

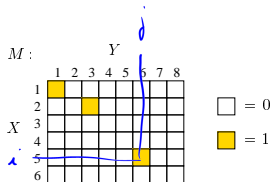
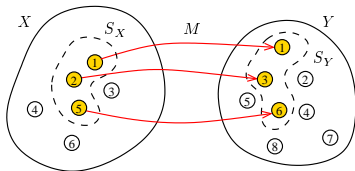
# ►The Full Problem of Matching and Fundamental Matrix Estimation

**Problem:** Given two sets of image points  $X = \{x_i\}_{i=1}^m$  and  $Y = \{y_j\}_{j=1}^n$  and their descriptors  $D$ , find the most probable

1. inliers  $S_X \subseteq X$ ,  $S_Y \subseteq Y$
2. one-to-one perfect matching  $M: S_X \rightarrow S_Y$
3. fundamental matrix  $\mathbf{F}$  such that  $\text{rank } \mathbf{F} = 2$
4. such that for each  $x_i \in S_X$  and  $y_j = M(x_i)$  it is probable that
  - a. the image descriptor  $D(x_i)$  is similar to  $D(y_j)$ , and
  - b. the total geometric error  $\sum_{ij} e_{ij}^2(\mathbf{F})$  is small
5. inlier-outlier and outlier-outlier matches are improbable

perfect matching: 1-factor of the bipartite graph

note a slight change in notation:  $e_{ij}$



$$(M^*, \mathbf{F}^*) = \arg \max_{M, \mathbf{F}} p(M, \mathbf{F} \mid X, Y, D) \quad (17)$$

- probabilistic model: an efficient language for task formulation
- the (17) is a p.d.f. for all the involved variables (there is a constant number of variables!)
- binary matching table  $M_{ij} \in \{0, 1\}$  of fixed size  $m \times n$ 
  - each row/column contains at most one unity
  - zero rows/columns correspond to unmatched point  $x_i/y_j$

# Deriving A Robust Matching Model by Marginalization

For algorithmic efficiency, instead of  $(M^*, \mathbf{F}^*) = \arg \max_{M, \mathbf{F}} p(M, \mathbf{F} | X, Y, D)$  we will solve

$$\sum_a p(a, b) = p(b)$$

$$\mathbf{F}^* = \arg \max_{\mathbf{F}} p(\mathbf{F} | X, Y, D)$$

$$p_{\mathbf{F}|X,Y,D}(\mathbf{F}|X,Y,D) \mathbf{F} \quad (18)$$

$$p_{M,\mathbf{F}|X,Y,D}(M, \mathbf{F}|X,Y,D)$$

by marginalization of  $p(M, \mathbf{F} | X, Y, D)$  over  $M$

this simplification changes the problem!

$$p(M, \mathbf{F} | X, Y, D) \simeq p(M, \mathbf{F}, X, Y, D) = p(X, Y, D, M | \mathbf{F}) \cdot p(\mathbf{F})$$

assuming correspondence-wise independence:

$$p(X, Y, D, M | \mathbf{F}) = \prod_{i=1}^m \prod_{j=1}^n p(x_i, y_j, D, m_{ij} | \mathbf{F}) \stackrel{\text{def}}{=} \prod_{i=1}^m \prod_{j=1}^n p_e(e_{ij}, d_{ij}, m_{ij} | \mathbf{F})$$

- $e_{ij}$  represents geometric error for match  $x_i \leftrightarrow y_j$ :  $e_{ij}(x_i, y_j | \mathbf{F})$
- $d_{ij}$  represents descriptor similarity for match  $x_i \leftrightarrow y_j$ :  $d_{ij} = \|\mathbf{d}(x_i) - \mathbf{d}(y_j)\|$

SIFT etc

Marginalization:

$$\sum_{m_{11} \in \{0,1\}} \sum_{m_{12}} \cdots \sum_{m_{mn}} p(X, Y, D, M | \mathbf{F}) = \sum_{m_{11}} \sum_{m_{12}} \cdots \sum_{m_{mn}} \prod_{i=1}^m \prod_{j=1}^n p_e(e_{ij}, d_{ij}, m_{ij} | \mathbf{F}) =$$

$$\sum_{m_{11}} \sum_{m_{12}} \cdots \sum_{m_{mn}} \left( \sum_{m_{11}} \sum_{m_{12}} \cdots \sum_{m_{mn}} p_e(e_{11}, d_{11}, m_{11} | \mathbf{F}) p_e(e_{12}, d_{12}, m_{12} | \mathbf{F}) \cdots p_e(e_{m1}, d_{m1}, m_{m1} | \mathbf{F}) \right) \prod_{i=1}^m \prod_{j=1}^n p_e(e_{ij}, d_{ij}, m_{ij} | \mathbf{F}) = p(X, Y, D | \mathbf{F})$$

we will continue with this term

## Robust Matching Model (cont'd)

$$\begin{aligned}
 \sum_{m_{ij} \in \{0,1\}} p_e(e_{ij}, d_{ij}, m_{ij} \mid \mathbf{F}) &= \sum_{m_{ij} \in \{0,1\}} p_e(e_{ij}, d_{ij} \mid m_{ij}, \mathbf{F}) \cdot p(m_{ij} \mid \mathbf{F}) = \\
 &= \underbrace{p_e(e_{ij}, d_{ij} \mid m_{ij} = 1, \mathbf{F})}_{p_1(e_{ij}, d_{ij} \mid \mathbf{F})} \underbrace{p(m_{ij} = 1 \mid \mathbf{F})}_{1-\alpha_0} + \underbrace{p_e(e_{ij}, d_{ij} \mid m_{ij} = 0, \mathbf{F})}_{p_0(e_{ij}, d_{ij} \mid \mathbf{F})} \underbrace{p(m_{ij} = 0 \mid \mathbf{F})}_{\alpha_0} = \\
 &= (1 - \alpha_0) \left[ p_1(e_{ij}, d_{ij} \mid \mathbf{F}) + \frac{\alpha_0}{1 - \alpha_0} p_0(e_{ij}, d_{ij} \mid \mathbf{F}) \right] \quad (19)
 \end{aligned}$$

- the  $p_0(e_{ij}, d_{ij} \mid \mathbf{F}) \approx \text{const}$  is a penalty for 'missing a correspondence' but it should be a p.d.f. (cannot be a constant) (see Slide 108 for a simplification)

$$\alpha_0 \rightarrow 1, \quad p_0 \rightarrow 0 \quad \text{so that} \quad \frac{\alpha_0}{1 - \alpha_0} p_0 \approx \text{const}$$

- the  $p_1(e_{ij}, d_{ij} \mid \mathbf{F})$  is typically an easy-to-design component: assuming independence of geometric error and descriptor similarity:

$$p_1(e_{ij}, d_{ij} \mid \mathbf{F}) = p_1(e_{ij} \mid \mathbf{F}) \cdot p_1(d_{ij})$$

- we choose, eg.

$$p_1(e_{ij} \mid \mathbf{F}) = \frac{1}{T_e(\sigma_1, \mathbf{F})} e^{-\frac{e_{ij}^2(\mathbf{F})}{2\sigma_1^2}}, \quad p_1(d_{ij}) = \frac{1}{T_d(\sigma_d, \dim \mathbf{d})} e^{-\frac{\|\mathbf{d}(x_i) - \mathbf{d}(y_j)\|^2}{2\sigma_d^2}} \quad (20)$$

- $\sigma_1, \sigma_d, \alpha_0$  are 'hyper-parameters'
- the form of  $T(\sigma_1, \mathbf{F})$  depends on error definition
- we will continue with the result from (19)

$$= \sqrt{2\sigma_d^2} \sigma_d^{\dim \mathbf{d}}$$

$$S = \sigma_d^2 I_{\dim \mathbf{d}}$$

## ► Simplified Robust Energy (Error) Function

- assuming the choice of  $p_1$  as in (20), we are simplifying

$$p(X, Y, D \mid \mathbf{F}) = \prod_{i=1}^m \prod_{j=1}^n \left[ (1 - \alpha_0) p_1(e_{ij}, d_{ij} \mid \mathbf{F}) + \alpha_0 p_0(e_{ij}, d_{ij} \mid \mathbf{F}) \right] \quad (21)$$

- we define 'energy' as:  $V(x) = -\log p(x)$   *$H = -\int p(x) \log p(x) dx$  this helps simplify the formulas*
- for simplicity, we omit  $d_{ij}$  *entropy = mean energy*
- we choose  $\sigma_0 \gg \sigma_1$  and the missed-correspondence penalty function as

$$p_0(e_{ij} \mid \mathbf{F}) = \frac{1}{T_e(\sigma_0, \mathbf{F})} e^{-\frac{e_{ij}^2(\mathbf{F})}{2\sigma_0^2}} \quad \sigma_0 \rightarrow \infty$$

- then

$$V(X, Y, D \mid \mathbf{F}) = \sum_{i=1}^m \sum_{j=1}^n \left[ \underbrace{-\log \frac{1 - \alpha_0}{T_e(\sigma_1, \mathbf{F})}}_{\Delta(\mathbf{F})} - \log \left( e^{-\frac{e_{ij}^2(\mathbf{F})}{2\sigma_1^2}} + \underbrace{\frac{\alpha_0}{1 - \alpha_0} \frac{T_e(\sigma_1, \mathbf{F})}{T_e(\sigma_0, \mathbf{F})} e^{-\frac{e_{ij}^2(\mathbf{F})}{2\sigma_0^2}}}_{t \approx \text{const}} \right) \right]$$

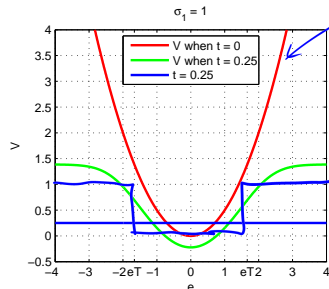
- by choosing representative of  $\mathbf{F}$  such that  $\Delta(\mathbf{F}) = \text{const}$ , we get

$$V(X, Y, D \mid \mathbf{F}) = m n \Delta + \sum_{i=1}^m \sum_{j=1}^n \underbrace{-\log \left( e^{-\frac{e_{ij}^2(\mathbf{F})}{2\sigma_1^2}} + t \right)}_{\hat{V}(e_{ij})} \quad (22)$$

note that  $m, n$  are fixed

# ►The Action of the Robust Matching Model on Data

Example for  $\hat{V}(e)$  from (22):  $e^2(F)$



red – the usual (non-robust) error when  $t = 0$   
 blue – the rejected correspondence penalty  $t$   
 green – ‘robust energy’ (22)

- if the error of a correspondence exceeds a limit, it is ignored
- then  $\hat{V}(e) = \text{const}$  and we essentially count outliers in (22)
- $t$  controls the ‘turn-off’ point
- the inlier/outlier threshold is  $e_T$  is the error for which  
 $(1 - \alpha_0) p_1(e_T) = \alpha_0 p_0(e_T)$ : note that  $t \approx 0$

$$e_T = \sigma_1 \sqrt{-\log t^2} \quad (23)$$

The full optimization problem is (18):

$$\begin{aligned} \mathbf{F}^* &= \arg \max_{\mathbf{F}} p(\mathbf{F} | X, Y, D) = \arg \max_{\mathbf{F}} \frac{\overbrace{p(X, Y, D | \mathbf{F})}^{\text{likelihood}} \cdot \overbrace{p(\mathbf{F})}^{\text{prior}}}{\underbrace{p(X, Y, D)}_{\text{evidence}}} = \\ &= \arg \min_{\mathbf{F}} \{V(X, Y, D | \mathbf{F}) + V(\mathbf{F})\} \end{aligned}$$

- typically we take  $V(\mathbf{F}) = 0$  unless we need to stabilize a computation, e.g. when video camera moves smoothly (on a high-mass vehicle) and we have a prediction for  $\mathbf{F}$
- evidence is not needed unless we want to compare different models

# Discussion: On The Art of Probabilistic Model Design...

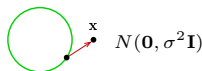
- a few models for fitting zero-centered circle  $C$  of radius  $r$  to points in  $\mathbb{R}^2$

marginalized over  $C$

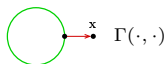
orthogonal deviation from  $C$

Sampson approximation

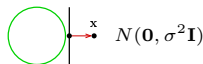
error model



$$N(\mathbf{0}, \sigma^2 \mathbf{I})$$

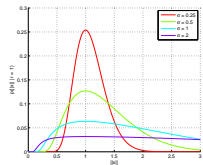
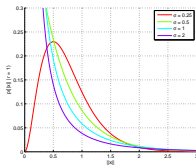
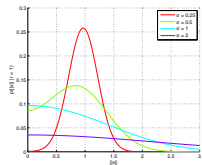


$$\Gamma(\cdot, \cdot)$$

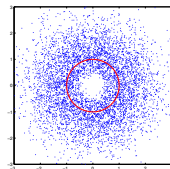
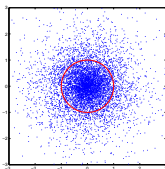
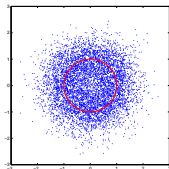


$$N(\mathbf{0}, \sigma^2 \mathbf{I})$$

radial p.d.f.



random sample



$p(\mathbf{x} | r)$

$$\approx \frac{1}{\sigma \sqrt{(2\pi)^3 r \|\mathbf{x}\|}} e^{-\frac{(\|\mathbf{x}\| - r)^2}{2\sigma^2}}$$

$$\frac{1}{2\pi\Gamma(\frac{r^2}{\sigma})} \frac{1}{\|\mathbf{x}\|^2} \left( \frac{r\|\mathbf{x}\|}{\sigma} \right)^{\frac{r^2}{\sigma}} e^{-\frac{r\|\mathbf{x}\|}{\sigma}}$$

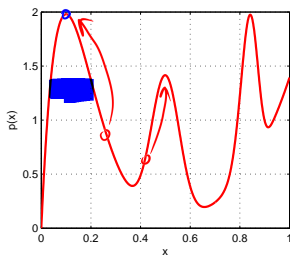
$$\frac{1}{r\sigma\sqrt{(2\pi)^3}} e^{-\frac{e^2(\mathbf{x};r)}{2\sigma^2}}$$

- mode inside the circle
- models the inside well
- tends to normal distrib.

- peak at the center
- unusable for small radii
- tends to Dirac distrib.

- mode at the circle
- hole at the center
- tends to normal distrib.

# How To Find the Global Maxima (Modes) of a PDF?



- consider the function  $p(x)$  at left p.d.f. on  $[0, 1]$ , mode at 0.1
- consider several methods:

## 1. exhaustive search

$\arg \max_x p(x)$

```
step = 1/(iterations-1);
for x = 0:step:1
    if p(x) > bestp
        bestx = x; bestp = p(x);
    end
end
```

- slow algorithm (definite quantization); faster variants exist
- fast to implement

## 2. randomized search with uniform sampling

```
x = rand(1); x = prnd;
if p(x) > bestp
    bestx = x; bestp = p(x);
end
```

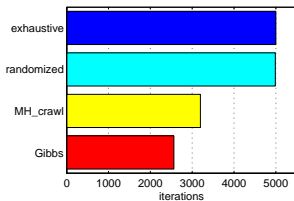
- slow algorithm but better convergence
- fast to implement
- how to stop it?

## 3. random sampling from $p(x)$ (Gibbs sampler)

- faster algorithm
- fast to implement but often infeasible (e.g. when  $p(x)$  is data dependent (our case))

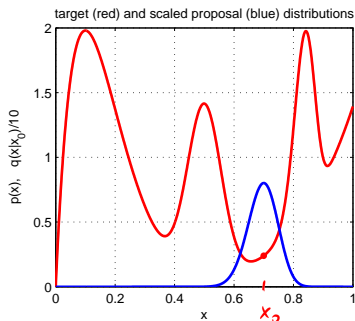
## 4. Metropolis-Hastings sampling

- almost as fast (with care)
- not so fast to implement
- rarely infeasible
- RANSAC belongs here



- averaged over  $10^4$  trials
- number of proposals before  $|x - x_{\text{true}}| \leq \text{step}$
- uniform and Gibbs give the theoretical result

# How To Generate Random Samples from a Complex Distribution?



- red: probability density function  $p(x)$  of a toy distribution on the unit interval **target distribution**

$$p(x) = \sum_{i=1}^4 \alpha_i \text{Be}(x; \alpha_i, \beta_i), \quad \sum_{i=1}^4 \alpha_i = 1, \quad \alpha_i \geq 0$$

$$\text{Be}(x; \alpha, \beta) = \frac{1}{\text{B}(\alpha, \beta)} \cdot x^{\alpha-1} (1-x)^{\beta-1}$$

$(\alpha_1, \alpha_2, \alpha_3, \alpha_4) \quad (0.1, 0.2, 0.5, 0.2) \quad | \quad (0.1, 0.3, 0.8, 1.0)$

- note we can generate samples from this  $p(x)$  **how?**
- suppose we cannot sample from  $p(x)$  but we can sample from some 'simple' distribution, given the last sample  $x_0$  (blue) **proposal distribution**

$$q(x | x_0) = \begin{cases} \text{U}_{0,1}(x) & \text{(independent) uniform sampling} \\ \text{Be}(x; \frac{x_0}{T} + 1, \frac{1-x_0}{T} + 1) & \text{'beta' diffusion (crawler) } \quad T = \text{temperature} \\ p(x) & \text{(independent) Gibbs sampler} \end{cases}$$

- note we have unified all the random sampling methods on the previous slide
- how to transform proposal samples  $q(x | x_0)$  to target distribution  $p(x)$  samples?



Thank You

