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# Robust model estimation from data contaminated by outliers

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# Fitting a Line





# • Select sample of m points at random



- Select sample of m points at random
- Calculate model parameters that fit the data in the sample



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- Calculate error function for each data point



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On average

N	number of point
Ι	number of inliers
т	size of the sample

$$P(\text{good}) = \frac{\binom{I}{m}}{\binom{N}{m}} = \prod_{j=0}^{m-1} \frac{I-j}{N-j}$$

mean time before the success E(k) = 1 / P(good)

With confidence *p* 

How large k?

... to hit at least one pair of points on the line l with probability <u>larger</u> than p (0.95)

Equivalently

... the probability of not hitting any pair of points on l is  $\leq 1-p$ 

With confidence *p* 

N	number of point
Ι	number of inliers
т	size of the sample

$$P(\text{good}) = \frac{\binom{I}{m}}{\binom{N}{m}} = \prod_{j=0}^{m-1} \frac{I-j}{N-j}$$

P(bad) = 1 - P(good)

$$P(bad k times) = (1 - P(good))^k$$

With confidence *p* 

$$P(bad k times) = (1 - P(good))^k \le 1 - p$$

$$k \log (1 - P(\text{good})) \le \log(1 - p)$$

$$k \ge \log(1-p) / \log (1 - P(\text{good}))$$

#### I/N[%]

ш		15%	20%	30%	40%	50%	70%
le	2	132	73	32	17	10	4
du	4	5916	1871	368	116	46	11
sai	7	$1.75 \cdot 10^{6}$	$2.34 \cdot 10^{5}$	$1.37 \cdot 10^{4}$	1827	382	35
Je	8	$1.17\cdot 10^7$	$1.17 \cdot 10^{6}$	$4.57 \cdot 10^{4}$	4570	765	50
f tł	12	$2.31 \cdot 10^{10}$	$7.31 \cdot 10^{8}$	$5.64 \cdot 10^{6}$	$1.79 \cdot 10^{5}$	$1.23 \cdot 10^4$	215
Õ	18	$2.08 \cdot 10^{15}$	$1.14 \cdot 10^{13}$	$7.73 \cdot 10^{9}$	$4.36 \cdot 10^{7}$	$7.85 \cdot 10^{5}$	1838
ize	30	$\infty$	$\infty$	$1.35 \cdot 10^{16}$	$2.60 \cdot 10^{12}$	$3.22 \cdot 10^{9}$	$1.33 \cdot 10^{5}$
S	40	$\infty$	$\infty$	$\infty$	$2.70 \cdot 10^{16}$	$3.29 \cdot 10^{12}$	$4.71 \cdot 10^{6}$



$$k = \frac{\log(1-p)}{\log\left(1-\frac{I}{N}\frac{I-1}{N-1}\right)}$$

k ... number of samples drawn
N ... number of data points
I ... time to compute a single model
p ... confidence in the solution (.95)

	RANSAC [Fischler, Bolles '81]
In: $U = \{x_i\}$	set of data points, $ \mathbf{U}  = \mathbf{N}$
$f(C) \cdot C \to \sigma$	function f computes model parameters p given a sample S from U
$f(S): S \to p$	the cost function for a single data point x
<b>Out:</b> $p^*$	p*, parameters of the model maximizing the cost function
k := 0	

Repeat until P{better solution exists} <  $\eta$  (a function of C<sup>\*</sup> and no. of steps k) k := k + 1

I. Hypothesis

(1) select randomly set  $S_k \subset U$ , sample size (2) compute parameters  $S_k \subset U$   $|S_k| = m$   $p_k = f(S_k)$ II. Verification (3) compute cost  $C_k = \sum_{x \in U} \rho(p_k, x)$ (4) if  $C^* < C_k$  then  $C^* := C_k$ ,  $p^* := p_k$ 

end

#### (III. Is the solution valid?)

# **RANSAC Issues**:

RANSAC is a very general robust estimation method, any problems?

#### - Correctness of the results. Degeneracy.

Solution: DegenSAC.

#### - Stopping criterion?:

Repeat until

 $P\{\text{better solution exists}\} < \eta \text{ (a function of } C^* \text{ and no. of steps } k)$ 

#### - Cost function:

Solutions: Least median of Squares, MINPRAN

- Accuracy (model parameters are estimated from minimal samples): Solution: Locally Optimized RANSAC

- Speed:

Running time grows with

1. number of data points,

2. number of iterations (polynomial in inlier ratio)

Addressing the problem:

RANSAC with SPRT (WaldSAC), PROSAC

#### RANSAC – Time Complexity



- Select a sample of *m* data points
- Calculate parameters of the model(s)
- 2. Model verification
- Find the support (consensus set) by
- verifying all N data points
- $t_M$  time needed to draw a sample
- $\overline{m}_s$  average number of models per sample
- I— the number of inliers
- N the number of data points
- $\eta$  confidence in the solution



 $\overline{m}_s\cdot N$ 

Total running time:

$$t = k(t_M + \overline{m}_s N)$$





### RANSAC time complexity

$$t = k(t_M + \overline{m}_s N)$$

The number of samples  $k \sim \frac{1}{P}$ 

#### where P is a probability of drawing an all-inlier sample

$$P = \frac{\binom{I}{m}}{\binom{N}{m}} \approx \varepsilon^m$$

where m is size of the sample and  $\varepsilon$  is the fraction of inliers I/N

 $t = \frac{k}{1 - \alpha} (t_M + \overline{m}_s V)$ 

## Randomised RANSAC [Matas, Chum 02]

#### Repeat $k/(1-\alpha)$ times

J. Matas, O. Chum, 25RANSAC workshop CVI

2006 éž

- 1. Hypothesis generation
- 2. Model pre-verification  $T_{d,d}$  test
- Verify  $d \ll N$  data points, reject
- the model if not all d data points
- are consistent with the model
- 3. Model verification

Verify the rest of the data points

V- average number of data points verified

 $\alpha$  – probability that a good model is rejected by  $T_{d,d}$  test

![](_page_19_Picture_12.jpeg)

 $t_M$ 

Time

 $\overline{m}_s\cdot V$ 

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### **Optimal Randomised Strategy**

![](_page_20_Picture_2.jpeg)

Model Verification is Sequential Decision Making  $H_g: P(x_i = 1|H_g) \ge \varepsilon$   $H_b: P(x_i = 1|H_b) = \delta$  $x_i = 1$   $x_i$  is consistent with the model

where

 $H_{g}$  - hypothesis of a `good` model ( $\approx$  from an uncontaminated sample)  $H_{b}$  - hypothesis of a `bad` model, ( $\approx$  from a contaminated sample)  $\delta$  - probability of a data point being consistent with an arbitrary model

Optimal (the fastest) test that ensures with probability  $\alpha$  that that  $H_g$  is not incorrectly rejected

is the

Sequential probability ratio test (SPRT) [Wald47]

#### SPRT [simplified from Wald 47]

![](_page_21_Picture_2.jpeg)

Compute the likelihood ratio

![](_page_21_Picture_4.jpeg)

if  $\lambda_i > A$  reject the model if i = N accept model as 'good'

Two important properties of SPRT:

- 1. probability of rejecting a  $\log \alpha < 1/A$
- 2. average number of verifications  $V=C \log(A)$

 $C \approx \left( P(0|H_b) \log \frac{P(0|H_b)}{P(0|H_g)} + P(1|H_b) \log \frac{P(1|H_b)}{P(1|H_g)} \right)^{-1}$ 

## SPRT properties

![](_page_22_Picture_2.jpeg)

1. Probability of rejecting a  $good \mod \alpha = 1/A$ 

 $\lambda_i = \prod_{j=1}^i \frac{P(x_j | H_b)}{P(x_j | H_g)} = \frac{P(x | H_b)}{P(x | H_g)}, x = (x_1, \dots, x_i)$ If  $\lambda_i > A$  then  $P(x | H_g) < P(x | H_b)/A$ , therefore

 $\alpha = \int_{\lambda_i > A} P(x|H_g) dx < \int_{\lambda_i > A} P(x|H_b) / A dx =$  $= \frac{1}{A} \int_{\lambda_i > A} P(x|H_b) dx \le \frac{1}{A} \int P(x|H_b) dx = \frac{1}{A}$ 

![](_page_23_Picture_1.jpeg)

![](_page_23_Picture_2.jpeg)

Time

Repeat k/(1-1/A) times 1. Hypothesis generation 2. Model verification use SPRT

 $t_M$  $\overline{m}_s \cdot C \log A$  $C \approx ((1 - \delta) \log \frac{1 - \delta}{1 - \epsilon} + \delta \log \frac{\delta}{\epsilon})^{-1}$  $t(A) = \frac{k}{(1 - 1/A)}(t_M + \overline{m}_S C \log A)$ 

In sequential statistical decision problem decision errors are traded off for time. These are two incomparable quantities, hence the constrained optimization.

In WaldSAC, decision errors cost time (more samples) and there is a single minimised quantity, time t(A), a function of a single parameter A.

## Optimal test (optimal A) given $\varepsilon$ and $\delta$

Center for Machine Percention Prague

Optimal  $A^*$ 

Optimal  $A^*$  found by solving

![](_page_24_Picture_5.jpeg)

 $A^* = \arg\min_A t(A)$ 

 $\frac{\partial t}{\partial A} = 0$ 

 $A^* = \lim_{n \to \infty} A_n$ 

 $A_0 = \frac{t_M}{\overline{m}_s C} + 1, \quad A_{n+1} = \frac{t_M}{\overline{m}_s C} + 1 + \log A_n$ 

![](_page_25_Picture_1.jpeg)

![](_page_25_Picture_2.jpeg)

![](_page_25_Figure_3.jpeg)

Note: the Wald's test is equivalent to series of T(d,c), where  $c = \lceil (\log A - d \log \lambda_1) / \log \lambda_0 \rceil$ 

WaldSAC - Optimal Randomised RANSAC

## Exp. 1: Wide-baseline matching

![](_page_26_Picture_2.jpeg)

![](_page_26_Picture_3.jpeg)

![](_page_26_Picture_4.jpeg)

	samples	models	V	time	spd-up
R	2914	7347	110.0	1099504	1.0
R-R	7825	19737	- 3.0	841983	1.3
Wald	3426	8648	8.2	413227	2.7

#### 25RANSAC workshop CVPR Exp. 2 Narrow-baseline stereo

![](_page_27_Picture_1.jpeg)

![](_page_27_Picture_2.jpeg)

J. Matas, O. Chum,

2006 éž

![](_page_27_Picture_3.jpeg)

	samples	models	V	time	spd-up
R	155	367	600.0	235904	1.0
R-R	247	587	86.6	75539	3.1
Wald	162	384	-23.1	25032	9.4

WaldSAC - Optimal Randomised RANSAC

#### J. Matas, O. Chum. 2SRANSAC workshop CVPR Randomised Verification in RANSAC: 2006 // Conclusions Perception Prague

- The same confidence  $\eta$  in the solution reached faster (data dependent,  $\approx$ 10x)
- No change in the character of the algorithm, it was randomised anyway.
- Optimal strategy derived using Wald's theory for known  $\varepsilon$  and  $\delta$ .
- Results with  $\varepsilon$  and  $\delta$  estimated during the course of RANSAC are not significantly different. Performance of SPRT is insensitive to errors in the estimate.
- $\delta$  can be learnt, an initial estimate can be obtained by geometric consideration
- Lower bound on  $\varepsilon$  is given by the best-so-far support •
- Note that the properties of WaldSAC are quite different from preemptive **RANSAC!**

## PROSAC – PROgressive SAmple Consensus

- Not all correspondences are created equally
- Some are better than others
- Sample from the best candidates first

$$1 2 3 4 5 \dots N-2 N-1 N$$

Sample from here

# **PROSAC** Samples

Draw  $T_l$  samples from  $(1 \dots l)$ Draw  $T_{l+1}$  samples from  $(1 \dots l+1)$ 

Samples from  $(1 \dots l)$  that are not from  $(1 \dots l+1)$  contain

![](_page_30_Picture_4.jpeg)

Draw  $T_{l+1}$  -  $T_l$  samples of size *m*-1 and add

![](_page_30_Picture_6.jpeg)

## Locally Optimized RANSAC

![](_page_31_Picture_2.jpeg)

It was observed experimentally, that RANSAC takes several times longer than theoretically expected. This is due to the noise on inlier measurement – not every all-inlier sample generates a good hypothesis.

By applying local optimization (LO) to the-best-so-far hypotheses: (i) a near perfect agreement with theoretical (i.e. optimal) performance (ii) lower sensitivity to noise and poor conditioning. The LO is shown to be executed so rarely that it has minimal impact on the execution time.

Chum, Matas, Kittler: Locally Optimized RANSAC, DAGM 2003

## Locally Optimized RANSAC

![](_page_32_Picture_2.jpeg)

Perception Pragu

Estimation of (approximate) models with lower complexity (less data points in the sample) followed by LO step estimating the desired model speeds the estimation up significantly.

The estimation of epipolar geometry is up to 10000 times faster when using 3 region-to-region correspondences rather than 7 point-to-point correspondences.

![](_page_32_Picture_5.jpeg)

![](_page_32_Picture_6.jpeg)

Simultaneous estimation of radial distortion and epipolar geometry with LO is superior to the state-of the art in both speed a precision of the model.

Fish-eye images by Braňo Mičušík Chum, Matas, Obdržálek: Enhancing RANSAC by Generalized Model Optimization, ACCV 2004

## **Degenerate** Configurations

![](_page_33_Picture_2.jpeg)

The DEGENSAC algorithm handles scenes with:
all points in a single plane
majority of the points in a single plane and the rest off the plane
no dominant plane present

No a-priori knowledge of the type of the scene is required

![](_page_33_Picture_5.jpeg)

**Chum, Werner, Matas**: Epipolar Geometry Estimation unaffected by dominant plane, *CVPR* 2005

![](_page_34_Picture_1.jpeg)

![](_page_34_Picture_2.jpeg)

![](_page_35_Picture_0.jpeg)

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