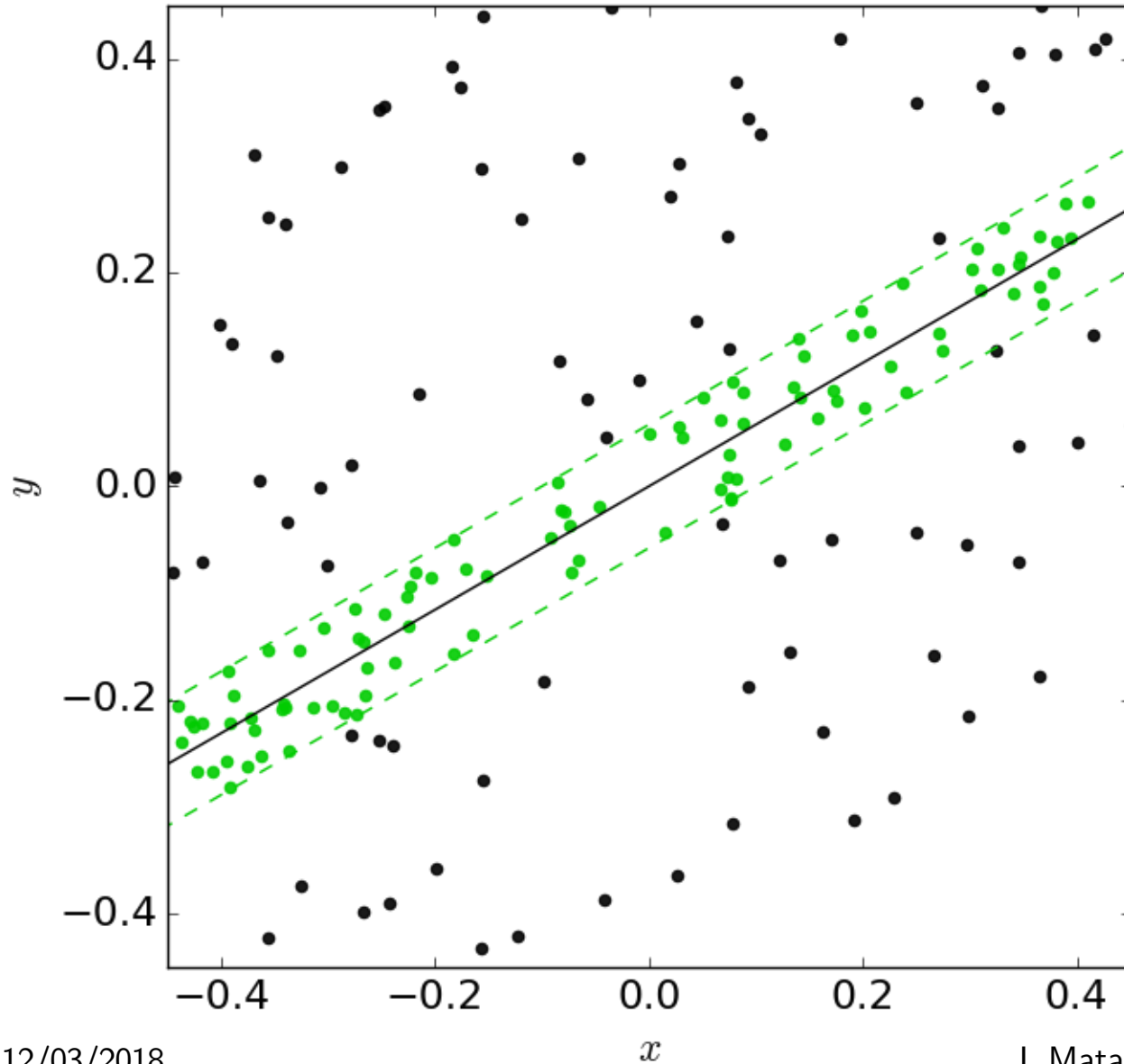


Init
Iteration 1
Iteration 2
...
Iteration 7
...
Iteration 15

LO-RANSAC: Example



Comparison with model (100 inliers):

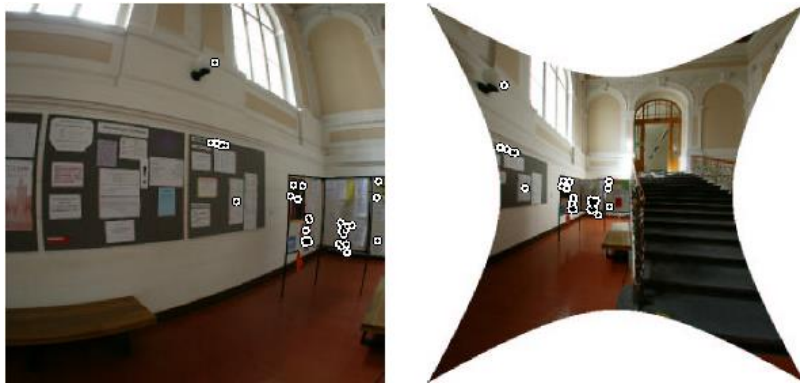


Locally Optimized RANSAC



Estimation of (approximate) models with lower complexity (less data points in the sample) followed by LO step estimating the desired model speeds the estimation up significantly.

The estimation of epipolar geometry is up to 10000 times faster when using 3 region-to-region correspondences rather than 7 point-to-point correspondences.



Fish-eye images by Braňo Mičušík

Simultaneous estimation of radial distortion and epipolar geometry with LO is superior to the state-of-the-art in both speed and precision of the model.

It was observed experimentally that RANSAC takes several times longer than theoretically expected. This is due to the noise – not every all-inlier sample generates a good hypothesis.

By applying local optimization (LO) to the-best-so-far hypotheses:

- (i) a near perfect agreement with theoretical performance
- (ii) lower sensitivity to noise and poor conditioning.

The LO is shown to be executed so rarely, $\log(\textit{iter})$ times, that it has minimal impact on the execution time.

RANSAC – Time Complexity

Repeat k times (k is a function of η , Q , N)

1. Hypothesis generation

- Select a sample of m data points
- Calculate parameters of the model(s)

2. Model verification

- Find the support (consensus set) by
- verifying all N data points

t_M – time needed to draw a sample

\bar{m}_s – average number of models per sample

Total **running** time:

$$t = k(t_M + \bar{m}_s N)$$

Repeat $k/(1-\alpha)$ times

1. Hypothesis generation
2. Model pre-verification $T_{d,d}$ test
 - Verify $d \ll N$ data points, reject
 - the model if not all d data points
 - are consistent with the model
3. Model verification
 - Verify the rest of the data points

V – average number of data points verified

α – probability that a good model is rejected by $T_{d,d}$ test

$$t = \frac{k}{1 - \alpha} (t_M + \bar{m}_s V)$$

Optimal Randomised Strategy

Model Verification is Sequential Decision Making

$$H_g: P(x_i = 1 | H_g) \geq \varepsilon$$

$$H_b: P(x_i = 1 | H_b) = \delta$$

$x_i = 1$ x_i is consistent with the model

where

H_g - hypothesis of a `good` model (\approx from an uncontaminated sample)

H_b - hypothesis of a `bad` model, (\approx from a contaminated sample)

δ - probability of a data point being consistent with an arbitrary model

Optimal (the fastest) test that ensures with probability α that that H_g is not incorrectly rejected is the Sequential probability ratio test (SPRT) [Wald47]

Compute the likelihood ratio

$$\lambda_i = \prod_{j=1}^i \frac{P(x_j|H_b)}{P(x_j|H_g)}$$

if $\lambda_i > A$ reject the model

if $i = N$ accept model as 'good'

Two important properties of SPRT:

1. probability of rejecting a 'good' model $\alpha < 1/A$
2. average number of verifications $V = C \log(A)$

$$C \approx \left(P(0|H_b) \log \frac{P(0|H_b)}{P(0|H_g)} + P(1|H_b) \log \frac{P(1|H_b)}{P(1|H_g)} \right)^{-1}$$

1. Probability of rejecting a \good\ model $\alpha=1/A$

$$\lambda_i = \prod_{j=1}^i \frac{P(x_j|H_b)}{P(x_j|H_g)} = \frac{P(x|H_b)}{P(x|H_g)}, x = (x_1, \dots, x_i)$$

If $\lambda_i > A$ then $P(x|H_g) < P(x|H_b)/A$, therefore

$$\begin{aligned} \alpha &= \int_{\lambda_i > A} P(x|H_g) dx < \int_{\lambda_i > A} P(x|H_b)/A dx = \\ &= \frac{1}{A} \int_{\lambda_i > A} P(x|H_b) dx \leq \frac{1}{A} \int P(x|H_b) dx = \frac{1}{A} \end{aligned}$$

WaldSAC

Repeat $k/(1-1/A)$ times

1. Hypothesis generation
2. Model verification, use SPRT

$$\bar{m}_S \cdot C \log A$$
$$C \approx \left((1 - \delta) \log \frac{1 - \delta}{1 - \varepsilon} + \delta \log \frac{\delta}{\varepsilon} \right)^{-1}$$

$$t(A) = \frac{k}{(1 - 1/A)} (t_M + \bar{m}_S C \log A)$$

In sequential statistical decision problem decision errors are traded off for time. These are two incomparable quantities, hence the constrained optimization.

In WaldSAC, decision errors cost time (more samples) and there is a single minimised quantity, time $t(A)$, a function of a single parameter A .

Optimal test (optimal A) given ε and δ



Optimal A^*

$$A^* = \arg \min_A t(A)$$

Optimal A^* found by solving

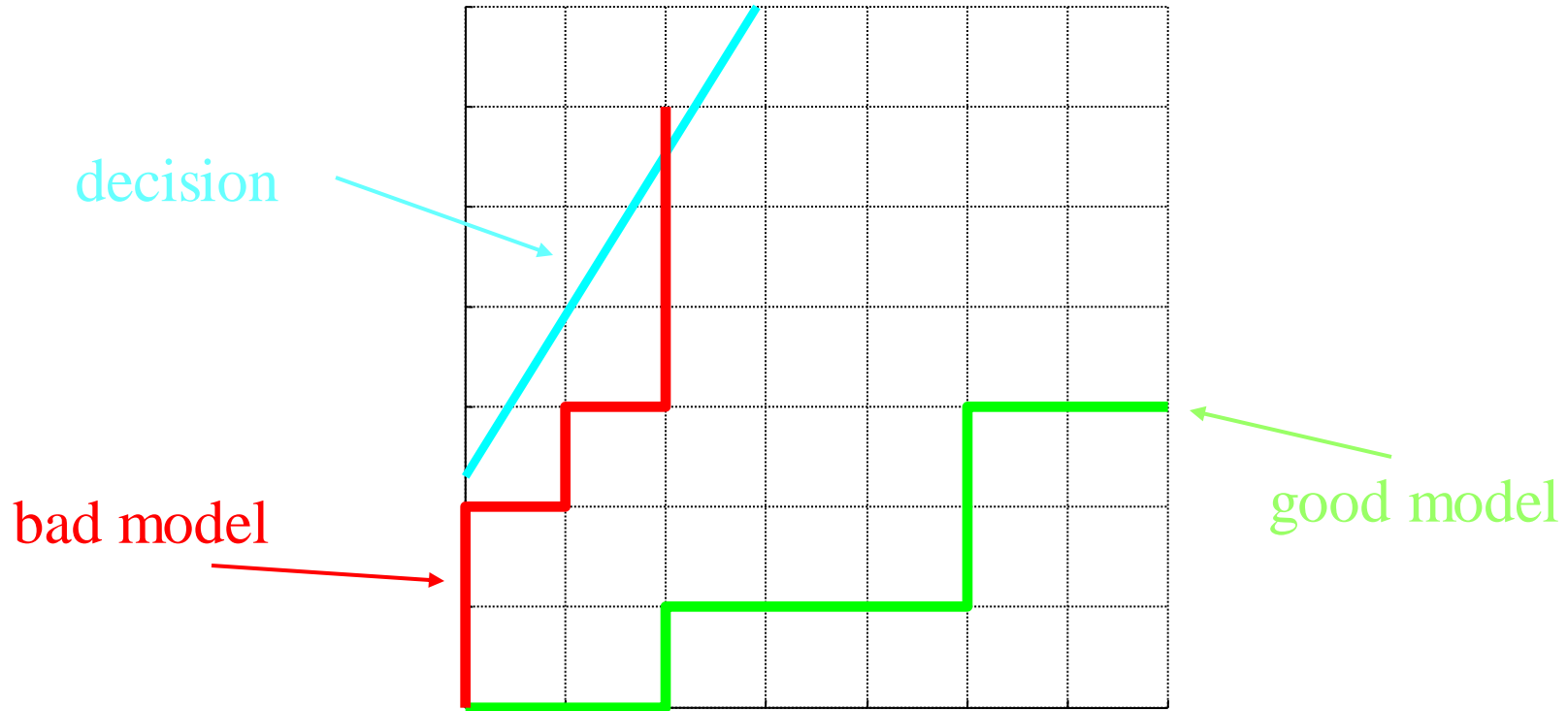
$$\frac{\partial t}{\partial A} = 0$$

$$A^* = \frac{t_M}{\bar{m}_s C} + 1 + \log A^*$$

$$A^* = \lim_{n \rightarrow \infty} A_n$$

$$A_0 = \frac{t_M}{\bar{m}_s C} + 1, \quad A_{n+1} = \frac{t_M}{\bar{m}_s C} + 1 + \log A_n$$

SPRT



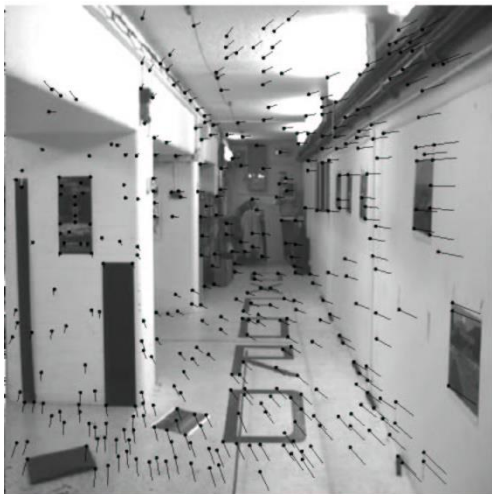
Note: the Wald's test is equivalent to series of $T(d, c)$, where $c = \lceil (\log A - d \log \lambda_1) / \log \lambda_0 \rceil$

Exp. 1: Wide-baseline matching



	samples	models	V	time	spd-up
R	2914	7347	110.0	1099504	1.0
R-R	7825	19737	3.0	841983	1.3
Wald	3426	8648	8.2	413227	2.7

Exp. 2 Narrow-baseline stereo

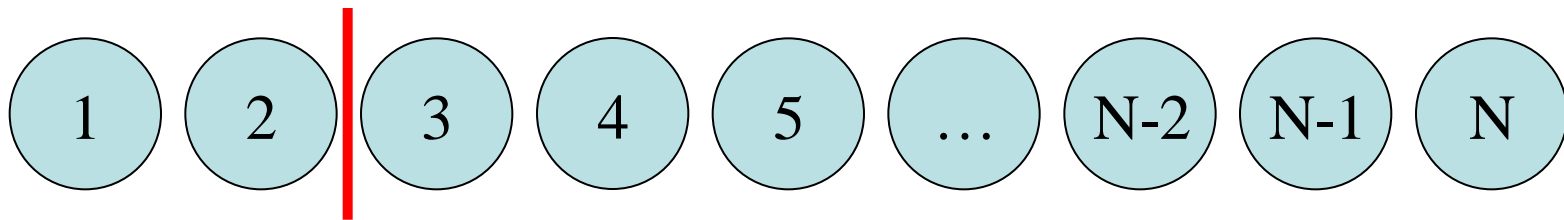


	samples	models	V	time	spd-up
R	155	367	600.0	235904	1.0
R-R	247	587	86.6	75539	3.1
Wald	162	384	23.1	25032	9.4

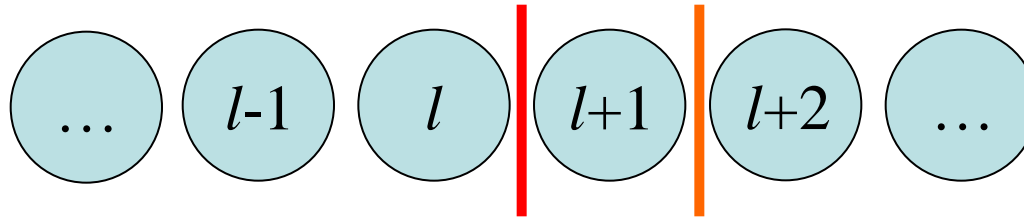


- The same confidence η in the solution reached faster (data dependent, $\approx 10x$)
- No change in the character of the algorithm, it was randomised anyway.
- Optimal strategy derived using Wald`s theory for known ε and δ .
- Results with ε **and** δ estimated during the course of RANSAC are not significantly different. Performance of SPRT is insensitive to errors in the estimate.
- δ can be learnt, an initial estimate can be obtained by geometric consideration
- Lower bound on ε is given by the best-so-far support
- Note that the properties of WaldSAC are quite different from preemptive RANSAC!

- Not all correspondences are created equally
- Some are better than others
- Sample from the best candidates first



Sample from here



Draw T_l samples from $(1 \dots l)$

Draw T_{l+1} samples from $(1 \dots l+1)$

Samples from $(1 \dots l)$ that are not from $(1 \dots l+1)$ contain

$l+1$

Draw $T_{l+1} - T_l$ samples of size $m-1$ and add

$l+1$

Degenerate Configurations

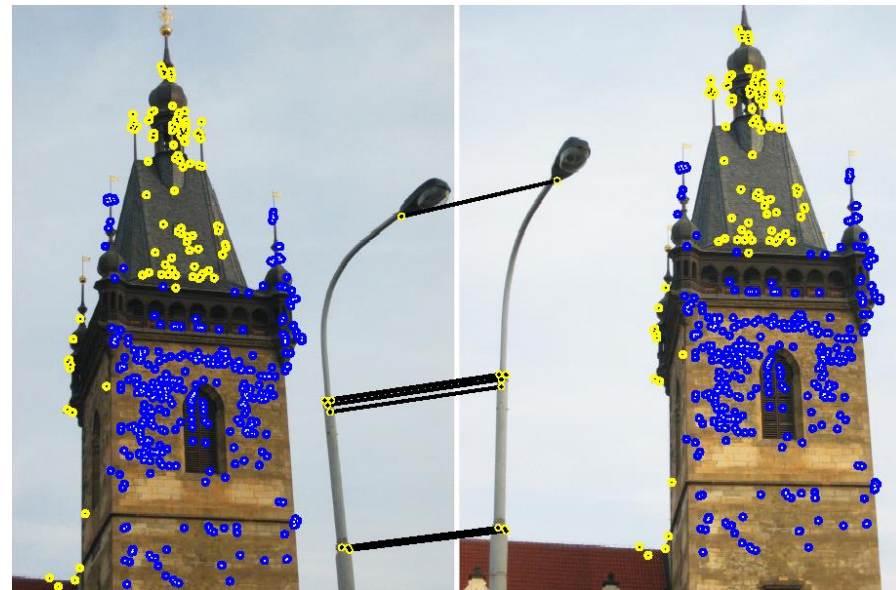


The presence of degenerate configuration causes RANSAC to fail in estimating a correct model, instead a model consistent with the degenerate configuration and some outliers is found.

The DEGENSAC algorithm handles scenes with:

- all points in a single plane
- majority of the points in a single plane and the rest off the plane
- no dominant plane present

No a-priori knowledge of the type of the scene is required



Chum, Werner, Matas: Epipolar Geometry Estimation unaffected by dominant plane, *CVPR 2005*