Multivariate Analysis of Variance

Jiří Kléma

Department of Computer Science, Czech Technical University in Prague



http://cw.felk.cvut.cz/wiki/courses/b4m36san/start

Agenda

- Explain ANOVA
 - relationship between continuous variables and a categorical variable
 - * categorical variable = treatment, factor,
 - relationship with t-test for two groups,
 - posthoc tests to find out which groups contributed most,
 - relationship with linear regression,
- Generalize towards MANOVA
 - procedure for comparing multivariate sample means,
 - two-way modification, non-parametric.

Categorical dependent vs continuous independent variable

Review t-test for two groups

- a test in which the test statistic follows a Student's t-distribution ...

- under the null hypothesis,
- consider a two sample t-test, $H_0: \mu_1 = \mu_2$, $H_a: \mu_1 \neq \mu_2$
 - the two populations should follow a normal distribution,
 - variances of the two populations assumed equal \rightarrow Student's t-tests,
 - variances can differ \rightarrow Welch's test (see below),

$$t_{obs} = \frac{\bar{X}_1 - \bar{X}_2}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} \sim t_{df}$$

 $- \bar{X}_i$, s_i^2 and n_i ...sample means, variances and sizes, $- df \leq n_1 + n_2 - 2$, the exact formula complicated, $- \text{ reject } H_0 \text{ if } |t_{obs}| \geq t_{df,1-\alpha/2}.$

t-distribution



Statlect: The Digital Textbook

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T-test for multiple groups

 \blacksquare Concern a categorical variable with many levels \rightarrow multiple groups

- the hypotheses of interest

*
$$H_0: \mu_1 = \mu_2 = \cdots = \mu_g$$
,
* $H_a: \mu_i \neq \mu_j$ for at least one $i \neq j$.

- conduct a two-sample t-test for a difference in means for each pair of groups
 - the number of comparisons grows quadratically with the number of groups/levels,
- $\hfill \ensuremath{\,\bullet\)}$ for $\alpha=0.05$ for each comparison
 - there is a 5% chance that each comparison will falsely be called significant,
 - the overall probability of Type I error is elevated above 5%,
 - we falsely reject at least one of the partial null hypothesis with probability

$$1 - (1 - \alpha)^{\binom{g}{2}}$$

- e.g., for g=4 it makes $0.26 \gg \alpha$,
- multiple comparisons must be corrected.

Multiple comparisons must be corrected

- often we control family-wise error rate (FWER)
 - the probability of making one or more false discoveries (type I errors) when performing multiple hypotheses tests,
 - the most simple FWER control is the **Bonferroni correction**,
 - test each hypothesis at level $\alpha_{indiv}=\alpha_{overall}/m$,
 - $\ast\ m$ stands for the number of individual pair tests,
 - \ast follows from Bonferroni inequality for independent tests

$$\alpha_{overall} = 1 - (1 - \alpha)^m \le m\alpha_{indiv}$$

* in our case with 4 groups $m = \binom{4}{2} = 6$, * the B. inequality obviously holds

$$0.26 = 1 - 0.95^6 < 0.05 * 6 = 0.3$$

- however, this adjustment may be too conservative

* insufficient power, often does not reject H_0 although H_a is true.

Analysis of variance (ANOVA)

- compares means for multiple (usually $g \ge 3$) independent populations
 - parametric and unpaired, one-way,
 - relationship between a categorical factor F and a continuous outcome Y,
 - extends a two sample t-test to multiple groups,

Subject	F	Y	
1	f_1	y_1	
2	f_2	y_2	
Ν	f_N	y_N	



- y_{ij} ... observation for subject j in group i,
- $n_i \dots$ number of subjects in group i,
- $N = n_1 + n_2 + \ldots + n_g \ldots$ total sample size.

Analysis of variance (ANOVA)

assumptions

- the subjects are independently sampled
 - * employ repeated measures ANOVA otherwise,
- the data are normally distributed in each group
 - * $E(Y_{i.}) = \mu_i$, e.g., no group sub-populations with different means,
 - \ast residuals of the model below show the normal distribution

 $y_{ij} = \mu + \alpha_i + \epsilon_{ij} = \mu_i + \epsilon_{ij}$

- * employ non-parametric Kruskal-Wallis test otherwise,
- the data are **homoscedastic**
 - * the variability in the data does not depend on group membership,
 - * there is a common variance $var(Y_{ij}) = \sigma^2$,
- the hypotheses of interest

$$H_0$$
 : $\mu_1=\mu_2=\dots=\mu_g$,

 $-H_a: \mu_i \neq \mu_j$ for at least one $i \neq j$.

Analysis of variance (ANOVA)

method

- partition SS_{total} , the total variation in a response variable,
- distinguish within groups variability SS_{error} ,
- and between groups variability SS_{treat} ,

$$SS_{total} = \sum_{i=1}^{g} \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_{..})^2 =$$

$$= \sum_{i=1}^{g} \sum_{j=1}^{n_i} ((y_{ij} - \bar{y}_{i.}) + (\bar{y}_{i.} - \bar{y}_{..}))^2 =$$

$$= \sum_{i=1}^{g} \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_{i.})^2 + \sum_{i=1}^{g} n_i (\bar{y}_{i.} - \bar{y}_{..})^2$$

$$\underbrace{SS_{error}}_{SS_{treat}}$$

* $\bar{y}_{i.} = \frac{1}{n_i} \sum_{j=1}^{n_i} y_{ij} \dots$ group i sample mean, * $\bar{y}_{..} = \frac{1}{N} \sum_{i=1}^{g} \sum_{j=1}^{n_i} y_{ij} \dots$ grand mean. method

 in a similar manner, partition the number of degrees of freedom that stand behind the observed sums of the squared deviations

$$DF_{total} = N - 1 = DF_{error} + DF_{treat} = (N - g) + (g - 1) = N - 1$$

- decide whether group averages differ more than based on random variability observed in the dependent variable under the null hypothesis,
- employ mean square variability, both within groups and between groups

$$MS_{error} = \frac{SS_{error}}{DF_{error}} = \frac{SS_{error}}{N-g} \qquad MS_{treat} = \frac{SS_{treat}}{DF_{treat}} = \frac{SS_{treat}}{g-1}$$

method

- compare the variance between the groups and within the groups,

$$F_{obs} = \frac{MS_{treat}}{MS_{error}} \sim F_{g-1,N-g}$$

- if F_{obs} is small (close to 1), then variability between groups is negligible compared to variation within groups and the grouping does not explain much variation in the data,
- if F_{obs} is large, then variability between groups is large compared to variation within groups and the grouping explains a lot of the variation in the data
- decision rule based on F_{obs}
 - reject H_0 if $F_{obs} \geq F_{\alpha,g-1,N-g}$,
 - fail to reject H_0 if $F_{obs} < F_{\alpha,g-1,N-g}$.

- after performing ANOVA (and rejecting the null hypothesis)
 - we only assume that there is some difference in group means,
- a post-hoc test identifies which particular groups stand behind the test outcome,
- Tukey's HSD (honest significant difference) test
 - a t-test that controls for family-wise error rate (FWER),
 - compares all pairs of group means,
 - identifies all pairs whose difference is larger than expected standard error,
 - observed test statistics related to the studentized range distribution,

$$q_{obs} = \frac{\bar{y}_{i.} - \bar{y}_{j.}}{\sqrt{\frac{MS_{error}}{n^*}}} \sim q_{g,N-g}$$

 $-n^*$. . . number of observations per group (their harmonic mean if not equal), - always positive, sort the means before its application.

ANOVA vs linear regression

Is there any link between testing of linear models and ANOVA?

$$F_{ANOVA} = \frac{SS_{treat}/(g-1)}{SS_{error}/(N-g)}$$
$$F_{LR} = \frac{(TSS - RSS)/p}{RSS/(m-p-1)}$$

ANOVA vs linear regression

Is there any link between testing of linear models and ANOVA?

$$F_{ANOVA} = \frac{SS_{treat}/(g-1)}{SS_{error}/(N-g)}$$
$$F_{LR} = \frac{(TSS - RSS)/p}{RSS/(m-p-1)}$$

- the same principle,

- exactly the same outcome for single independent categorical variable.

ANOVA extensions/alternatives

up to now we talked about ANOVA that

— is parametric,

- deals with independent measurements,
- is one-way (with a single factor),
- concerns a single target variable only,
- other options
 - non-parametric analysis (Wilcoxon test \rightarrow Kruskal-Wallis analysis),
 - compares all possible group means (repeated measures ANOVA, Friedman test if non-parametric too),
 - main effects ANOVA and factorial ANOVA,
 - multivariate ANOVA (MANOVA).





Sir Ronald Fisher (1890-1962), evolutionary biologist and statistician

Statistical Methods for Research Workers

> R. A. FISHER, M.A. Fellow of Gonville and Caius College, Cambridge Chief Statistician, Rothamsted Experiment Station

BY

OLIVER AND BOYD EDINBURGH: TWEEDDALE COURT LONDON: 33 PATERNOSTER ROW, E.C. 1925

His work considered to define modern statistics

Multivariate analysis of variance (MANOVA)

- $\hfill p$ variables measured on each subject, objects categorized into g disjoint groups.
- y_{ijk} ... an observation for variable k from subject j in group i,
- $y_{ij} \dots a$ vector of dependent variables for subject j in group i,
- assumptions
 - the subjects are **independently sampled**,
 - the data are **multivariate normally distributed** in each group,
 - the data from all groups have common covariance matrix $\Sigma_{\rm r}$
 - the data from group i has common mean vector $\mu_{\mathbf{i}}$ of length p,
- the hypotheses of interest

$$-H_0: \mu_1=\mu_2=\cdots=\mu_{\mathbf{g}}$$
,

 $-H_a: \mu_{ik} \neq \mu_{jk}$ for at least one $i \neq j$ and at least one variable k.

Multivariate analysis of variance (MANOVA)

method

- the analogy of SS_{total} in ANOVA is a $p \times p$ cross products matrix T,
- similarly to ANOVA, it can be decomposed into the Error Sum of Squares and Cross Products E, and the Hypothesis Sum of Squares and Cross Products H.

$$\begin{split} \mathbf{T} &= \sum_{i=1}^{g} \sum_{j=1}^{n_{i}} (\mathbf{y}_{ij} - \bar{\mathbf{y}}_{..}) (\mathbf{y}_{ij} - \bar{\mathbf{y}}_{..})' = \\ &= \sum_{i=1}^{g} \sum_{j=1}^{n_{i}} \{ (\mathbf{y}_{ij} - \bar{\mathbf{y}}_{i.}) + (\bar{\mathbf{y}}_{i.} - \bar{\mathbf{y}}_{..}) \} \{ (\mathbf{y}_{ij} - \bar{\mathbf{y}}_{i.}) + (\bar{\mathbf{y}}_{i.} - \bar{\mathbf{y}}_{..}) \}' = \\ &= \sum_{i=1}^{g} \sum_{j=1}^{n_{i}} (\mathbf{y}_{ij} - \bar{\mathbf{y}}_{i.}) (\mathbf{y}_{ij} - \bar{\mathbf{y}}_{i.})' + \sum_{i=1}^{g} \mathbf{n}_{i} (\bar{\mathbf{y}}_{i.} - \bar{\mathbf{y}}_{..}) (\bar{\mathbf{y}}_{i.} - \bar{\mathbf{y}}_{..})' \\ &= \underbrace{\sum_{i=1}^{g} \sum_{j=1}^{n_{i}} (\mathbf{y}_{ij} - \bar{\mathbf{y}}_{i.}) (\mathbf{y}_{ij} - \bar{\mathbf{y}}_{i.})'}_{\mathbf{E}} + \underbrace{\sum_{i=1}^{g} \mathbf{n}_{i} (\bar{\mathbf{y}}_{i.} - \bar{\mathbf{y}}_{..}) (\bar{\mathbf{y}}_{i.} - \bar{\mathbf{y}}_{..})'}_{\mathbf{H}} \end{split}$$

* $\bar{\mathbf{y}}_{..} = \frac{1}{N} \sum_{i=1}^{g} \sum_{j=1}^{n_i} \mathbf{y}_{ij} \dots$ grand mean vector of length p.

*

 ${\scriptstyle \bullet}$ explanation of the elements of ${\bf T},\, {\bf E}$ and ${\bf H}$

- the element $\mathbf{t}_{k,l}$ is

$$\sum_{i=1}^{g} \sum_{j=1}^{n_i} (y_{ijk} - \bar{y}_{..k})(y_{ijl} - \bar{y}_{..l})$$

- for k = l it is the total sum of squares for variable k, and measures the total variation in the kth variable, for $k \neq l$, this measures the dependence between variables k and l across all of the observations,
- the element $\mathbf{e}_{k,l}$ is

$$\sum_{i=1}^{g} \sum_{j=1}^{n_i} (y_{ijk} - \bar{y}_{i.k}) (y_{ijl} - \bar{y}_{i.l})$$

- for k = l it is the error sum of squares for variable k, and measures the within treatment variation for the kth variable, for $k \neq l$ it measures the dependence between variables k and l after taking into account the treatment,

- ${\scriptstyle \bullet}$ explanation of the elements of ${\bf T},\, {\bf E}$ and ${\bf H}$
 - the element $\mathbf{h}_{k,l}$ is

$$\sum_{i=1}^{g} n_i (\bar{y}_{i.k} - \bar{y}_{..k}) (\bar{y}_{i.l} - \bar{y}_{..l})$$

- for k = l it is the treatment sum of squares for variable k, and measures the between treatment variation for the kth variable, for $k \neq l$, this measures dependence of variables k and l across treatments.
- consequently, if the hypothesis sum of squares and cross products H is large relative to the error sum of squares and cross products matrix E we wish to reject H₀.

Multivariate analysis of variance (MANOVA)

- Wilk's lambda test statistics for MANOVA (several other statistics exist too)
 - the determinant of the error matrix \mathbf{E} is divided by the determinant of the total matrix $\mathbf{T} = \mathbf{H} + \mathbf{E}$, we will reject the null hypothesis if Wilk's lambda is small/close to zero as then \mathbf{H} is large relative to \mathbf{E} too.

$$\Lambda^* = \frac{|\mathbf{E}|}{|\mathbf{H} + \mathbf{E}|}$$

- can also be computed using the eigenvalues $\hat{\lambda}$ of $\mathbf{E^{-1}H}(s = min(p, g-1))$

$$\Lambda^* = \prod_{i=1}^s \frac{1}{1+\hat{\lambda}_i}$$

- the distribution of Λ^* is not tractable, we can only have approximations,

- e.g., Bartlett's approximation can be used if N is large

$$-(N-1-\frac{p+g}{2})\ln\Lambda^* > \chi^2_{p(g-1),\alpha}$$

A typical case in which MANOVA helps

- Mechanical engineering domain
 - 90 samples of three different alloys (A, B, C),
 - samples differ in flexibility and strength,
 - flexibility and strength correlated, strength in C slightly increased,
 - goal: decide (detect) the influence of alloy on flexibility and strength.



A typical case in which MANOVA helps

• ANOVA outcome:

MANOVA outcome:

Summary

- MANOVA compares multivariate sample means
 - it deals with multiple dependent variables at the same time,
- MANOVA advantages over ANOVA
 - better chance to discover which factor is truly important,
 - protects against Type I errors in multiple independent ANOVA runs,
 - increased power, it can reveal differences not discovered by ANOVA tests,
- MANOVA cautions
 - a complicated design, more difficult to disambiguate,
 - one degree of freedom is lost for each dependent variable that is added,
 - unsuitable if the dependent variables are perfectly correlated or uncorrelated,
- typically followed by significance tests on individual dependent variables.

The main references

- :: Resources (slides, scripts, tasks) and reading
 - STAT 505 course on Applied Multivariate Statistical Analysis, PennState University, https://onlinecourses.science.psu.edu/stat505/.
 - G. James, D. Witten, T. Hastie and R. Tibshirani: An Introduction to Statistical Learning with Applications in R. Springer, 2014.
 - A. C. Rencher, W. F. Christensen: Methods of Multivariate Analysis.
 3rd Edition, Wiley, 2012.
 - T. Hastie, R. Tibshirani and J. Friedman: The Elements of Statistical Learning: Data Mining, Inference, and Prediction. Springer, 2009.